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# BAYESIAN SINGLE SAMPLING attribute plans For discrete PRIOR DISTRIBUTIONS 

A. HALD


København 1965

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# Bayesian single sampling attribute plans For discrete PRIOR DISTRIBUTIONS 

BY
A. HALD


København 1965
Kommissionær: Ejnar Munksgaard

## Synopsis

The paper gives a rather complete tabulation and discussion of properties of a system of single sampling attribute plans obtained by minimizing average costs under the assumptions that costs are linear in $p$, the fraction defective, and that the distribution of lot quality is a double binomial distribution. The optimum sampling plan $(n, c)$ depends on 6 parameters $\left(N, p_{r}, p_{s}, p_{1}, p_{2}, w_{2}\right)$ where $N$ denotes lot size, $\left(p_{r}, p_{s}\right)$ are suitably normalized cost parameters, and $\left(p_{1}, p_{2}, w_{2}\right)$ are the parameters of the prior distribution. It may be shown, however, that the weights combine with the $p$ 's in such a way that only 5 independent parameters are left.

A procedure to obtain the exact solution of the problem has been developed in a previous paper and this procedure is used for computing a set of master tables in which $p_{r}=p_{s}=0.01$ and $0.10, w_{2}=0.05$, ( $p_{1}, p_{2}$ ) take on suitably chosen values in relation to the value of $p_{r}$, and $1 \leqq N \leqq 200,000$.

The properties of the optimum plans are studied, and simple conversion formulas are derived which makes it possible to find the optimum plan for an arbitrary set of parameters from a plan in the master tables with a "corresponding" set of parameters. The main tool for this investigation is the asymptotic expressions for the acceptance number and for the sample size, viz. $c=n p_{0}+a+o(1)$ and $n=\frac{1}{\varphi_{0}}\left(\ln N-\frac{1}{2} \ln \ln N+\right.$ $\left.\ln \lambda+\frac{3}{2} \ln \varphi_{0}\right)+o(1)$, where $p_{0}$ and $\varphi_{0}$ are functions of $\left(p_{1}, p_{2}\right)$ only, whereas $a$ and $\lambda$ depend on the other parameters also. It is furthermore shown that the minimum value of the standardized costs per lot asymptotically equals the costs of sampling inspection plus a constant and that the producer's and the consumer's risks tend to zero inversely proportional to lot size. By means of the asymptotic formulas it is possible to find out how $(n, c)$ vary with the individual parameters and derive two general conversion formulas.

Efficiency of various other systems of sampling plans is studied in relation to the present model and some general recommendations are made.

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## 1. Introduction and Summary

TThe main purpose of the present paper is to give a rather complete tabulation and discussion of properties of a system of single sampling attribute plans obtained by minimizing average costs under the assumptions that costs are linear in $p$, the fraction defective, and that the distribution of lot quality is a double binomial distribution.

Starting from a cost function containing 6 parameters and a mixed binomial prior distribution it is shown how the average costs may be written in a standard form containing only two parameters, $p_{r}$ and $p_{s}$, besides the parameters defining the prior distribution. The one parameter, $p_{r}$, is the economic break-even quality and depends on the costs of acceptance and rejection only, whereas the second parameter, $p_{s}$, also depends on the costs of sampling inspection and the average quality. In a simple and practically important case $p_{r}$ and $p_{s}$ denote the costs of rejection and the costs of sampling inspection, respectively, divided by the costs of accepting a defective item.

Specializing the prior distribution to a double binomial distribution defined by the two quality levels $\left(p_{1}, p_{2}\right)$ and the weights $\left(w_{1}, w_{2}\right), w_{1}+w_{2}=1$, it will be seen that the optimum sampling plan $(n, c)$ depends on the 6 parameters $\left(N, p_{r}, p_{s}, p_{1}, p_{2}, w_{2}\right)$ where $N$ denotes lot size. It may be shown, however, that the weights combine with the $p$ 's in such a way that only 5 (independent) parameters are left.

A procedure to obtain the exact solution of the problem has been developed in a previous paper and this has been used for computing a set of master tables in which $p_{r}=p_{s}=0.01$ and $0.10, w_{2}=0.05,\left(p_{1}, p_{2}\right)$ take on suitably chosen values in relation to the value of $p_{r}$, and $1 \leqq N \leqq 200,000$.

In the remaining part of the paper the properties of the optimum plans are studied with the purpose to derive simple conversion formulas which will make it possible to find the optimum plan for an arbitrary set of parameters from a plan in the master table with a "corresponding" set of parameters. The main tool for this investigation is the asymptotic expressions for the acceptance number and for the sample size, viz.

$$
c=n p_{0}+a+o(1) \quad \text { and } \quad n=\frac{1}{\varphi_{0}}\left(\ln N-\frac{1}{2} \ln \ln N+\ln \lambda+\frac{3}{2} \ln \varphi_{0}\right)+o(1)
$$

where $p_{0}$ and $\varphi_{0}$ are functions of $\left(p_{1}, p_{2}\right)$ only, whereas $a$ and $\lambda$ depend on the other parameters also. It is furthermore shown that the minimum value of the standardized costs per lot asymptotically equals the costs of sampling inspection plus a constant (depending on $\left(p_{1}, p_{2}\right)$ ) and that the producer's and the consumer's risks tend to zero inversely proportional to lot size. Numerical investigations show that the asymptotic expressions give good approximations to the optimum plan even for quite small values of $N$.

By means of the asymptotic formulas it is possible to find out how ( $n, c$ ) vary with the individual parameters. One of the most important results is found by letting all the $p$ 's tend to zero which leads to "the proportionality law": The optimum sampling plan corresponding to $\left(N, \lambda p_{r}, \lambda p_{s}, \lambda p_{1}, \lambda p_{2}, w_{2}\right)$ is approximately equal to $\left(n^{*} / \lambda, c^{*}\right)$ where $\left(n^{*}, c^{*}\right)$ is the plan corresponding to $\left(N^{*}, p_{r}, p_{s}, p_{1}, p_{2}, w_{2}\right)$ with $N^{*}=N \lambda$.

This theorem combined with other similar results regarding the effect of varying the individual parameters lead to two general conversion formulas stated in sections 8 and 11. A summary of these formulas is given at the end of the paper in connection with the tables.

Efficiency of a sampling plan is defined as the ratio of the standardized costs (loss) of the optimum plan and the costs of the plan in question. Efficiency is discussed for various alternative systems and the efficiency of using optimum plans determined from wrong values of the parameters is studied.

Finally the present system is discussed in relation to other systems and it is pointed out that from an economic point of view it is not advisable to fix the consumer's or the producer's risk. If one wants a system with a fixed risk then the risk should be fixed to 50 per cent at a point between $p_{1}$ and $p_{2}$. Two such IQL systems are then briefly discussed.

## 2. The model

Several authors have studied economic models, mostly linear, for the determination of single sampling inspection plans by attributes, see for instance [1] and [2].

We shall here start from the formulation proposed by Guthrie and Johns [3] and show how the model may be reduced to a standard form as previously used by Hald [4].

Let $N$ and $n$ denote lot size and sample size and let $X$ and $x$ denote number of defectives in the lot and the sample, respectively. The acceptance number is denoted by $c$.

Let the costs be
and

$$
\begin{equation*}
n S_{1}+x S_{2}+(N-n) A_{1}+(X-x) A_{2} \quad \text { for } x \leqq c \tag{1}
\end{equation*}
$$

$$
\begin{equation*}
n S_{1}+x S_{2}+(N-n) R_{1}+(X-x) R_{2} \quad \text { for } x>c \tag{2}
\end{equation*}
$$

The interpretation of the six cost parameters depends on the kind of inspection envisaged, i. e. whether inspection is a consumer's receiving inspection, a producer's
inspection of finished goods, or "internal inspection" by delivery of goods from one department to another within the same firm. The cost parameters may have quite different values when considered exclusively from a producer's or a consumer's point of view because certain costs are borne primarily by one of the parties involved. The values of the cost parameters also depend on whether the inspection is rectifying or non-rectifying, destructive or non-destructive. In the following the two cost expressions are discussed and a few examples of interpretation are given.

Costs associated with the sample, $n S_{1}+x S_{2}$, for brevity called "costs of sampling inspection" consist of two parts: one part, $n S_{1}$, proportional to the number of items in the sample so that $S_{1}$ includes sampling and testing costs per item, and another part, $x S_{2}$, proportional to the number of defectives in the sample, i.e. $S_{2}$ denotes additional costs for an inspected defective item. If defective items found in the sample are repaired, say, then $S_{2}$ includes the repair costs per item.
"Costs of acceptance" are similarly composed of a part, $(N-n) A_{1}$, proportional to the number of items in the remainder of the lot, and another part, $(X-x) A_{2}$, proportional to the number of defective items accepted. Whereas $A_{1}$ usually will be zero or negligible, $A_{2}$ will often be considerable. If accepted items are used as parts in an assembly operation, say, $A_{2}$ may include the manufacturing costs (or the price) of an item, the costs of handling the defective item in assembling and disassembling, and the damage to other parts used in the assembly. In case of inspection of finished goods $A_{2}$ may include costs of repair, service and guarantees plus loss of good-will.
"Costs of rejection" consist of a part, $(N-n) R_{1}$, proportional to the number of items in the remainder of the lot, and another part, $(X-x) R_{2}$, proportional to the number of defective items rejected. Rejection is here taken in a broad sense meaning only that the lot cannot be accepted according to the sampling plan used. Rejection may therefore lead to sorting, price reduction, scrapping, or salvaging. If rejection means sorting, say, then $R_{1}$ includes sorting costs per item and $R_{2}$ denotes additional costs for defective items found, for example costs of repair or replacement.

It is obvious that from a practical point of view it will in general be easiest to obtain information on the values of the cost parameters in the case of "internal inspection".

Denoting the hypergeometric probability by

$$
p\{x \mid X\}=\binom{n}{x}\binom{N-n}{X-x} /\binom{N}{X}
$$

the average costs for lots of size $N$ with $X$ defectives become

$$
\begin{gather*}
K(N, n, c, X)=\sum_{x=0}^{n}\left(n S_{1}+x S_{2}\right) p\{x \mid X\}+\sum_{x=0}^{c}\left((N-n) A_{1}+(X-x) A_{2}\right) p\{x \mid X\} \\
+\sum_{x=c+1}^{n}\left((N-n) R_{1}+(X-x) R_{2}\right) p\{x \mid X\} \tag{3}
\end{gather*}
$$

Let $f_{N}(X)$ denote the (prior) distribution of $X$, i. e. the distribution of lot quality. The average costs then become

$$
\begin{equation*}
K(N, n, c)=\sum_{X} K(N, n, c, X) f_{N}(X) . \tag{4}
\end{equation*}
$$

As shown in [4] this expression becomes linear in $N$ for the important class of mixed binomial distributions, i.e. for

$$
\begin{equation*}
f_{N}(X)=\int_{0}^{1}\binom{N}{X} p^{X} q^{N-X} d W(p) \tag{5}
\end{equation*}
$$

where $W(p)$ denotes a cumulative distribution function (independent of $N$ ).
From (3)-(5) we find

$$
\begin{equation*}
K(N, n, c)=\int_{0}^{1} K(N, n, c, p) d W(p) \tag{6}
\end{equation*}
$$

where

$$
\begin{gather*}
K(N, n, c, p)=n\left(S_{1}+S_{2} p\right)+(N-n)\left(\left(A_{1}+A_{2} p\right) P(p)+\left(R_{1}+R_{2} p\right) Q(p)\right),  \tag{7}\\
P(p)=B(c, n, p)=\sum_{x=0}^{c}\binom{n}{x} p^{x} q^{n-x}, \tag{8}
\end{gather*}
$$

and $Q(p)=1-P(p)$.
For convenience the frequency function corresponding to $W(p)$ will be called the distribution of the process average or the distribution of $p$ as distinct from $f_{N}(X)$ which gives the distribution of $X / N$, i.e. the distribution of lot quality. (The following discussion will be in terms of $p$ ).

Limiting the prior distributions to mixed binomials, (6) shows that the average costs may be considered as an average of the cost function (7), which is a function of $p$, with respect to the distribution of $p$. It should be noted that this result is valid for any $(N, n)$ for a mixed binomial prior distribution and that a similar result holds for $N \rightarrow \infty, n \rightarrow \infty$, and $n / N \rightarrow 0$, for any prior distribution. The limit theorems derived in the following may therefore be applied in general.

The sampling plans discussed are obtained by minimizing $K(N, n, c)$ according to (6) with respect to ( $n, c$ ) for given cost parameters and prior distribution and they will be called Bayesian single sampling plans or optimum plans.

Starting from (7) we introduce the three cost functions

$$
\begin{align*}
& k_{s}(p)=S_{1}+S_{2} p,  \tag{9}\\
& k_{a}(p)=A_{1}+A_{2} p, \tag{10}
\end{align*}
$$

and

$$
\begin{equation*}
k_{r}(p)=R_{1}+R_{2} p, \tag{11}
\end{equation*}
$$

defined for $0 \leqq p \leqq 1$. We shall make the following assumptions regarding these functions:

1. All three functions are non-negative and none of them is identical zero.
2. $k_{a}(0)<k_{r}(0)$ and $k_{a}(1)>k_{r}(1)$, from which follows that the equation $k_{a}(p)=k_{r}(p)$ has the solution

$$
\begin{equation*}
p_{r}=\left(R_{1}-A_{1}\right) /\left(A_{2}-R_{2}\right), \quad 0<p_{r}<1, \tag{12}
\end{equation*}
$$

$p_{r}$ being called the (economic) break-even quality.
3. $k_{s}(p) \geqq k_{m}(p)$ for $0 \leqq p \leqq 1$, where

$$
k_{m}(p)=\left\{\begin{array}{ll}
k_{a}(p) & \text { for } p \leqq p_{r}  \tag{13}\\
k_{r}(p) & \text { for } p>p_{r}
\end{array}\right\}
$$

The function $k_{m}(p)$ gives the unavoidable (minimum) costs, i.e. the costs corresponding to the situation where perfect knowledge of quality exists without costs and all lots are classified correctly on basis of the corresponding process average, viz. accepted for $p \leqq p_{r}$ and rejected for $p>p_{r}$.

Averages over the prior distribution are denoted by $k_{s}, k_{a}$, etc., i.e.

$$
\begin{equation*}
k_{a}=\int_{0}^{1} k_{a}(p) d W(p)=k_{a}(\bar{p})=A_{1}+A_{2} \bar{p}, \tag{14}
\end{equation*}
$$

and

$$
\begin{equation*}
k_{m}=\int_{0}^{1} k_{m}(p) d W(p)=\int_{0}^{p_{r}} k_{a}(p) d W(p)+\int_{p_{r}}^{1} k_{r}(p) d W(p) \tag{15}
\end{equation*}
$$

Costs per item are denoted by $k$, costs per lot by the corresponding $K$, i.e. $K=N k$.
The average costs for the three cases without sampling inspection, i.e. the cases where
(a) all lots are classified correctly,
(b) all lots are accepted, and
(c) all lots are rejected,
then become $k_{m}, k_{a}$, and $k_{r}$, respectively. These cases are useful "reference cases" since sampling inspection is justified only if $k-k_{m}<\min \left\{k_{a}-k_{m}, k_{r}-k_{m}\right\}$, where $k=K(N, n, c) / N$.

Case (a) will usually be considered as the basic reference case and average costs for other cases will therefore be reduced by $k_{m}$, since $k_{m}$ represents the average fixed costs per item which will be incurred irrespective of the decision made. The cost differences

$$
k_{a}-k_{m}=\int_{p_{r}}^{1}\left(k_{a}(p)-k_{r}(p)\right) d W(p)
$$

and

$$
k_{r}-k_{m}=\int_{0}^{p_{r}}\left(k_{r}(p)-k_{a}(p)\right) d W(p)
$$

represent average decision losses in case (b) and (c) respectively, and $k_{s}-k_{m}$ represents the average "loss" by inspection.

From (6) and (15) we find

$$
\begin{equation*}
K=n k_{s}+(N-n) \int_{0}^{1}\left(k_{a}(p) P(p)+k_{r}(p) Q(p)\right) d W(p) \tag{16}
\end{equation*}
$$

and

$$
K_{m}=n k_{m}+(N-n)\left\{\int_{0}^{p_{r}} k_{a}(p) d W(p)+\int_{p_{r}}^{1} k_{r}(p) d W(p)\right\}
$$

leading to

$$
\begin{gather*}
K-K_{m}=n\left(k_{s}-k_{m}\right) \\
+(N-n)\left\{\begin{array}{l}
\int_{0}^{p_{r}}\left(k_{r}(p)-k_{a}(p)\right) Q(p) d W(p)+\int_{p_{r}}^{1}\left(k_{a}(p)-k_{r}(p)\right) P(p) d W(p)
\end{array}\right\} \\
=n\left(k_{s}-k_{m}\right)+(N-n)\left(A_{2}-R_{2}\right)\left\{\int_{0}^{p_{r}}\left(p_{r}-p\right) Q(p) d W(p)+\int_{p_{r}}^{1}\left(p-p_{r}\right) P(p) d W(p)\right\}, \tag{17}
\end{gather*}
$$

the two terms giving the average costs of sampling inspection and the average decision losses, respectively.

Instead of minimizing $K$ with respect to ( $n, c$ ) we might just as well minimize $K-K_{m},\left(K-K_{m}\right) /\left(A_{2}-R_{2}\right)$, or $\left(K-K_{m}\right) /\left(k_{s}-k_{m}\right)$, since $K_{m}, A_{2}-R_{2}$, and $k_{s}-k_{m}$ are independent of $(n, c)$. It will be seen from (17) that it is practical to use $A_{2}-R_{2}$ or $k_{s}-k_{m}$ as "economic unit".

Defining

$$
\begin{equation*}
p_{m}=\int_{0}^{p_{r}} p d W(p)+\int_{p_{r}}^{1} p_{r} d W(p)=p_{r}-\int_{0}^{p_{r}}\left(p_{r}-p\right) d W(p), \tag{18}
\end{equation*}
$$

we find $0 \leqq p_{m} \leqq p_{r}$ and

$$
\begin{equation*}
p_{r}-p_{m}=\left(k_{r}-k_{m}\right) /\left(A_{2}-R_{2}\right) . \tag{19}
\end{equation*}
$$

Defining $p_{s}$ by means of

$$
\begin{equation*}
p_{s}-p_{m}=\left(k_{s}-k_{m}\right) /\left(A_{2}-R_{2}\right) \tag{20}
\end{equation*}
$$




Fig. 1. Example of cost functions.
we find

$$
\begin{equation*}
p_{s}-p_{r}=\left(k_{s}-k_{r}\right) /\left(A_{2}-R_{2}\right) \tag{21}
\end{equation*}
$$

and

$$
\begin{equation*}
p_{s}=\left\{\left(S_{1}-A_{1}\right)+\left(S_{2}-R_{2}\right) \bar{p}\right\} /\left(A_{2}-R_{2}\right) . \tag{22}
\end{equation*}
$$

Introducing

$$
\begin{equation*}
R^{*}(N, n, c)=\left\{K(N, n, c)-K_{m}\right\} /\left(A_{2}-R_{2}\right) \tag{23}
\end{equation*}
$$

we find the standard form

$$
\begin{equation*}
R^{*}=n\left(p_{s}-p_{m}\right)+(N-n)\left\{\int_{0}^{p_{r}}\left(p_{r}-p\right) Q(p) d W(p)+\int_{p_{r}}^{1}\left(p-p_{r}\right) P(p) d W(p)\right\} \tag{24}
\end{equation*}
$$

containing only two parameters $p_{r}$ and $p_{s}-p_{m}$, instead of the six cost parameters in the original model, see [4]. It should be noted that $p_{s}-p_{m}$ depends on the prior distribution besides on the cost parameters.

Consider the special case given by $k_{a}(p)=A_{2} p, k_{r}(p)=R_{1}$, and $k_{s}(p)=S_{1}$, which is a model commonly used in practice. It follows that $p_{r}=R_{1} / A_{2}$ and $p_{s}=S_{1} / A_{2}$, i.e. $p_{r}$ and $p_{s}$ are the costs of rejection and of sampling and testing, respectively, measured with the cost of accepting a defective item. This simple interpretation of $p_{r}$ and $p_{s}$ is one of the reasons for using them as parameters.

It is often useful to discuss the problem in terms of the simple cost functions $k_{a}(p)=p, k_{r}(p)=p_{r}$, and $k_{s}(p)=p_{s}$, which immediately lead to the form (24). The corresponding form of (7) becomes

$$
K_{0}(p)=n p_{s}+(N-n)\left(p P(p)+p_{r} Q(p)\right)
$$

from which the general form may be found as

$$
K(p)=\left(A_{2}-R_{2}\right) K_{0}(p)+\left(n S_{2}+(N-n) R_{2}\right) p+\left(N A_{1}-n\left(S_{2}-R_{2}\right) \bar{p}\right)
$$

A sketch of the cost functions for a typical case has been given in Fig. 1, which is based on the data in section 13.

For some purposes it is useful to use $k_{s}-k_{m}$ as economic unit instead of $A_{2}-R_{2}$. Putting

$$
R(N, n, c)=\left\{K(N, n, c)-K_{m}\right\} /\left(k_{s}-k_{m}\right),
$$

i.e.

$$
R=R^{*} /\left(p_{s}-p_{m}\right),
$$

we find

$$
\begin{equation*}
R=n+\frac{N-n}{p_{s}-p_{m}}\left\{\int_{0}^{p_{r}}\left(p_{r}-p\right) Q(p) d W(p)+\int_{p_{r}}^{1}\left(p-p_{r}\right) P(p) d W(p)\right\}, \tag{25}
\end{equation*}
$$

the two terms again giving the costs of sampling inspection and the average decision losses, respectively, but here using the average costs of sampling inspection (minus $k_{m}$ ) per item in the sample as economic unit.

In the next section we shall discuss the determination of ( $n, c$ ) for a double binomial distribution as prior distribution. This means that $p$ is a random variable taking on only two values, $p_{1}<p_{r}<p_{2}$, with probabilities $w_{1}$ and $w_{2}=1-w_{1}$, respectively. From (25) we then find

$$
\begin{equation*}
R=n+(N-n)\left(\gamma_{1} Q\left(p_{1}\right)+\gamma_{2} P\left(p_{2}\right)\right) \tag{26}
\end{equation*}
$$

where

$$
\begin{gather*}
\gamma_{i}=\left|p_{i}-p_{r}\right| w_{i}\left|\left(p_{s}-p_{m}\right)=\left|k_{a}\left(p_{i}\right)-k_{r}\left(p_{i}\right)\right| w_{i}\right|\left(k_{s}-k_{m}\right), \quad i=1,2,  \tag{27}\\
p_{m}=p_{1} w_{1}+p_{r} w_{2} \tag{28}
\end{gather*}
$$

i.e. $R$ depends on four parameters only, viz. $p_{1}, p_{2}, \gamma_{1}, \gamma_{2}$.

The correspondingly standardized costs for the cases of acceptance and rejection without inspection are

$$
\begin{equation*}
R_{a}=N\left(k_{a}-k_{m}\right) /\left(k_{s}-k_{m}\right)=N \gamma_{2} \tag{29}
\end{equation*}
$$

and

$$
\begin{equation*}
R_{r}=N\left(k_{r}-k_{m}\right) /\left(k_{s}-k_{m}\right)=N \gamma_{1} \tag{30}
\end{equation*}
$$

These results may also be obtained from (26) for $n=0$ by setting $P(p)=1$ and 0 , respectively.

If acceptance without inspection is cheaper than rejection without inspection, i. e. $k_{a}<k_{r}$ we find $\bar{p}<p_{r}$ and $\gamma_{2}<\gamma_{1}$.

In the special case $k_{s}=k_{r}$ we have $p_{s}=p_{r}$ and $\gamma_{1}=1$ so that the model contains only three parameters.

It should be noted that

$$
\begin{equation*}
\gamma_{2}=\frac{\bar{p}-p_{m}}{p_{s}-p_{m}}=1-\frac{p_{s}-\bar{p}}{p_{s}-p_{m}} \tag{31}
\end{equation*}
$$

and

$$
\begin{equation*}
\gamma_{1}=\frac{p_{r}-p_{m}}{p_{s}-p_{m}}=1-\frac{p_{s}-p_{r}}{p_{s}-p_{m}} \tag{32}
\end{equation*}
$$

## 3. The exact solution and the tables

In a previous paper [4] we have proved the following theorem:
For a double binomial (prior) distribution of lot quality given by the parameters $\left(p_{1}, p_{2}, w_{2}\right)$ and for linear cost functions (1) and (2) the Bayesian single sampling plan may be found by minimizing $R(N, n, c)$, see (26), with respect to $(n, c)$. The solution satisfies the two inequalities

$$
\begin{equation*}
\alpha+\beta c \leqq n<\alpha+\beta(c+1) \tag{33}
\end{equation*}
$$

and

$$
\begin{equation*}
F(n-1, c) \leqq N<F(n, c) \tag{34}
\end{equation*}
$$

where

$$
\begin{gather*}
\alpha=\log \frac{w_{2}\left(p_{2}-p_{r}\right)}{w_{1}\left(p_{r}-p_{1}\right)}\left|\log \frac{q_{1}}{q_{2}}=\log \frac{\gamma_{2}}{\gamma_{1}}\right| \log \frac{q_{1}}{q_{2}},  \tag{35}\\
\left.\beta=\log \frac{p_{2} q_{1}}{q_{2} p_{1}} \right\rvert\, \log \frac{q_{1}}{q_{2}}, \tag{36}
\end{gather*}
$$

and

$$
\begin{equation*}
F(n, c)=n+1+\frac{p_{s}-p_{r}+\sum_{i} w_{i}\left(p_{r}-p_{i}\right) B\left(c, n, p_{i}\right)}{\sum_{i} w_{i}\left(p_{i}-p_{r}\right) p_{i} b\left(c, n, p_{i}\right)} \tag{37}
\end{equation*}
$$

For two plans $\left(n_{1}, c_{1}\right)$ and $\left(n_{2}, c_{2}\right), c_{1}<c_{2}$ say, satisfying (33) and having overlapping $N$-intervals according to (34) $R\left(N, n_{1}, c_{1}\right) \lesseqgtr R\left(N, n_{2}, c_{2}\right)$ for $N \lesseqgtr N_{12}$ where

$$
\begin{equation*}
N_{12}=\frac{\left(p_{s}-p_{r}\right)\left(n_{2}-n_{1}\right)+n_{2} \gamma\left(n_{2}, c_{2}\right)-n_{1} \gamma\left(n_{1}, c_{1}\right)}{\gamma\left(n_{2}, c_{2}\right)-\gamma\left(n_{1}, c_{1}\right)} \tag{38}
\end{equation*}
$$

and

$$
\begin{equation*}
\gamma(n, c)=\sum_{i} w_{i}\left(p_{r}-p_{i}\right) B\left(c, n, p_{i}\right) \tag{39}
\end{equation*}
$$

In [4] the theorem was derived as a special case of a more general one. We shall here derive the theorem directly from (26) using the same method as in [4].

Values of $(n, c)$ minimizing $R$ must satisfy the two inequalities
and

$$
\begin{equation*}
\Delta_{c} R(N, n, c-1) \leqq 0<\Delta_{c} R(N, n, c), \quad 0 \leqq c \leqq n, \tag{40}
\end{equation*}
$$

$$
\begin{equation*}
\Delta_{n} R(N, n-1, c) \leqq 0<\Delta_{n} R(N, n, c), \quad c \leqq n \leqq N, \tag{41}
\end{equation*}
$$

$\Delta$ denoting the usual forward difference operator.
Noting that $\Delta_{c} B(c, n, p)=b(c+1, n, p)$ and $\Delta_{n} B(c, n, p)=-p b(c, n, p)$ we find from (26)

$$
\begin{equation*}
\Delta_{c} R(N, n, c)=(N-n)\left\{-\gamma_{1} b\left(c+1, n, p_{1}\right)+\gamma_{2} b\left(c+1, n, p_{2}\right)\right\} \tag{42}
\end{equation*}
$$

and

$$
\begin{equation*}
\Delta_{n} R(N, n, c)=1-\left\{\gamma_{1} Q\left(p_{1}\right)+\gamma_{2} P\left(p_{2}\right)\right\}+(N-n-1)\left\{\gamma_{1} p_{1} b\left(c, n, p_{1}\right)-\gamma_{2} p_{2} b\left(c, n, p_{2}\right)\right\} . \tag{43}
\end{equation*}
$$

Inserting these expressions into (40) and (41) and solving for $n$ and $N$, respectively, immediately leads to (33) and (34). From $R\left(N, n_{1}, c_{1}\right)=R\left(N, n_{2}, c_{2}\right)$ we next determine $N_{12}$ by solving for $N$.

A sketch of $R$ as function of $n$ and $c$ for fixed $N$ has been given in Fig. 2 for a typical case.

The economic interpretation of (40) and (42) is the following: For given $n$ the optimum value of $c$ is determined such that a change of $c$, an increase by 1 say, will give nearly no change of the total decision loss, since the loss due to the increased consumer's risk is nearly balanced by the gain due to the smaller producer's risk.

Similarly the interpretation of (41) and (43) is that for given $c$ the optimum value of $n$ is determined such that a change of $n$, an increase by 1 say, will give nearly no change of total costs, since the increase of sampling inspection costs by 1 minus the average decision loss for one item is nearly balanced by the decrease in decision losses for the remainder of the lot.


Fig. 2. $R(N, n, c)$ as function of $n$ and $c$ for $N=1000, p_{r}=p_{s}=0.10, p_{1}=0.06, p_{2}=0.20$, and $w_{2}=0.05$.

Tabulation of optimum plans may be carried out by starting from the smallest value of $c$ giving a positive $n(n \geqq c)$ according to (33), i.e. $c_{m}=[-\alpha /(\beta-1)]$, [ ] denoting "the integer part of". For consecutive values of $c, n$ - and $N$-intervals are computed from (33) and (34) and in case of overlapping $N$-intervals costs are compared by means of (38). A detailed example may be found in [4]. The tables have been computed by this method on an electronic computer.

The sampling plans have been tabulated for two "quality levels", viz. $p_{r}=$ $=p_{s}=0.01$ and 0.10 , for one value of the weight function $w_{2}=0.05$, for 8 values of $p_{1} / p_{r}$, and for 10 , respectively 5 , values of $p_{2} / p_{r}$, giving a total of 120 tables. Each table gives $(n, c)$ as function of $N$ for $N \leqq 200,000$.

For $p_{r}=0.01$ the search for optimum plans has been limited to values of $n$ which are multiples of 5 .

These tables will be referred to as "master tables" since optimum plans for other values of the parameters may easily be found from the tabulated ones by means of conversion formulas developed in the following sections.

The exact solution has been modified in one respect. For a given value of $c$ the first and last $N$-interval may be rather short as compared to the other intervals. As an example consider the following section of the original table for $p_{r}=p_{s}=0.010$, $p_{1}=0.006, p_{2}=0.020, w_{2}=0.05$ :

| $N$ | $n$ | $c$ | $\Delta N$ |
| :---: | :---: | :---: | ---: |
| $4010-4370$ | 165 | 3 | 360 |
| $4370-4420$ | 170 | 3 | 50 |
| $4420-4430$ | 240 | 4 | 10 |
| $4430-4920$ | 245 | 4 | 490 |
| $4920-5570$ | 250 | 4 | 650 |
| $5570-5590$ | 255 | 4 | 20 |
| $5590-5610$ | 325 | 5 | 20 |
| $5610-6250$ | 330 | 5 | 640 |

The example shown is an extreme one with small intervals occurring at the beginning as well as at the end of each section of the table. It is naturally without any interest to use the sampling plan $(240,4)$ for $4420<N<4430$ and then change to $(245,4)$ for $4430<N<4920$. To eliminate such small intervals from the final table it was decided to discard the first and the last sampling plan for a given $c$ if the length of the corresponding $N$-interval was less than $1 / 5$ of the length of the neighbouring interval. In such cases the value of $N$ according to (38) was computed for the new neighbouring plans, $(165,3)$ and $(245,4)$ say, to find the optimum $N$-intervals for the remaining plans. The result of this procedure is in most cases practically equal to incorporating the small N -intervals into the larger neighbouring intervals, for example using $(245,4)$ for $4420<N<4920$.

To save space every second $N$-interval for a given value of $c$ has been omitted because the corresponding sampling plans may be found by adding $1\left(p_{r}=0.10\right)$ and $5\left(p_{r}=0.01\right)$, respectively, to $n$ for the preceding interval.

Values of $N$ have been rounded to 3 significant figures and tabulation has been stopped at $N=200,000$.

As mentioned above the tables were designed as master tables from which optimum plans may be derived for other values of the parameters and for this reason it was decided to tabulate the complete solution with respect to $N$ to make interpolation superfluous.

The user of the tables in practice may easily derive a simplified set of tables from the given ones, either by using a set of fixed $N$-intervals, or a set of fixed $N$ arguments. An example has been given in the table on page 17.

The "natural" parameters of the model are ( $p_{1}, p_{2}, w_{2}$ ), which characterize the prior distribution, and ( $p_{r}, p_{s}$ ), which depend on the costs. The tables and the properties of the solution will be discussed in terms of these parameters on basis of the results in the next section. However, one property may be stated immediately from the observation that the solution depends on four parameters only, viz. ( $p_{1}, p_{2}, \gamma_{1}, \gamma_{2}$ ). The three parameters ( $p_{r}, p_{s}, w_{2}$ ) may therefore in respect to the solution be considered as functionally related, i. e. combinations of $\left(p_{r}, p_{s}, w_{2}\right)$ giving the same $\left(\gamma_{1}, \gamma_{2}\right)$ will lead to the same sampling plan.

Single Sampling Tables for $100 p_{r}=100 p_{s}=1.0,100 p_{1}=0.5$, and $w_{2}=0.05$.


## From

$$
\begin{aligned}
& \gamma_{2}=\begin{array}{l}
\left(p_{2}-p_{r}\right) w_{2} \\
\gamma_{1}
\end{array}=\begin{array}{l}
\left(p_{r}-p_{1}\right) w_{1}
\end{array}, ~
\end{aligned}
$$

we find

$$
\begin{equation*}
p_{r}-p_{1}=\left(p_{2}-p_{1}\right) /\left(1+\frac{\gamma_{2} w_{1}}{\gamma_{1} w_{2}}\right) \tag{44}
\end{equation*}
$$

From

$$
\gamma_{1}\left(p_{s}-p_{m}\right)=\left(p_{r}-p_{1}\right) w_{1}
$$

and

$$
p_{s}-p_{m}=p_{s}-p_{1}-w_{2}\left(p_{r}-p_{1}\right)
$$

we find

$$
\begin{equation*}
p_{s}-p_{1}=\left(p_{r}-p_{1}\right)\left(\frac{w_{1}}{\gamma_{1}}+w_{2}\right) . \tag{45}
\end{equation*}
$$

These formulas show how $p_{r}$ and $p_{s}$ depend on $w_{2}$ for given $\left(p_{1}, p_{2}, \gamma_{1}, \gamma_{2}\right)$. To use them in connection with the master tables we put $p_{r}=p_{s}$ and $w_{2}=0.05 \lambda$ which leads to

$$
p_{r}(\lambda)=p_{10}+\left(p_{20}-p_{10}\right) /\left(1-\gamma_{20}+\frac{20 \gamma_{20}}{\lambda}\right)
$$

where

$$
\gamma_{20}=\frac{p_{20}-p_{r 0}}{19\left(p_{r 0}-p_{10}\right)}=\frac{\varrho_{2}-1}{19\left(1-\varrho_{1}\right)}
$$

the index 0 denoting an argument in the master table, $p_{r 0}=0.01$ or $0.10, \varrho_{i}=p_{i 0} / p_{r 0}$. Dividing by $p_{r 0}$ gives

$$
\begin{equation*}
p_{r}(\lambda) / p_{r 0}=\varrho_{1}+\left(\varrho_{2}-\varrho_{1}\right) /\left(1-\gamma_{20}+\frac{20 \gamma_{20}}{\lambda}\right)=f\left(u_{2}, \varrho_{1}, \varrho_{2}\right) \tag{46}
\end{equation*}
$$

which has been tabulated in the appendix.
The field of application of the master tables may therefore be considerably enlarged by making use of the following rule:

The optimum sampling plan for $\left(N, p_{r 0}, p_{10}, p_{20}, w_{2}=0.05\right), p_{r 0}=p_{s 0}$, is the same as the plan for $\left(N, p_{r 0} f\left(w_{2}, \varrho_{1}, \varrho_{2}\right), p_{10}, p_{20}, w_{2}\right)$.

Consider for example the case with $p_{r 0}=p_{s 0}=0.010, p_{1}=0.006, p_{2}=0.040$, and $w_{2}=0.05$ for which the optimum plans have been given in the master table. The same plans are also optimum for $w_{2}=0.20$, say, and $p_{r}=p_{s}=0.019$, $p_{1}=0.006$, and $p_{2}=0.040$ which may be seen by interpolation in the table of $f\left(w_{2}, \varrho_{1}, \varrho_{2}\right)$ for $\varrho_{1}=0.6$ and $\varrho_{2}=4.0$.

## 4. The asymptotic solution

In this section we shall give a somewhat simpler and more direct proof of the asymptotic results found by Guthrie and Johns [3] and by Hald [4], and furthermore carry the asymptotic expansion so far that we get a useful approximation to the exact solution also for small values of $c$.

The proof is based on the following lemma which is a special case of a theorem proved by Blackwell and Hodges [5]:

For $c / n=h=p_{0}+\varepsilon, p_{0}$ being a constant and $\varepsilon \rightarrow 0$ for $n \rightarrow \infty$, we have

$$
\begin{equation*}
P(p)=\frac{1}{\sqrt{2 \pi n p_{0} q_{0}}} \frac{q_{0} p}{\left(p-p_{0}\right)} e^{-n \varphi(h, p)}(1+O(\sqrt{\varepsilon})) \quad \text { for } p_{0}<p \tag{47}
\end{equation*}
$$

where

$$
\begin{equation*}
\varphi(h, p)=h \ln \frac{h}{p}+(1-h) \ln \frac{1-h}{1-p} \tag{48}
\end{equation*}
$$

For $p_{0}>p$ the same expression is valid for $Q(p)$ if only $p-p_{0}$ is replaced by $p_{0}-p$.

Writing

$$
\begin{equation*}
\varphi(h, p)=\varphi\left(p_{0}, p\right)+\varepsilon \varphi^{\prime}\left(p_{0}, p\right)+O\left(\varepsilon^{2}\right) \tag{49}
\end{equation*}
$$

where

$$
\begin{equation*}
\varphi^{\prime}\left(p_{0}, p\right)=\ln \frac{p_{0} q}{q_{0} p} \tag{50}
\end{equation*}
$$

we find from (26) and (47) the asymptotic expression

$$
\begin{equation*}
R=n+(N-n) \frac{q_{0}}{\sqrt{2 \pi n p_{0} q_{0}}} \sum_{i=1}^{2} \frac{\gamma_{i} p_{i}}{\mid p_{0}-p_{i}} e^{-n \varphi\left(p_{0}, p_{i}\right)-n \varepsilon \varphi^{\prime}\left(p_{0}, p_{i}\right)}(1+O(\sqrt{\varepsilon})) \tag{51}
\end{equation*}
$$

on the assumption that $p_{1}<p_{0}<p_{2}$. (As will be shown later $\varepsilon=O(1 / n)$, and we may therefore disregard $n \varepsilon^{2}$ ). We shall first determine the value of $h=p_{0}+\varepsilon$ which minimize $R$ for given $n$ and next determine the value of $n$ giving the absolute minimum by treating $R$ as a differentiable function of $n$.

The essential feature of (51) is that the two binomial risks, $Q\left(p_{1}\right)$ and $P\left(p_{2}\right)$, have been expressed as functions tending exponentially to zero for $n \rightarrow \infty$.

As explained in [4] the optimum plan must have the property that $R / N \rightarrow 0$ for $N \rightarrow \infty, n \rightarrow \infty$, and $n / N \rightarrow 0$. It follows that $p_{0}$ must satisfy the inequality $p_{1}<p_{0}<p_{2}$ because otherwise $R / N$ would not tend to zero but to $\gamma_{1}$ or $\gamma_{2}$.

We shall state the theorem to be proved for the double binomial distribution only, but it is valid for a more general class of distributions, viz. for a distribution having probability density $w(p)=0$ for $p_{1}<p<p_{2}, w\left(p_{1}\right)=w_{1}>0, w\left(p_{2}\right)=$ $w_{2}>0, w_{1}+w_{2} \leqq 1$, and

$$
\int_{0}^{p_{1}^{*}} d W(p)+\int_{p_{2}^{*}}^{1} d W(p)=1-w_{1}-w_{2}
$$

for $0 \leqq p_{1}^{*}<p_{1}$ and $p_{2}<p_{2}^{*} \leqq 1$, which means that the probability distribution may be arbitrary outside the interval $p_{1}^{*}<p<p_{2}^{*}$. The result of such a generalization will only be to add a term to (51) of form

$$
\frac{N-n}{p_{s}-p_{m}} \frac{q_{0}}{\sqrt{2 \pi n p_{0} q_{0}}} \int_{I} \frac{\left(p_{r}-p\right) p}{p_{0}-p} e^{-n \varphi(h, p)} d W(p),
$$

(I denoting the intervals $\left(0 \leqq p \leqq p_{1}^{*}\right)$ and $\left(p_{2}^{*} \leqq p \leqq 1\right)$ ) which obviously is $O\left(e^{-n}\right)$ times the last term of (51) since $\varphi(h, p)>\varphi\left(h, p_{1}\right)$ for $p<p_{1}$ and $\varphi(h, p)>\varphi\left(h, p_{2}\right)$ for $p>p_{2}$.

Because of the factor $p_{r}-p$ in the cost function we might also have assumed that $w\left(p_{r}\right)>0$ without altering the result.

It is reasonable to assume that the two exponential terms in (51) tend to zero with the same speed, i.e. that $p_{0}$ is determined from

$$
\varphi\left(p_{0}, p_{1}\right)=\varphi\left(p_{0}, p_{2}\right)
$$

which gives

$$
\begin{equation*}
p_{0}=\left(\ln \frac{q_{1}}{q_{2}}\right) /\left(\ln \frac{p_{2} q_{1}}{q_{2} p_{1}}\right)=\frac{1}{\beta} \tag{52}
\end{equation*}
$$

and

$$
\begin{equation*}
\varphi_{0}=p_{0} \ln \frac{p_{0}}{p_{i}}+q_{0} \ln \frac{q_{0}}{q_{i}}, \quad i=1 \text { or } 2 . \tag{53}
\end{equation*}
$$

Under this assumption we shall determine $\varepsilon$ by minimization of (51). The part of $R$ depending on $\varepsilon$ is

$$
f(\varepsilon)=\sum_{i} \frac{\gamma_{i} p_{i}}{\left|p_{0}-p_{i}\right|} e^{-n \varepsilon \varphi^{\prime}\left(p_{0}, p_{i}\right)} .
$$

From $f^{\prime}(\varepsilon)=0$ we find

$$
\begin{equation*}
\sum_{i=1}^{2} \frac{\gamma_{i} p_{i} \varphi_{i}^{\prime}}{\mid p_{0}-p_{i}} e^{-n \varepsilon \varphi_{i}^{\prime}}-0 \tag{54}
\end{equation*}
$$

where-according to (50)-

$$
\begin{equation*}
\varphi_{i}^{\prime}=\ln \frac{p_{0} q_{i}}{q_{0} p_{i}} . \tag{55}
\end{equation*}
$$

Solving for $a=n \varepsilon$ we find

$$
\begin{equation*}
a \delta_{0}^{\prime}=\ln \frac{\gamma_{1} p_{1}\left(p_{2}-p_{0}\right) \varphi_{1}^{\prime}}{\gamma_{2} p_{2}\left(p_{0}-p_{1}\right)\left(-\varphi_{2}^{\prime}\right)} \tag{56}
\end{equation*}
$$

where

$$
\begin{equation*}
\delta_{0}^{\prime}-\varphi_{1}^{\prime}-\varphi_{2}^{\prime}=\ln \frac{p_{2} q_{1}}{q_{2} p_{1}} . \tag{57}
\end{equation*}
$$

We thus have the result that $c=n p_{0}+a+o(1)$ in accordance with what could be expected from (33).

Inserting these results into (51) we find

$$
\begin{equation*}
R=n+(N-n) \frac{\lambda}{\sqrt{n}} e^{-n \varphi_{0}} \tag{58}
\end{equation*}
$$

with

$$
\begin{equation*}
\lambda=\frac{q_{0}}{\sqrt{2 \pi p_{0} q_{0}}} \sum_{i=1}^{2} \frac{\gamma_{i} p_{i}}{\left|p_{0}-p_{i}\right|} e^{-a \varphi_{i}^{\prime}} . \tag{59}
\end{equation*}
$$

To prove (indirectly) that $h=p_{0}+\varepsilon$ minimizes $R$ let us assume that $h=p_{0}+\varepsilon$, given by (52) and (56), does not minimize $R$ but that min $R$ is obtained for $h=h_{0}+\varepsilon_{0}$, $h_{0} \neq p_{0}$ and $\varepsilon_{0} \rightarrow 0$. Denoting the part of $R$ depending on $h$ by $g(h)$ we find for sufficiently large $n$ and for $h_{0}<p_{0}$, say, that

$$
g\left(h_{0}\right)=\lambda_{1}\left(h_{0}\right) e^{-n \varphi\left(h_{0}, p_{1}\right)}\left(1+O\left(e^{-n}\right)\right)
$$

since $\varphi\left(h_{0}, p_{2}\right)>\varphi\left(h_{0}, p_{1}\right)$ for $h_{0}<p_{0}$. However, $g\left(h_{0}\right)$ cannot be min $g(h)$ since $\varphi\left(h_{0}, p_{1}\right)<\varphi\left(p_{0}, p_{1}\right)$, i. e. we have reached a contradiction by assuming $h_{0} \neq p_{0}$.

From $d R / d n=0$ we find

$$
\begin{equation*}
1-(N-n) \frac{\lambda}{\sqrt{n}} e^{-n \varphi_{0}}\left(\varphi_{0}+\frac{1}{2 n}\right)-\frac{\lambda}{\sqrt{n}} e^{-n \varphi_{0}}=0 \tag{60}
\end{equation*}
$$

or

$$
\begin{equation*}
\ln (N-n)=\varphi_{0} n+\frac{1}{2} \ln n-\ln \left(\lambda \varphi_{0}\right)+o(1) . \tag{61}
\end{equation*}
$$

From (58) and (60) we also have that

$$
\begin{equation*}
\min _{(n, c)} R=n+\frac{1}{\varphi_{0}}+o(1) \tag{62}
\end{equation*}
$$

where $n$ may be determined by inversion of (61), i.e.

$$
n=\frac{1}{\varphi_{0}}\left(\ln N-\frac{1}{2} \ln \ln N+\ln \lambda+\frac{3}{2} \ln \varphi_{0}\right)+o(1) .
$$

We have thus found that asymptotically $c$ is a linear function of $n$, and $n$ is proportional to $\ln N-\frac{1}{2} \ln \ln N$ plus a constant. Furthermore it follows from (62) that the average decision loss per lot tends to a constant $1 / \varphi_{0}$ so that for large lots decision losses divided by sampling inspection costs tend to zero.

To investigate the two risks asymptotically we find from (54)

$$
\begin{aligned}
& \gamma_{1} p_{1} \varphi_{1}^{\prime} \\
& p_{0}-p_{1}
\end{aligned} e^{-a \varphi_{1}^{\prime}}=\frac{\gamma_{2} p_{2}\left(-\varphi_{2}^{\prime}\right)}{p_{2}-p_{0}} e^{-a \varphi_{2}^{\prime}}
$$

so that (59) gives

$$
\lambda=\frac{q_{0}}{\sqrt{2 \pi p_{0} q_{0}}} \frac{\gamma_{1} p_{1} \delta_{0}^{\prime}}{\left(p_{0}-p_{1}\right)\left(-\varphi_{2}^{\prime}\right)} e^{-a \varphi_{1}^{\prime}}
$$

which together with (60) may be used to reduce

$$
Q\left(p_{1}\right)=\frac{1}{\sqrt{2 \pi p_{0} q_{0}}} \frac{q_{0} p_{1}}{p_{0}-p_{1}} e^{-a \varphi_{1}^{\prime}} \frac{1}{\sqrt{n}} e^{-n \varphi_{0}}
$$

to

$$
\begin{equation*}
Q\left(p_{1}\right)=\frac{-\varphi_{2}^{\prime}}{\varphi_{0} \gamma_{1} \delta_{0}^{\prime}} \frac{1}{N-n} . \tag{63}
\end{equation*}
$$

Similarly we have

$$
\begin{equation*}
P\left(p_{2}\right)=\frac{\varphi_{1}^{\prime}}{\varphi_{0} \gamma_{2} \delta_{0}^{\prime}} \frac{1}{N-n} \tag{64}
\end{equation*}
$$

so that

$$
\begin{equation*}
P\left(p_{2}\right) / Q\left(p_{1}\right)=\gamma_{1} \varphi_{1}^{\prime} / \gamma_{2}\left(-\varphi_{2}^{\prime}\right) \tag{65}
\end{equation*}
$$

We have thus proved the following theorem:
Asymptotically the optimum sampling plan is given by

$$
\begin{equation*}
c=n p_{0}+a+o(1) \tag{66}
\end{equation*}
$$

and

$$
\begin{equation*}
n=\frac{1}{\varphi_{0}}\left(\ln N-\frac{1}{2} \ln \ln N+\ln \lambda+\frac{3}{2} \ln \varphi_{0}\right)+o(1) \tag{67}
\end{equation*}
$$

which lead to

$$
\begin{aligned}
& \min R=\frac{1}{\varphi_{0}}\left(\ln N-\frac{1}{2} \ln \ln N+\ln \lambda+\frac{3}{2} \ln \varphi_{0}+1\right)+o(1) \\
& Q\left(p_{1}\right)=\frac{-\varphi_{2}^{\prime}}{\varphi_{0} \gamma_{1} \delta_{0}^{\prime}} \frac{1}{N-n}+o\binom{1}{N}
\end{aligned}
$$

and

$$
P\left(p_{2}\right)=\frac{\varphi_{1}^{\prime}}{\varphi_{0} \gamma_{2} \delta_{0}^{\prime}} \frac{1}{N-n}+o\left(\frac{1}{N}\right)
$$

It will be noted that $p_{0}$ and $\varphi_{0}$ depend on $\left(p_{1}, p_{2}\right)$ only, i. e. they are independent of the cost parameters and $w_{2}$.

The asymptotic solution supplements the exact one in several respects. Since the optimum plan is a function of 5 parameters $\left(N, p_{1}, p_{2}, \gamma_{1}, \gamma_{2}\right)$ a complete tabulation is rather hopeless even if a program has been worked out for an electronic computer. Furthermore the properties of the exact solution are not easily to be found from the procedure by which the solution is obtained. The advantages of the asymptotic solution are that
(1) it clearly shows how the optimum plan and various derived quantities depend on the parameters,
(2) it may be used as starting point for developing approximations which are valid also for small $N$,
(3) it may be used for developing interpolation and extrapolation formulas in connection with "master tables" of the exact solution, and
(4) it shows the sensitivity of the solution with respect to changes of the parameters.

These aspects of the solution will be discussed in the following sections.

## 5. Comparison of exact and approximate solution

Looking at the relation between $n$ and $c$ in the tables it will be seen that the optimum values of $n$ for a given value of $c$ tend to cluster around

$$
\begin{equation*}
n_{c}=\alpha+\beta\left(c+\frac{1}{2}\right) \tag{69}
\end{equation*}
$$

as might be expected from (33). Comparing with the asymptotic result $c=n p_{0}+a$, $p_{0}=1 / \beta$ and $a$ being defined by (56), agreement between the two expressions would require that

$$
\left.\left(\ln \frac{p_{2}\left(p_{0}-p_{1}\right)\left(-\varphi_{2}^{\prime}\right)}{p_{1}\left(p_{2}-p_{0}\right) \varphi_{1}^{\prime}}\right)\right)\left(\ln \frac{p_{2} q_{1}}{q_{2} p_{1}}\right)=\frac{1}{2} .
$$

It can be proved that the ratio on the left hand side above is positive and less that 1. Numerical investigations show that in typical cases in practice the ratio does not deviate much from $1 / 2$. As examples consider the following results:

| $100 p_{1}$ | $100 p_{2}$ | $p_{2} / p_{1}$ | Ratio |
| :---: | :---: | :---: | :---: |
| 0.2 | 4.0 | 20 | 0.528 |
| 0.2 | 2.0 | 10 | 0.517 |
| 0.6 | 4.0 | 6.7 | 0.512 |
| 0.6 | 2.0 | 3.3 | 0.505 |

The ratio depends primarily on $p_{2} / p_{1}$ and practically the same results will be found for values of $\left(p_{1}, p_{2}\right)$ which are 10 times as large or $1 / 10$ of the values considered. We shall therefore in the following use the simpler expression (69) instead of $c=n p_{0}+a$ as the starting point for finding $n$ from $c$ or reversely.

The asymptotic formulas may be used in two ways:
(1) Starting from $c$ we may determine the corresponding $N$-interval and within that the relation between $n$ and $N$.
(2) Starting from $N$ we may determine the corresponding $n$ and from $n$ determine $c$.

The first method is useful for making a systematic tabulation of sampling plans whereas the second is suitable for computing "isolated" plans for a given $N$.

Starting from an integer value of $c$ we first find $n_{c}$ from (69) and the corresponding $N_{c}$ from (61). Similarly we find $N_{c-0.5}$ and $N_{c+0.5}$, being the lower and upper limit for $N$ having $c$ as optimum acceptance number.

In the asymptotic solution we have disregarded the discreteness of $c$ and $n$. We may, however, afterwards try to take the effect of the discreteness of $c$ into account by investigating the relationship between $n$ and $N$ for given (integer) value of $c$. From $d R(N, n, c) / d n=0$ it can be found that $n$ is approximately a linear function of
$\log N$ with slope $-1 / \log q_{2}{ }^{*}$. Within the interval $\left(N_{c-0.5}, N_{c+0.5}\right)$ we may therefore determine $n$ from the approximate formula

$$
\begin{equation*}
n=n_{c}-\left(\log N-\log N_{c}\right) / \log q_{2}, \quad N_{c-0.5}<N<N_{c+0.5}, \tag{70}
\end{equation*}
$$

which for small $p_{2}$ and small intervals may be replaced by

$$
\begin{equation*}
n=n_{c}+\left(N-N_{c}\right) / N_{c} P_{2}, \quad N_{c-0.5}<N<N_{c+0.5} . \tag{71}
\end{equation*}
$$

It follows that the values of $n$ belong to the interval

$$
n_{c} \pm \beta\left(\varphi_{0}+\frac{1}{2 n_{c}}\right) / 2 p_{2} .
$$

For applications in practice we give the formula corresponding to (61) with logarithms to base 10, i.e.

$$
\begin{equation*}
\log \left(N_{c}-n_{c}\right)=\varphi n_{c}+\frac{1}{2} \log n_{c}+\delta \tag{72}
\end{equation*}
$$

where

$$
\begin{gather*}
\varphi=p_{0} \log \frac{p_{0}}{p_{i}}+q_{0} \log \frac{q_{0}}{q_{i}}, \quad i=1 \text { or } 2,  \tag{73}\\
\delta=-\log \left(\lambda \varphi_{o}\right), \tag{74}
\end{gather*}
$$

and

$$
\begin{equation*}
\lambda \varphi_{0}=10^{\varphi\left(\alpha+\frac{\beta}{2}\right)} \varphi_{\log e} \sqrt{\frac{q_{0}}{2 \pi p_{0}}} \sum_{i=1}^{2} \frac{\gamma_{i} p_{i}}{\left|p_{0}-p_{i}\right|}\binom{q_{i}}{q_{0}}^{\alpha+\frac{\beta}{2}} \tag{75}
\end{equation*}
$$

$-a$ having been replaced by $\frac{\alpha}{\beta}+\frac{1}{2}$ in $\lambda \varphi_{o}$.
In the following we shall make much use of (72) with $N_{c}-n_{c}$ replaced by $N_{c}$ which only means that we disregard terms of order $n_{c} / N_{c}$ and less.

The approximation obtained by using (69), (70), and (72) is usually very good even for quite small values of $c$. Normally the approximate value of $c$ will deviate at most 1 from the correct value. The approximation depends essentially on $p_{2} / p_{1}$, being good for large values of $p_{2} / p_{1}$ and poorer for small values. Two examples for $p_{2} / p_{1}=6.7$ and 3.3 , respectively, will show the results obtained for a typical good and poor case. Table 1 and Fig. 3 show that the approximate and the exact solution are practically identical in the first case whereas the approximate solution in the second case often will lead to a value of $c$ being 1 too large and a corresponding value of $n$.

* This results is due to Mrs. K. West Andersen.

Table 1.
Comparisons of exact and approximate sampling plans computed from
(69), (70), and (72).

| c |  |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: |
|  | Approximation |  |  | Exact |  |
|  | $n_{c}$ | $n$ | $N_{c \pm 0.5}$ | $n$ | N |
| 1 | 57 | 43-66 | 269-714 | 45-65 | 280- 714 |
| 2 | 112 | 104-120 | 715- 1400 | 105-120 | 715-1420 |
| 4 | 223 | 216-230 | 2490- 4300 | 220-230 | 2550- 4390 |
| 6 | 334 | 328-340 | 7190-12000 | 330-340 | 7390-12300 |
| 8 | 445 | 439-451 | 19700-32300 | 440-450 | 20200-33000 |
| 10 | 556 | 550-562 | 52400-85300 | 550-560 | 53600-87000 |
| 12 | 667 | 661-673 | 137000-200000 | 665-670 | 140000-200000 |


| c | $\begin{gathered} p_{r}=p_{s}=0.010, p_{1}=0.006, p_{2}=0.020, w_{2}=0.05 . \\ \alpha=-143.0, \beta=85.879, \varphi=0.0009088, \delta=2.0785,-1 / \log q_{2}=113.97 . \end{gathered}$ |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: |
|  | Approximation |  |  | Exact |  |
|  | $n_{c}$ |  | $N_{c \pm 0.5}$ | $n$ | N |
| 2 | 72 | 44- 88 | 715-1750 | - | - |
| 4 | 243 | 234-251 | 2790-3970 | 245-250 | $4420-5590$ |
| 6 | 415 | 408-422 | 5410-7150 | 415-420 | 7100- 8980 |
| 8 | 587 | 581-593 | 9270-11900 | 585-595 | 11300-14200 |
| 10 | 759 | 753-765 | 15100-19000 | 755-765 | 17700-22000 |
| 12 | 931 | 925-936 | 23800-29600 | 930-935 | 27300-33700 |
| 14 | 1102 | 1097-1107 | 36800-45700 | 1100-1105 | 41500-51100 |
| 16 | 1274 | 1269-1279 | 56500-69700 | 1270-1280 | 62800-77200 |
| 18 | 1446 | 1441-1451 | 86000-106000 | 1445-1450 | 94600-116000 |
| 20 | 1618 | 1613-1623 | 130000-159000 | 1615-1620 | 142000-173000 |

It is essential for the efficiency of the approximation to use the right relation between $n$ and $c$, see the discussion in section 12, and it is therefore fortunate that this relation is a simple linear one.

The approximation formula breaks down for values of $N$ for which the cheapest solution is acceptance without inspection (or rejection without inspection). As will be seen from Table 1 the approximation formula may in such cases lead to a sampling plan even if no optimum plan exists. The difference in costs by using such a plan instead of accepting without sampling inspection will, however, normally be small.

Turning to the inverse formula (67) numerical investigations show that the results are not as accurate as those found from (61). Taking one more term in the inversion of (61) and changing to logarithms with base 10 we find


$$
\begin{equation*}
n_{N}=\frac{1}{\varphi}\left\{\log N-\left(\frac{1}{2} \log \log N+d\right)\left(1-\frac{1}{3 \log N}\right)\right\} \tag{76}
\end{equation*}
$$

where

$$
\begin{equation*}
d=-\log \lambda \varphi_{0}-\frac{1}{2} \log \varphi=\delta-\frac{1}{2} \log \varphi \tag{77}
\end{equation*}
$$

The exact inversion leads to the correction term $(\log e) /(2 \log N)=0.22 / \log N$ which, however, on the basis of numerical investigations has been replaced by $1 /(3 \log N)$. If $(76)$ is to be used extensively it pays to tabulate

$$
\begin{equation*}
g(N)=\log N-\frac{1}{2}\left(1-\frac{1}{3 \log N}\right) \log \log N \tag{78}
\end{equation*}
$$

and use (76) in the form

$$
\begin{equation*}
n_{N}=\frac{1}{\varphi}\left\{g(N)-d\left(1-\frac{1}{3 \log N}\right)\right\} . \tag{79}
\end{equation*}
$$

From $n$ we may then find

$$
c_{N}=p_{0}\left(n_{N}-\alpha\right)-\frac{1}{2}
$$

Table 2.
Comparisons of exact and approximate sampling plans computed from (76).

and round to the nearest integer. To obtain more accurate results $n_{c}$ may be computed from the rounded value of $c_{N}$ and $n$ could then be found from (70) or (71).

Table 2 shows that (76) leads to good results for the two previously discussed typical examples.

As a general conclusion of the many numerical comparisons which have been carried out we may state that the asymptotic formulas give sufficiently good approximations to the optimum sampling plans for most practical purposes. If one wants to be sure to find the optimum plan one may start from the approximation and compare the costs of this plan with the costs of suitably chosen neighbouring plans thus finding the optimum one by trial and error.

The formulas (72) and (77) have, however, the serious drawback from the point of view of application that the constants $\delta$ and $d$ are rather hard to compute. The asymptotic formulas have therefore in the following only been used to derive relationships between sampling plans under variation of the parameters. It is to be expected that these relationships will prove to be rather accurate in view of the good approximation demonstrated above.

According to (62) we have for the optimum plans that the average decision loss asymptotically is constant, i.e. $R-n \sim 1 / \varphi_{0}$. For small $N$ this gives an upper limit to the decision loss but the formula is not of much value as an approximation.

Fig. 4 sketches for the two previously considered examples $R-n$ as function of $N$. The discontinuities correspond to changes in $c$; each time $c$ is increased by one $n$

increases approximately by $\beta$ and $R-n$ decreases with the same quantity. The asymptotic result corresponds to the mid-points of the intervals. It will be seen that the asymptote is nearly being reached for $N=100,000$ in the case with $p_{2} / p_{1}=6.7$ but not for $p_{2} / p_{1}=3.3$.

For small $N$ a useful upper limit to the average decision loss may be obtained by noticing that $R<N \gamma_{2}$ if an optimum plan exists and the alternative is acceptance without inspection.

According to (63) and (64) the probabilities of wrong decisions, $Q\left(p_{1}\right)$ and $P\left(p_{2}\right)$, are asymptotically inversely proportional to N. Fig. 5 sketches for the two examples $Q\left(p_{1}\right)$ and $P\left(p_{2}\right)$ as functions of $N$. The asymptotic formula gives a reasonable approximation to $P\left(p_{2}\right)$ in both cases, whereas the approximation to $Q\left(p_{1}\right)$ is rather poor, particularly for the case $p_{2} / p_{1}=3.3$. The discontinuities resulting from changes of $c$ are very pronounced for $Q\left(p_{1}\right)$.

## 6. Proportional change of $\left(p_{r}, p_{s}, p_{1}, p_{2}\right)$ for fixed $w_{2}$

We shall first study the asymptotic formulas for all "quality levels" tending to zero with the same speed. Introducing the auxiliary quantities

$$
\begin{equation*}
\varrho_{s}=\frac{p_{s}}{p_{r}}, \varrho_{1}=\frac{p_{1}}{p_{r}}, \varrho_{2}=\frac{p_{2}}{p_{r}}, \varrho_{m}=\frac{p_{m}}{p_{r}}, \varrho=\frac{p_{2}}{p_{1}}, \tag{80}
\end{equation*}
$$



Fig. 5. Probabilities of wrong decisions as functions of lot size.
we find for $p_{r} \rightarrow 0$ and fixed $\left(\varrho_{s}, \varrho_{1}, \varrho_{2}, w_{2}\right)$

$$
\begin{aligned}
\alpha p_{r} & \rightarrow\left(\ln \frac{w_{2}\left(\varrho_{2}-1\right)}{w_{1}\left(1-\varrho_{1}\right)}\right) /\left(\varrho_{2}-\varrho_{1}\right)=\alpha_{0}, \\
\beta p_{r} & \rightarrow\left(\ln \frac{\varrho_{2}}{\varrho_{1}}\right) /\left(\varrho_{2}-\varrho_{1}\right)=\beta_{0}, \\
p_{0} / p_{r} & \rightarrow 1 / \beta_{0}=\varrho_{0}, \\
p_{0} / p_{r} & \rightarrow \varrho_{0} \ln \frac{\varrho_{0}}{\varrho_{i}}+\left(\varrho_{i}-\varrho_{0}\right)=\varphi^{*}, \quad i=1 \text { or } 2,
\end{aligned}
$$

and
$\lambda \varphi_{0} / \sqrt{p_{r}} \rightarrow \exp \left\{\varphi^{*}\left(\alpha_{0}+\frac{\beta_{0}}{2}\right)\right\} \frac{\varphi^{*}}{\sqrt{2 \pi \varrho_{0}}} \sum_{i=1}^{2} \frac{\omega_{i} \varrho_{i}\left(\varrho_{i}-1\right)}{\left(\varrho_{s}-\varrho_{m}\right)\left(\varrho_{i}-\varrho_{0}\right)} \exp \left\{\left(\varrho_{0}-\varrho_{i}\right)\left(\alpha_{0}+\frac{\beta_{0}}{2}\right)\right\}=\exp \left\{-\delta_{0}\right\}$,
where in the last expression $-a$ has been replaced by $\frac{\alpha}{\beta}+\frac{1}{2}$ as in (75).
Inserting these results into (69) and (72) we find

$$
n_{c} p_{r} \rightarrow \alpha_{0}+\beta_{0}\left(c+\frac{1}{2}\right)=n_{0}(c)
$$

and

$$
\ln \left(N_{c} p_{r}\right) \rightarrow \varphi^{*} n_{0}(c)+\frac{1}{2} \ln n_{0}(c)+\delta_{0}=\ln N_{0}(c) .
$$

It follows that for small $p_{r}$ we have approximately
and

$$
n_{c} \sim n_{0}(c) / p_{r}
$$

$$
N_{c} \sim N_{0}(c) / p_{r}
$$

where $n_{0}(c)$ and $N_{0}(c)$ are independent of $p_{r}$, i.e. $n$ and $N$ vary inversely proportional to $p_{r}$ for given $c$.

Suppose that the optimum sampling plans have been tabulated for a small value of $p_{r}, p_{r}=0.01$ say, and certain values of $\left(\varrho_{s}, \varrho_{1}, \varrho_{2}, w_{2}\right)$. The above result may then be used to find the optimum plans for $\lambda p_{r}$, say, from the plans in the given table. Denoting the quantities required by $n_{c}\left(\lambda p_{r}\right)$ and $N_{c}\left(\lambda p_{r}\right)$ we have for given $c$

$$
\begin{equation*}
n_{c}\left(\lambda p_{r}\right) \sim n_{c}\left(p_{r}\right) / \lambda \tag{81}
\end{equation*}
$$

and

$$
\begin{equation*}
N_{c}\left(\lambda p_{r}\right) \sim N_{c}\left(p_{r}\right) / \lambda, \tag{82}
\end{equation*}
$$

i.e. we have found the following important "proportionality law":

The optimum sampling plan corresponding to $\left(N, \lambda p_{r}, \lambda p_{s}, \lambda p_{1}, \lambda p_{2}, w_{2}\right)$ is approximately equal to $\left(n^{*} / \lambda, c^{*}\right)$ where $\left(n^{*}, c^{*}\right)$ is the plan corresponding to ( $N^{*}, p_{r}$, $\left.p_{s}, p_{1}, p_{2}, w_{2}\right)$ with $N^{*}=N \lambda$.

The theorem has been illustrated in Fig. 6 which shows that the approximation holds good also for quite large values of $p_{r}$.

This theorem greatly enlarges the field of application of the two master tables. The table with $p_{r}=0.01$ may be used for $\lambda<5$ and the table with $p_{r}=0.10$ for $0.5<\lambda<2$, in that way covering all cases with $p_{r}<0.20$ which is the domain of practical interest.

A large number of numerical investigations has shown that the proportionality law gives rather accurate results. The value of $c$ found will seldom deviate more than 1 from the correct value. For $\lambda>1$ the formula will normally tend to give too large a value of $c$ and for $\lambda<1$ too small a value.

The example in Table 3 shows the derivation of sampling plans with a breakeven quality of $p_{r}=0.03$, partly from the first master table using $\lambda=3$ and partly from the second using $\lambda=0.3$. Both results are remarkably close to the exact solution, see also Fig. 6.


Fig 6. Relation between lot size and acceptance number by proportional change of ( $p_{r}, p_{s}, p_{1}, p_{2}$ ) for fixed $w_{2}$.

Table 3.
Comparisons of exact sampling plans for $p_{r}=p_{s}=0.030, p_{1}=0.018, p_{2}=0.060$, $w_{2}=0.05$, and approximate plans derived from the master tables by the proportionality law.

| $N$ | Exact |  | Derived from$p_{r}=0.01 \quad(\lambda=3)$ |  |  | N* | $\begin{aligned} & p_{r}= \\ = & 0.3 \mathrm{~N} \end{aligned}$ | from $\begin{aligned} & (\lambda= \\ & n^{*} / 0.3 \end{aligned}$ | $c^{*}$ |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| 1000 | Accept |  | 3000 | Accept |  |  | 300 |  |  |
| 2000 | 110 | 5 | 6000 | 110 | 5 |  | 600 | 115 | 5 |
| 3000 | 140 | 6 | 9000 | 165 | 7 |  | 900 | 145 | 6 |
| 5000 | 225 | 9 | 15000 | 225 | 9 |  | 1500 | 200 | 8 |
| 7000 | 255 | 10 | 21000 | 255 | 10 |  | 2100 | 255 | 10 |
| 10000 | 310 | 12 | 30000 | 310 | 12 |  | 3000 | 285 | 11 |
| 20000 | 395 | 15 | 60000 | 395 | 15 |  | 6000 | 365 | 14 |
| 30000 | 455 | 17 | 90000 | 455 | 17 |  | 9000 | 425 | 16 |
| 50000 | 510 | 19 | 150000 | 540 | 20 |  | 15000 | 480 | 18 |
| 70000 | 570 | 21 | 210000 | 570 | 21 |  | 21000 | 535 | 20 |
| 100000 | 625 | 23 | - | - | - |  | 30000 | 565 | 21 |
| 200000 | 710 | 26 | - | - | - |  | 60000 | 675 | 25 |

Consider now the inverse formula (79). From

$$
d+\log p_{r} \rightarrow \delta_{0} \log e-\frac{1}{2} \log \varphi^{*}=d_{0}
$$

we find

$$
\begin{equation*}
n_{N}\left(\lambda p_{r}\right) \sim \frac{n_{N}\left(p_{r}\right)}{\lambda}+\left(1-\frac{1}{3 \log N}\right) \frac{\log \lambda}{\lambda \varphi\left(p_{r}\right)} \tag{83}
\end{equation*}
$$

where $\varphi\left(p_{r}\right)$ denotes the value of $\varphi$ for the given (basic) set of parameters. This formula shows how the sample size for a given lot size changes with the "quality level". This result is, however, not as accurate as the previous one for small $N$ and it is neither as convenient for use in connection with the tables.

An example has been given in the following table for $N=50,000, p_{r}=p_{s}=$ $=0.010, p_{1}=0.006, p_{2}=0.040$, and $w_{2}=0.05$.

Comparisons of exact and approximate sampling plans derived from (83).

|  | Exact |  | Approximation |  |
| :---: | ---: | ---: | :---: | ---: |
| $\lambda$ | $n$ | $c$ | $n$ | $c$ |
| 0.1 | 1850 | 3 | 2330 | 4 |
| 0.3 | 1300 | 7 | 1210 | 7 |
| 1.0 | 505 | 9 | - | - |
| 3.0 | 205 | 11 | 210 | 11 |
| 10.0 | 63 | 12 | 78 | 15 |

## 7. Change of $\boldsymbol{p}_{s}$ for fixed $\left(\boldsymbol{p}_{r}, \boldsymbol{p}_{1}, \boldsymbol{p}_{2}, w_{2}\right)$

The master tables contain sampling plans for $p_{s}=p_{r}$ only, because a simple and rather accurate rule exists for deriving plans for $p_{s} \neq p_{r}$ from the tabulated ones.

From (69) and (72) it will be seen that $p_{s}$ influences $N_{c}$ only through $\delta$. Writing

$$
p_{s}-p_{m}=p_{s}-p_{r}+w_{1}\left(p_{r}-p_{1}\right)=w_{1}\left(p_{r}-p_{1}\right)\left(1+\frac{p_{s}-p_{r}}{w_{1}\left(p_{r}-p_{1}\right)}\right)
$$

it follows from (72) that

$$
\log N_{c}\left(p_{r}, p_{s}\right)=\log N_{c}\left(p_{r}, p_{r}\right)+\log \left(1+\frac{p_{s}-p_{r}}{w_{1}\left(p_{r}-p_{1}\right)}\right)
$$

or

$$
\begin{equation*}
N_{c}\left(p_{r}, p_{s}\right)=N_{c}\left(p_{r}, p_{r}\right) / \lambda_{s}, \tag{84}
\end{equation*}
$$

say, where

$$
\begin{equation*}
\lambda_{s}=\left(1+\frac{p_{s}-p_{r}}{w_{1}\left(p_{r}-p_{1}\right)}\right)^{-1} . \tag{85}
\end{equation*}
$$

Table 4.
Comparisons of exact sampling plans for $p_{r}=0.010, p_{s}=0.020, p_{1}=0.006$, $p_{2}=0.040, w_{2}=0.05$ with approximate plans derived from the master table.

| $N$ | Exact |  | Approximation |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: |
|  | $n$ | c | $N^{*}=0.275 N$ | $n^{*}$ | $c^{*}$ |
| 300 | Accept |  | 83 | 5 | 0 |
| 500 | 5 | 0 | 138 | 10 | 0 |
| 700 | 10 | 0 | 193 | 15 | 0 |
| 1000 | 15 | 0 | 275 | 15 | 0 |
| 2000 | 60 | 1 | 550 | 60 | 1 |
| 3000 | 110 | 2 | 825 | 110 | 2 |
| 5000 | 120 | 2 | 1380 | 120 | 2 |
| 7000 | 170 | 3 | 1930 | 170 | 3 |
| 10000 | 220 | 4 | 2750 | 220 | 4 |
| 20000 | 280 | 5 | 5500 | 280 | 5 |
| 30000 | 330 | 6 | 8250 | 330 | 6 |
| 50000 | 390 | 7 | 13800 | 390 | 7 |
| 70000 | 395 | 7 | 19300 | 395 | 7 |
| 100000 | 450 | 8 | 27500 | 445 | 8 |
| 200000 | 505 | 9 | 55000 | 550 | 10 |

We have thus proved the following theorem:
The optimum sampling plan corresponding to ( $N, p_{r}, p_{s}, p_{1}, p_{2}, w_{2}$ ) is approximately equal to the plan ( $n^{*}, c^{*}$ ) corresponding to $\left(N^{*}, p_{r}, p_{r}, p_{1}, p_{2}, w_{2}\right)$ with $N^{*}=N \lambda_{s}$.

This theorem makes it possible to use the master tables also for $p_{s} \neq p_{r}$ if only $N$ is replaced by $N^{*}$. The error in $c$ by using this procedure will seldom be more than $\pm 1$. An example has been given in Table 4 with

$$
\lambda_{s}=\left(1+\frac{2-1}{0.95(1-0.6)}\right)^{-1}=0.275
$$

The corresponding "inverse" formula becomes

$$
\begin{equation*}
n_{N}\left(p_{r}, p_{s}\right)=n_{N}\left(p_{r}, p_{r}\right)+\frac{1}{\varphi}\left(1-\frac{1}{3 \log N}\right) \log \lambda_{s} . \tag{86}
\end{equation*}
$$

Using this result for $N=50,000$ and the parameters given in Table 4 we find

$$
n=505-293 \times 0.9291 \times 0.5607=350
$$

as compared to the exact solution 390 .
In the following sections we shall limit ourselves to consider cases with $p_{s}=p_{r}$ since we may always begin the analysis by replacing $N$ by $N^{*}$ if $p_{s} \neq p_{r}$. The "conversion factor" $\lambda_{s}$ depends on $w_{2}$ and the ratios $\left(\varrho_{s}, \varrho_{1}\right)$, i. e. $\lambda_{s}$ is independent of $p_{2}$ and the general quality level.

## 8. Proportional change of $\left(p_{r}, p_{1}, p_{2}\right)$ and change of $w_{2}$

Consider the problem of finding the optimum plans for an arbitrary set of parameter values ( $p_{r}, p_{1}, p_{2}, w_{2}$ ) by combining the proportionality law with the relation between $p_{r}$ and $w_{2}$ for given $\gamma_{2}$ and using the tabulated plans in the master table for parameter values $\left(p_{r 0}, p_{10}, p_{20}, w_{20}\right)$, say.

The problem is to determine $\lambda$ so that $\left(p_{r} ; p_{1}, p_{2}\right)=\left(\lambda p_{r 0}^{*}, \lambda p_{10}, \lambda p_{20}\right)$ and $\left(p_{r 0}^{*}\right.$, $\left.p_{10}, p_{20}, w_{2}\right)$ give the same value of $\gamma_{2}$ as $\left(p_{r 0}, p_{10}, p_{20}, w_{20}\right)$. For this value of $\lambda$ we may find the plans for ( $p_{r}, p_{1}, p_{2}, w_{2}$ ) from the plans for ( $p_{r 0}, p_{10}, p_{20}, w_{2}$ ) by means of the proportionality law, and the plans for ( $p_{r 0}^{*}, p_{10}, p_{20}, w_{2}$ ) are identical to the plans for $\left(p_{r 0}, p_{10}, p_{20}, w_{20}\right)$. (It will be noted that $p_{r 0}^{* /} / p_{r 0}$ is identical to the function defined by (46)).

Since the value of $\gamma_{2}$ is the same for $\left(p_{r}, p_{1}, p_{2}, w_{2}\right)$ and ( $p_{r 0}^{*}, p_{10}, p_{20}, w_{2}$ ) we have the equation $\gamma_{20}=\gamma_{2}$ for the determination of $\lambda$, i.e.

$$
\frac{w_{20}\left(p_{20}-p_{r 0}\right)}{w_{10}\left(p_{r 0}-p_{10}\right)}=\frac{w_{2}\left(p_{2}-p_{r}\right)}{w_{1}\left(p_{r}-p_{1}\right)} .
$$

Introducing $p_{20}=p_{2} / \lambda$ and $p_{10}=p_{1} / \lambda$ we find

$$
\begin{equation*}
\lambda p_{r 0}=\left(p_{2}+\frac{w_{10}}{w_{20}} \gamma_{2} p_{1}\right) /\left(1+\frac{w_{10}}{w_{20}} \gamma_{2}\right) . \tag{87}
\end{equation*}
$$

For the master table with $p_{r 0}=0.01$ and $w_{20}=0.05$ the result is

$$
\begin{equation*}
\lambda=100\left(p_{2}+19 \gamma_{2} p_{1}\right) /\left(1+19 \gamma_{2}\right) \tag{88}
\end{equation*}
$$

For the other master table $\left(p_{r 0}=0.10\right)$ the factor 100 should be replaced by 10 .
The results of sections $6-8$ may be combined to the following theorem:
The optimum sampling plan corresponding to $\left(N, p_{r}, p_{s}, p_{1}, p_{2}, w_{2}\right)$ is approximately equal to $\left(n^{*} / \lambda, c^{*}\right)$ where $\left(n^{*}, c^{*}\right)$ may be found in the master table for $N^{*}=N \lambda_{s} \lambda$, $p_{10}=p_{1} / \lambda$, and $p_{20}=p_{2} / \lambda$, the conversion factors being equal to

$$
\lambda_{s}=\left(1+\frac{p_{s}-p_{r}}{w_{1}\left(p_{r}-p_{1}\right)}\right)^{-1}
$$

and

$$
\lambda=100\left(p_{2}+19 \gamma_{2} p_{1}\right) /\left(1+19 \gamma_{2}\right), \quad \gamma_{2}=\frac{w_{2}\left(p_{2}-p_{r}\right)}{w_{1}\left(p_{r}-p_{1}\right)}
$$

for the 0.01-table, 100 being replaced by 10 for the 0.10-table.
By means of this theorem it is rather easy to find the optimum plan corresponding to an arbitrary set of parameter values if only $p_{1} / \lambda$ and $p_{2} / \lambda$ fall within the range of arguments in the master tables. If that is not the case the method given in the next section may be used.

Usually $p_{1} / \lambda$ and $p_{2} / \lambda$ will not be equal to the arguments used in the master tables. One might then interpolate but this is hardly worth while since the arguments in the table have been chosen in such a way that by rounding to the nearest argument the rounding error will ordinarily be less than $10 \%$.

If one wants to be sure to obtain a sufficiently large sample the value of $p_{1} / \lambda$ should be rounded up and the value of $p_{2} / \lambda$ rounded down.

As an example consider the problem of finding the sampling plans for $\left(p_{r}, p_{1}, p_{2}\right.$, $\left.w_{2}\right)=(0.03,0.01,0.07,0.08)$ and $p_{s}=p_{r}$. Since $p_{r}<0.05$, say, we choose to use the 0.01 -table. From

$$
\gamma_{2}=\frac{8}{92} \frac{7-3}{3-1}=0.174, \quad 19 \gamma_{2}=3.31
$$

we find $\lambda=(7+3.31) /(1+3.31)=2.39, \quad p_{1} / \lambda=0.01 / 2.39=0.0042, \quad$ and $\quad p_{2} / \lambda$ $=0.07 / 2.39=0.029$. The master table should thus be entered with $p_{10}=0.004$ and $p_{20}=0.030$. For $N=2000$, say, we find $N^{*}=4780$ and $\left(n^{*}, c^{*}\right)=(210,3)$ leading to $(n, c)=(210 / 2.39,3)=(90,3)$ which is the correct solution.

If $w_{2}=0.02$ instead of 0.08 we find similarly $\lambda=4.38, p_{1} / \lambda=0.0023 \simeq 0.0025$ and $p_{2} / \lambda=0.0160 \cong 0.0150$. For $N=2000$ we get $N^{*}=8760$ leading to acceptance without inspection as the most economical decision.

## 9. Change of $\boldsymbol{w}_{\mathbf{2}}$ for fixed $\left(\boldsymbol{p}_{\boldsymbol{r}}, \boldsymbol{p}_{s}, \boldsymbol{p}_{1}, \boldsymbol{p}_{2}\right)$

In the following we shall develop a method for evaluating the effect of changing one of the five parameters only, and use it first for $w_{2}$ and then for $p_{r}$.

From (69) and (72) we find for given $c$

$$
\begin{equation*}
\frac{\partial n_{c}}{\partial \log w_{2}}=\frac{\partial \alpha}{\partial \log w_{2}}=1 /\left(w_{1} \log \frac{q_{1}}{q_{2}}\right) \tag{89}
\end{equation*}
$$

and

$$
\begin{equation*}
\frac{\partial \log N_{c}}{\partial \log w_{2}}=\varphi \frac{\partial n_{c}}{\partial \log w_{2}}+\frac{1}{2} \frac{\partial \log n_{c}}{\partial \log w_{2}}+\frac{\partial \delta}{\partial \log w_{2}} . \tag{90}
\end{equation*}
$$

The last term on the right hand side is a rather complicated function of the parameters. Tabulation of $\delta$ and graphical analysis of $\delta$ as a function of $\log w_{2}$ has shown, however, that a least for $w_{2} \leqq 0.20$ and $p_{r} \leqq 0.10$ (and corresponding values of $\left.p_{s}, p_{1}, p_{2}\right) \delta$ is approximately a linear function of $\log w_{2}$ with a slope depending slightly on $\left(\varrho_{s}, \varrho_{1}, \varrho_{2}\right)$ and being practically independent of $p_{r}$.

Limiting ourselves to the case $p_{s}=p_{r}$ we thus have

$$
\frac{\partial \delta}{\partial \log w_{2}} \simeq-b_{1}\left(\varrho_{1}, \varrho_{2}\right)
$$

say, where $b_{1}\left(\varrho_{1}, \varrho_{2}\right)$ has been tabulated in the appendix.

Writing $\delta=\delta\left(p_{r}, \varrho_{1}, \varrho_{2}, w_{2}\right)$ and putting $w_{2}=0.02$ and 0.20 respectively, so that $\Delta \log w_{2}=\log 0.20-\log 0.02=1$, an approximation to $\partial \delta / \partial \log w_{2}$ may be found as $\delta\left(p_{r}, \varrho_{1}, \varrho_{2}, 0.20\right)-\delta\left(p_{r}, \varrho_{1}, \varrho_{2}, 0.02\right)$. This approximation has been computed for both $p_{r}=0.01$ and 0.10 , and finally the average of the two has been taken as $-b_{1}$.

For small $p_{r}$ we also have

$$
\varphi / \log \frac{q_{1}}{q_{2}} \simeq \varphi^{*} /\left(\varrho_{2}-\varrho_{1}\right)=b_{2}(\varrho) .
$$

The values given for $b_{2}$ have been computed as averages of $\varphi /\left(\log \frac{q_{1}}{q_{2}}\right)$ for $p_{r}=0.01$
$p_{r}=0.10$. and $p_{r}=0.10$.

For large $n$ we have that $(\log e) / 2 n$ is small as compared to $\varphi$ and we shall therefore disregard the second term on the right hand side of (90). We then have approximately

$$
\begin{gathered}
\partial \log N_{c} \\
\partial \log w_{2}
\end{gathered}=\begin{gathered}
b_{2}(\varrho) \\
w_{1}
\end{gathered}-b_{1}\left(\varrho_{1}, \varrho_{2}\right)
$$

which gives

$$
N_{c}\left(w_{2}\right)=A w_{2}^{-b_{1}}\left(\frac{w_{1}}{w_{2}}\right)^{-b_{3}}
$$

where $A$ denotes a constant of integration. Changing from $w_{2}$ to $\lambda w_{2}$ we get

$$
\begin{equation*}
N_{c}\left(\lambda w_{2}\right)=N_{c}\left(w_{2}\right) / f_{1}(\lambda) \tag{91}
\end{equation*}
$$

where

$$
\begin{equation*}
f_{1}(\lambda)=\lambda^{b_{1}-b_{2}}\left(1-(\lambda-1) \frac{w_{2}}{w_{1}}\right)^{b_{2}} . \tag{92}
\end{equation*}
$$

From (69) we further have

$$
\begin{equation*}
n_{c}\left(\lambda w_{2}\right)=n_{c}\left(w_{2}\right)+g_{1}(\lambda) \tag{93}
\end{equation*}
$$

where

$$
\begin{equation*}
g_{1}(\lambda)=\left(\log \frac{\lambda w_{1}}{1-\lambda w_{2}}\right) /\left(\log \frac{q_{1}}{q_{2}}\right) \sim\left(\ln \frac{\lambda w_{1}}{1-\lambda w_{2}}\right) /\left(\varrho_{2}-\varrho_{1}\right) p_{r} \tag{94}
\end{equation*}
$$

For convenience $f_{1}$ and $g_{1}$ have been written as functions of $\lambda$ only, even if they both depend also on other parameters. The function $f_{1}$ which will be called the conversion factor for $N$ due to a change in $w_{2}$ has been tabulated in the appendix for $w_{2}$ $=0.05$ as a function of $\left(\lambda, \varrho_{1}, \varrho_{2}\right)$. The function $g_{1}$ which gives the correction to $n$ due to a change in $w_{2}$ has similarly been tabulated in the appendix as function of $\left(\lambda, \varrho_{1}, \varrho_{2}\right)$ for $w_{2}=0.05$ and $p_{r}=0.01$. Values of this function for other values of $p_{r}$ may be obtained as $g_{1} / 100 p_{r}$ where $g_{1}$ represents the tabulated values.


Fig. 7. Relation between lot size and acceptance number by change of $w_{2}$ for fixed ( $p_{r}, p_{s}, p_{1}, p_{2}$ ).

The above results may be formulated as the following theorem:
The optimum sampling plan corresponding to $\left(N, p_{r}, p_{s}, p_{1}, p_{2}, \lambda w_{2}\right), p_{r}=p_{s}$, is approximately equal to $\left(n^{*}+g_{1}(\lambda), c^{*}\right)$ where $\left(n^{*}, c^{*}\right)$ is the plan corresponding to $\left(N^{*}, p_{r}, p_{s}, p_{1}, p_{2}, w_{2}\right)$ with $N^{*}=N f_{1}(\lambda)$.

The theorem has been illustrated in Fig. 7.
This theorem enlarges the field of application of the two master tables with respect to values of $w_{2}$ in a similar manner as the law of proportionality does with respect to the other parameters. The results of using the approximation have been compared with the exact solutions in a large number of cases and the deviations found between the approximate and the correct value of $c$ have never exceeded 1 for $\lambda<4$. There is a tendency for the approximation to give too small a value of $c$ for $\lambda>1$ and too large a value for $\lambda<1$, in particular for small $N$.

It should be noted that the formula breaks down in some cases for small $N$. Let $N_{a}$ denote the largest $N$ for which acceptance without inspection is cheaper than sampling inspection for the master table used. If $\lambda>1$ and $N f_{1}(\lambda)=N^{*}<N_{a}$ then the formula does not lead to a sampling plan even if there may exist a plan which for $\lambda w_{2}$ is cheaper than acceptance without inspection. Similarly, for $\lambda<1$ and $N f_{1}(\lambda)=N^{*}>N_{a}$ there may be some cases where the approximation formula leads to a sampling plan even if the cheapest solution is acceptance without inspection.

An example has been shown in Table 5. The approximation is remarkably good. Since $N_{a}=74$ the approximation formula leads to acceptance without inspection for all $N \leqq 57$. Sampling plans cheaper than acceptance without inspection do, however, exist for $12 \leqq N \leqq 57$.

## Table 5.

Comparisons of exact sampling plans for $p_{r}=p_{s}=0.010, p_{1}=0.006, p_{2}=0.040$, $w_{2}=0.10$, and approximate plans derived from the master table. $f_{1}=1.29, g_{1}=20$.

| $N$ | Exact |  | Approximate |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: |
|  | $n$ | c | $N^{*}=1.29 \mathrm{~N}$ | $n^{*}+20$ | $c^{*}$ |
| 50 | 15 | 0 | 65 |  |  |
| 70 | 20 | 0 | 90 | 25 | 0 |
| 100 | 25 | 0 | 129 | 30 | 0 |
| 200 | 35 | 0 | 258 | 35 | 0 |
| 300 | 75 | 1 | 387 | 75 | 1 |
| 500 | 85 | 1 | 645 | 85 | 1 |
| 700 | 130 | 2 | 903 | 130 | 2 |
| 1000 | 140 | 2 | 1290 | 140 | 2 |
| 2000 | 240 | 4 | 2580 | 240 | 4 |
| 3000 | 250 | 4 | 3870 | 250 | 4 |
| 5000 | 305 | 5 | 6450 | 300 | 5 |
| 7000 | 355 | 6 | 9030 | 355 | 6 |
| 10000 | 405 | 7 | 12900 | 405 | 7 |
| 20000 | 465 | 8 | 25800 | 465 | 8 |
| 30000 | 520 | 9 | 38700 | 520 | 9 |
| 50000 | 575 | 10 | 64500 | 575 | 10 |
| 70000 | 630 | 11 | 90300 | 630 | 11 |
| 100000 | 635 | 11 | 129000 | 635 | 11 |
| 200000 | 740 | 13 | - | - | - |

Using the method of section 8 we find $\gamma_{2}=0.833, \lambda=0.804, p_{1} / \lambda=0.0075$, and $p_{2} / \lambda=0.050$. Since $p_{1} / \lambda$ falls outside the range of arguments in the master table the method does not apply. Using $p_{1} / \lambda=0.007$ gives, however, a rather good approximation.

From the inverse formula (79) we get

$$
\begin{gathered}
\partial n_{N} \\
\partial \log w_{2}
\end{gathered}=\frac{b_{1}\left(\varrho_{1}, \varrho_{2}\right)}{\varphi}\left(1-\frac{1}{3 \log N}\right)
$$

and consequently

$$
n_{N}\left(w_{2}\right)=A+\frac{b_{1}}{\varphi}\left(1-\frac{1}{3 \log N}\right) \log w_{2}
$$

or

$$
\begin{equation*}
n_{N}\left(\lambda w_{2}\right)=n_{N}\left(w_{2}\right)+\frac{b_{1}}{\varphi}\left(1-\frac{1}{3 \log N}\right) \log \lambda . \tag{95}
\end{equation*}
$$

This shows that the difference between $n_{N}\left(\lambda w_{2}\right)$ and $n_{N}\left(w_{2}\right)$ for given $N$ is proportional to $\log \lambda$. This formula is, however, not as accurate as (93) for small $N$.

An example has been given in the following table for $N=20,000, p_{r}=p_{s}$ $=0.010, p_{1}=0.006, \quad p_{2}=0.020$, and $w_{2}=0.05$, which gives $b_{1}=0.59$ and $1 / \varphi=1100$.

Comparisons of exact and approximate sampling plans derived from (95).

|  | Exact |  | Approx. |  |  |
| :---: | :---: | ---: | ---: | ---: | ---: |
| $100 w_{2}$ | $\lambda$ | $n$ | $c$ | $n$ | $c$ |
| 2.5 | 0.5 | 540 | 8 | 580 | 9 |
| 5.0 | 1.0 | 760 | 10 | - | - |
| 10.0 | 2.0 | 980 | 12 | 940 | 11 |
| 20.0 | 4.0 | 1130 | 13 | 1120 | 13 |

## 10. Change of $\boldsymbol{p}_{r}=\boldsymbol{p}_{s}$ for fixed $\left(\boldsymbol{p}_{1}, \boldsymbol{p}_{2}, w_{2}\right)$

From (69) and (72) we find for given $c$

$$
\begin{equation*}
\frac{\partial n_{c}}{\partial \log p_{r}}=\frac{\partial \alpha}{\partial \log p_{r}}=-p_{r}\left(\frac{1}{p_{r}-p_{1}}+\frac{1}{p_{2}-p_{r}}\right) /\left(\log q_{1}\right) \tag{96}
\end{equation*}
$$

and

$$
\begin{equation*}
\frac{\partial \log N_{c}}{\partial \log p_{r}}=\varphi \frac{\partial n_{c}}{\partial \log p_{r}}+\frac{1}{2} \frac{\partial \log n_{c}}{\partial \log p_{r}}+\frac{\partial \delta}{\partial \log p_{r}} . \tag{97}
\end{equation*}
$$

Numerical investigations show that-for $w_{2}<0.20, p_{r}<0.20$, and $p_{r}=p_{s}-$ $\delta$ is approximately a linear function of $\log p_{r}$ with a slope depending on $\varrho$ and being practically independent of "the level of $\left(p_{1}, p_{2}\right)$ " and of $w_{2}$ if only $p_{r}$ does not come too close to $p_{1}$ or $p_{2}$, i.e.

$$
\frac{\partial \delta}{\partial \log p_{r}} \simeq b_{3}(\varrho) \text { for } p_{1} \varrho^{1 / 5}<p_{r}<p_{2} \varrho^{-1 / 5},
$$

say, where $b_{3}(\varrho)$ has been tabulated. (Another limitation of no practical importance is that $p_{2}$ must not be too close to 1 ). An approximation to $\partial \delta / \partial \log p_{r}$ may be computed as the corresponding difference-quotient setting $p_{r}=p_{1} \varrho^{1 / 5}$ and $p_{r}=p_{2} \varrho^{-1 / 5}$ respectively. This has been done for $w_{2}=0.05$ and for the "standard" values of $p_{1}$ and $p_{2}$, partly at the $1 \%$ and partly at the $10 \%$ level. The value of $b_{3}$ given in the table is the average of the two values found.

Proceeding as in section 9 we have approximately

$$
\frac{\partial \log N_{c}}{\partial \log p_{r}}=-b_{2}(\varrho) p_{r}\left(\frac{1}{p_{r}-p_{1}}+\frac{1}{p_{2}-p_{r}}\right)+b_{3}(\varrho)
$$

which on integration gives

Table 6.
Comparisons of exact sampling plans for $p_{r}=p_{s}=0.020, p_{1}=0.006, p_{2}=0.040$, $w_{2}=0.05$ with approximate plans derived from the master table. $\lambda=2, f_{2}=0.52$,

| $N$ | Exact |  | Approximate |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: |
|  | $n$ | c | $N^{*}=0.52 \mathrm{~N}$ | $n^{*}-55$ | $c^{*}$ |
| 2000 | Accept |  | 1040 | 60 | 2 |
| 3000 | 75 | 2 | 1560 | 110 | 3 |
| 5000 | 130 | 3 | 2600 | 165 | 4 |
| 7000 | 180 | 4 | 3640 | 170 | 4 |
| 10000 | 230 | 5 | 5200 | 225 | 5 |
| 20000 | 290 | 6 | 10400 | 280 | 6 |
| 30000 | 345 | 7 | 15600 | 335 | 7 |
| 50000 | 400 | 8 | 26000 | 390 | 8 |
| 70000 | 450 | 9 | 36400 | 445 | 9 |
| 100000 | 460 | 9 | 52000 | 450 | 9 |
| 200000 | 565 | 11 | 104000 | 555 | 11 |

$$
N_{c}\left(p_{r}\right)=A p_{r}^{b_{3}}\left(\frac{p_{2}-p_{r}}{p_{r}-p_{1}}\right)^{b_{2}}
$$

and

$$
\begin{equation*}
N_{c}\left(\lambda p_{r}\right)=N_{c}\left(p_{r}\right) / f_{2}(\lambda) \tag{98}
\end{equation*}
$$

where

$$
\begin{equation*}
f_{2}(\lambda)=\lambda^{-b_{3}}\left(\frac{\left(\varrho_{2}-1\right)\left(\lambda-\varrho_{1}\right)}{\left(1-\varrho_{1}\right)\left(\varrho_{2}-\lambda\right)}\right)^{b_{2}} . \tag{99}
\end{equation*}
$$

From (69) we further have

$$
\begin{equation*}
n_{c}\left(\lambda p_{r}\right)=n_{c}\left(p_{r}\right)+g_{2}(\lambda) \tag{100}
\end{equation*}
$$

where

$$
\begin{equation*}
g_{2}(\lambda) \simeq\left(\ln \frac{\left(1-\varrho_{1}\right)\left(\varrho_{2}-\lambda\right)}{\left(\varrho_{2}-1\right)\left(\lambda-\varrho_{1}\right)}\right) /\left(\varrho_{2}-\varrho_{1}\right) p_{r} . \tag{101}
\end{equation*}
$$

The conversion factor for $N$ due to a change in $p_{r}, f_{2}(\lambda)$, has been tabulated in the appendix as function of $\left(\lambda, \varrho_{1}, \varrho_{2}\right)$, and the correction to $n$ due to a change in $p_{r}, g_{2}(\lambda)$, has been tabulated as function of $\left(\lambda, \varrho_{1}, \varrho_{2}\right)$ for $p_{r}=0.01$. Values of $g_{2}(\lambda)$ for other values of $p_{r}$ may be found from the tabulated ones by dividing by $100 p_{r}$.

The above results may be formulated as the following theorem:
The optimum sampling plan corresponding to $\left(N, \lambda p_{r}, \lambda p_{s}, p_{1}, p_{2}, w_{2}\right), p_{r}=p_{s}$, is approximately equal to $\left(n^{*}+g_{2}(\lambda), c^{*}\right)$ where $\left(n^{*}, c^{*}\right)$ is the plan corresponding to $\left(N^{*}, p_{r}, p_{s}, p_{1}, p_{2}, w_{2}\right)$ with $N^{*}=N f_{2}(\lambda)$.

With the given set of tables this theorem is, however, not as important in practice as the previous ones, because the tables contain the optimum plans for so many
combinations of $\left(p_{r}, p_{1}, p_{2}\right)$ that an adjustment of the relative position of $p_{r}$ within the interval $\left(p_{1}, p_{2}\right)$ will seldom be felt necessary from a practical point of view.

In table 6 an example has been shown of the effect of changing $p_{r}=p_{s}$ from 0.010 to 0.020 within the interval $\left(p_{1}, p_{2}\right)=(0.006,0.040)$.

From the inverse formula (79) we get

$$
\frac{\partial n_{N}}{\partial \log p_{r}}=-\frac{b_{3}(\varrho)}{\varphi}\left(1-\frac{1}{3 \log N}\right)
$$

which leads to

$$
\begin{equation*}
n_{N}\left(\lambda p_{r}\right)=n_{N}\left(p_{r}\right)-\frac{b_{3}}{\varphi}\left(1-\frac{1}{3 \log N}\right) \log \lambda \tag{102}
\end{equation*}
$$

An example of the application of this formula has been given in the following table for $N=50,000, p_{r}=p_{s}=0.010, \lambda=0.5$ and $2.5, p_{1}=0.002, p_{2}=0.040$, and $w_{2}=0.05$, which give $b_{3}=1.09$ and $1 / \varphi=177$.

Comparisons of exact and approximate sampling plans derived from (102).

|  | Exact |  | Approximate |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: |
| $p_{r}$ | $\lambda$ | $n$ | $c$ | $n$ | $c$ |
| 0.005 | 0.5 | 350 | 4 | 370 | 4 |
| 0.010 | 1.0 | 315 | 4 | - | - |
| 0.025 | 2.5 | 250 | 4 | 245 | 4 |

## 11. Change of all parameters

The results of the preceding sections may be combined into a "chain formula" of the type

$$
\begin{equation*}
N_{c}\left(\lambda p_{r}, \varrho_{s} \lambda p_{r}, \lambda p_{1}, \lambda p_{2}, \lambda_{1} w_{2}\right)=N_{c}\left(p_{r}, p_{r}, p_{1}, p_{2}, w_{2}\right) / \lambda_{s} f_{1} \lambda \tag{103}
\end{equation*}
$$

and

$$
\begin{equation*}
n_{c}\left(\lambda p_{r}, \varrho_{s} \lambda p_{r}, \lambda p_{1}, \lambda p_{2}, \lambda_{1} w_{2}\right)=\left(n_{c}\left(p_{r}, p_{r}, p_{1}, p_{2}, w_{2}\right)+g_{1}\right) / \lambda \tag{104}
\end{equation*}
$$

where

$$
\lambda_{s}=\left(1+\frac{\varrho_{s}-1}{\left(1-\lambda_{1} w_{2}\right)\left(1-\varrho_{1}\right)}\right)^{-1}
$$

$f_{1}\left(\lambda_{1}\right)$ and $g_{1}\left(\lambda_{1}\right)$ being defined by (92) and (94) for $\varrho_{1}=p_{1} / p_{r}$ and $\varrho_{2}=p_{2} / p_{r}$.
In the master tables $p_{r}=p_{s}=0.01$ (or 0.10 ) and $w_{2}=0.05$ have been used as reference values. What has been denoted by $\lambda$ and $\lambda_{1}$ in the above formulas become $100 p_{r}\left(\right.$ or $\left.10 p_{r}\right)$ and $20 w_{2}$ if $p_{r}$ and $w_{2}$ denote the values for which the optimum plan is required.

Table 7.
Comparisons of exact sampling plans for $p_{r}=0.030, p_{s}=0.060, p_{1}=0.018$, $p_{2}=0.120, w_{2}=0.10$ and approximate plans derived from the master table for $p_{r}=p_{s}=0.010, p_{1}=0.006, p_{2}=0.040, w_{2}=0.05$.

| N | Exact |  | Approximate |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: |
| 50 |  |  | 51 |  |  |
| 70 | 5 | 0 | 71 |  |  |
| 100 | 5 | 0 | 102 | 10 | 0 |
| 200 | 10 | 0 | 204 | 10 | 0 |
| 300 | 10 | 0 | 306 | 25 | 1 |
| 500 | 25 | 1 | 510 | 25 | 1 |
| 700 | 30 | 1 | 714 | 30 | 1 |
| 1000 | 45 | 2 | 1020 | 45 | 2 |
| 2000 | 65 | 3 | 2040 | 65 | 3 |
| 3000 | 80 | 4 | 3060 | 80 | 4 |
| 5000 | 100 | 5 | 5100 | 100 | 5 |
| 7000 | 100 | 5 | 7140 | 100 | 5 |
| 10000 | 120 | 6 | 10200 | 120 | 6 |
| 20000 | 135 | 7 | 20400 | 155 | 8 |
| 30000 | 155 | 8 | 30600 | 155 | 8 |
| 50000 | 175 | 9 | 51000 | 175 | 9 |
| 70000 | 190 | 10 | 71400 | 195 | 10 |
| 100000 | 210 | 11 | 102000 | 210 | 11 |
| 200000 | 225 | 12 | 204000 | 230 | 12 |

We thus get the following rule for using the master table with $p_{r}=0.01$ :
The optimum plan for $\left(N, p_{r}, p_{s}, p_{1}, p_{2}, w_{2}\right)$ with $p_{r}<0.05$ is approximately equal to $\left(\left(n^{*}+g_{1}\right) / 100 p_{r}, c^{*}\right)$ where $\left(n^{*}, c^{*}\right)$ may be found by entering the master table with

$$
\begin{equation*}
N^{*}=N\left(100 p_{r}\right) f_{1}\left(20 w_{2}, \varrho_{1}, \varrho_{2}\right) /\left(1+\frac{p_{s}-p_{r}}{w_{1}\left(p_{r}-p_{1}\right)}\right) \tag{105}
\end{equation*}
$$

$\varrho_{1}=p_{1} / p_{r}, \varrho_{2}=p_{2} / p_{r}$, and $g_{1}=g_{1}\left(20 w_{2}, \varrho_{1}, \varrho_{2}\right)$, the arguments for $\left(p_{1}, p_{2}\right)$ in the master table being $\left(\varrho_{1} / 100, \varrho_{2} / 100\right)$.

For $0.05<p_{r}<0.20$ the master table with $p_{r}=0.10$ should be used accordingly.

If $\left(\varrho_{1} / 100, \varrho_{2} / 100\right)$ are not to be found in the table then use the "nearest" argument or interpolate. One may also use the results in section 10 to change $p_{r}$ in the master table so that the relations between $\left(p_{r}, p_{1}, p_{2}\right)$ in the table become closer to the ones for which the sampling plan is required. From a practical point of view, however, the master tables combined with the rule above will normally suffice.

An example has been given in Table 7. The conversion factor for $N$ is found as

$$
3 f_{1}(2,0.6,4.0)\left(\left(1+\frac{30}{0.90 \times 12}\right)=3 \times 1.29 / 3.78=1.02 .\right.
$$

The agreement between the approximate and the exact solution is very good.
Using instead the method of section 8 we get $\lambda_{s}=1 / 3.78=0.265, \lambda=2.40$, $p_{1} / \lambda=0.0075$, and $p_{2} / \lambda=0.050$, i.e. $N^{*}=0.636 N$ and $n=n^{*} / 2.40$. Since the master table does not contain the argument 0.0075 we may as an approximation use 0.0070 which, however, will tend to give too small values of $c$.

The corresponding inverse formula takes the form

$$
\begin{gather*}
n_{N}\left(p_{r}, p_{s}, p_{1}, p_{2}, w_{2}\right)=\frac{1}{100 p_{r}}\left[n_{0}+\frac{1}{\varphi}\left\{\log \left(100 p_{r}\right)+b_{1} \log \left(20 w_{2}\right)\right.\right. \\
\left.\left.-\log \left(1+\frac{p_{s}-p_{r}}{w_{1}\left(p_{r}-p_{1}\right)}\right)\right\}\left(1-\frac{1}{3 \log N}\right)\right] \tag{106}
\end{gather*}
$$

where $n_{0}$ denotes the sample size to be found in the master table for $p_{r}=0.01$, $\varrho_{1}=p_{1} / p_{r}, \varrho_{2}=p_{2} / p_{r}$, corresponding to the given lot size $N$.

As an example consider the determination of $n$ for $N=50,000$ and the parameters given in Table 7. The value of $1 / \varphi$ is 293 for $p_{1}=0.006$ and $p_{2}=0.040$, and

$$
\begin{aligned}
& \left(1-\frac{1}{3 \log N}\right)=0.929, \text { so that we find } \\
& \qquad \begin{aligned}
n & =\frac{1}{3}(505+293(\log 3+0.61 \log 2-\log 3.78) 0.929) \\
& =\frac{1}{3}(505+23)=176
\end{aligned}
\end{aligned}
$$

in agreement with the (rounded) exact solution, $n=175$, given in Table 7.

## 12. Efficiency

In a previous paper [6] it has been proposed to define the efficiency of a sampling plan as

$$
\begin{equation*}
e(N, n, c)=R_{0}(N) / R(N, n, c) \tag{107}
\end{equation*}
$$

where $R_{0}(N)$ denotes the costs of the optimum plan and $R(N, n, c)$ denotes the costs of the plan in question.

We shall first discuss the efficiency of a sampling plan on the assumption that the optimum relationship between $n$ and $c$ has been used so that the loss in efficiency is due to using a wrong relationship between $N$ and $n$. Looking at Fig. 2 it will be
seen that it does not matter much whether we use the value of $c$ giving the absolute minimum of $R$ or a neighbouring value of $c$ provided $n$ is chosen such that a (relative) minimum of $R$ is obtained.

For a given set of parameters let $\left(n_{0}, c_{0}\right)$ be optimum for $N_{0}$ and $\left(n_{1}, c_{1}\right)$ be optimum for $N_{1}$. From (26) it follows that

$$
R(N, n, c)=n+(N-n) h(n, c)
$$

where

$$
h(n, c)=\gamma_{1} Q\left(p_{1}\right)+\gamma_{2} P\left(p_{2}\right)
$$

Using the plan $\left(n_{1}, c_{1}\right)$ for lot size $N_{0}$ (instead of $N_{1}$ ) we find

$$
\begin{align*}
R\left(N_{0}, n_{1}, c_{1}\right) & =n_{1}+\left(N_{0}-n_{1}\right) h\left(n_{1}, c_{1}\right) \\
& =n_{1}+\left(N_{0}-n_{1}\right)\left(R_{0}\left(N_{1}\right)-n_{1}\right) /\left(N_{1}-n_{1}\right) . \tag{108}
\end{align*}
$$

It is therefore rather simple by means of the function $R_{0}(N)$ to evaluate the efficiency of plans contained in the master table in case such plans are used for the wrong value of $N$. The resulting efficiency is

$$
\begin{equation*}
e\left(N_{0}, n_{1}, c_{1}\right)=\frac{R_{0}\left(N_{0}\right)}{n_{1}+\left(N_{0}-n_{1}\right)\left(R_{0}\left(N_{1}\right)-n_{1}\right) /\left(N_{1}-n_{1}\right)} . \tag{109}
\end{equation*}
$$

Since $R_{0}(N) \sim n+1 / \varphi_{0}$ we have asymptotically

$$
\begin{equation*}
e\left(N_{0}, n_{1}, c_{1}\right) \sim\left(n_{0}+\frac{1}{\varphi_{0}}\right) /\left(n_{1}+\frac{N_{0}-n_{1}}{N_{1}-n_{1}} \frac{1}{\varphi_{0}}\right) . \tag{110}
\end{equation*}
$$

Introducing $n_{0}=\left(\ln N_{0}\right) / \varphi_{0}+o\left(\ln N_{0}\right)$ and considering $n_{1}$ as an arbitrary function of $N_{0}, n_{1}=g\left(N_{0}\right) / \varphi_{0}$ say, we find

$$
\begin{equation*}
e\left(N, n_{1}, c_{1}\right) \sim(\ln N) /\left(g(N)+N e^{-g(N)}\right) \tag{111}
\end{equation*}
$$

for $n_{1}=o(N)$, which is the result given without proof in [6].
For $g(N)=\lambda \ln N$ we get $e \rightarrow 1 / \lambda$ for $\lambda \geqq 1$ but $e \rightarrow 0$ for $0<\lambda<1$, i.e. if we use a semilogarithmic relationship between $n$ and $N$ differing from the correct one then it is important to use too large a sample. For $g(N)=N^{\lambda}, \lambda>0$, we get $e \rightarrow 0$.

A more accurate expression than (111) may be found by using all three term of (61) which leads to

$$
e\left(N_{0}, n_{1}, c_{1}\right) \sim\left(n_{0}+\frac{1}{\varphi_{0}}\right) /\left(n_{1}+\frac{1}{\varphi_{0}}\left(e^{\varphi_{0}\left(n_{0}-n_{1}\right)} \sqrt{\left.\left.\frac{n_{0}}{n_{1}}+\frac{n_{0}-n_{1}}{N_{1}-n_{1}}\right)\right) . . . . ~ . ~}\right.\right.
$$

Table 8.
Investigation of efficiency for sampling plans with an acceptance number deviating 1 from the optimum. ( $e^{*}=$ asymptotic efficiency).

| $N$ | $n_{0}$ | $c_{0}$ | $R$ | $n_{1}$ | $c_{1}$ | $100 e$ | $100 e^{*}$ |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| 145 | 10 | 0 | 53 | 60 | 1 | 72 | 94 |
| 447 | 60 | 1 | 126 | $\begin{array}{r} 10 \\ 115 \end{array}$ | $\begin{aligned} & 0 \\ & 2 \end{aligned}$ | $\begin{aligned} & 84 \\ & 86 \end{aligned}$ | $\begin{aligned} & 94 \\ & 95 \end{aligned}$ |
| 1010 | 115 | 2 | 200 | $\begin{array}{r} 60 \\ 170 \end{array}$ | $\begin{aligned} & 1 \\ & 3 \end{aligned}$ | $\begin{aligned} & 90 \\ & 92 \end{aligned}$ | $\begin{aligned} & 95 \\ & 96 \end{aligned}$ |
| 1900 | 170 | 3 | 265 | $\begin{aligned} & 115 \\ & 225 \end{aligned}$ | $\begin{aligned} & 2 \\ & 4 \end{aligned}$ | $\begin{aligned} & 93 \\ & 95 \end{aligned}$ | $\begin{aligned} & 96 \\ & 97 \end{aligned}$ |
| 3350 | 225 | 4 | 326 | $\begin{aligned} & 170 \\ & 280 \end{aligned}$ | $\begin{aligned} & 3 \\ & 5 \end{aligned}$ | $\begin{aligned} & 95 \\ & 96 \end{aligned}$ | $\begin{aligned} & 96 \\ & 97 \end{aligned}$ |
| 5700 | 280 | 5 | 386 | $\begin{aligned} & 225 \\ & 335 \end{aligned}$ | $\begin{aligned} & 4 \\ & 6 \end{aligned}$ | $\begin{aligned} & 96 \\ & 97 \end{aligned}$ | $\begin{aligned} & 97 \\ & 98 \end{aligned}$ |
| 9530 | 335 | 6 | 444 | $\begin{aligned} & 280 \\ & 390 \end{aligned}$ | $\begin{aligned} & 5 \\ & 7 \end{aligned}$ | $\begin{aligned} & 96 \\ & 97 \end{aligned}$ | $\begin{aligned} & 97 \\ & 98 \end{aligned}$ |
| 15800 | 390 | 7 | 502 | $\begin{aligned} & 335 \\ & 445 \end{aligned}$ | $\begin{aligned} & \hline 6 \\ & 8 \end{aligned}$ | $\begin{aligned} & 97 \\ & 98 \end{aligned}$ | $\begin{aligned} & 97 \\ & 98 \end{aligned}$ |
| 25800 | 445 | 8 | 559 | $\begin{aligned} & 390 \\ & 500 \end{aligned}$ | $\begin{aligned} & \hline 7 \\ & 9 \end{aligned}$ | $\begin{aligned} & 97 \\ & 98 \end{aligned}$ | $\begin{aligned} & 98 \\ & 98 \end{aligned}$ |
| 42100 | 500 | 9 | 616 | $\begin{aligned} & 445 \\ & 555 \end{aligned}$ | $\begin{array}{r} 8 \\ 10 \end{array}$ | $\begin{aligned} & 97 \\ & 98 \end{aligned}$ | $\begin{aligned} & 98 \\ & 98 \end{aligned}$ |
| 68300 | 555 | 10 | 672 | $\begin{aligned} & 500 \\ & 615 \end{aligned}$ | $\begin{array}{r} 9 \\ 11 \end{array}$ | $\begin{aligned} & 98 \\ & 98 \end{aligned}$ | $\begin{aligned} & 98 \\ & 98 \end{aligned}$ |

This formula is, however, not of direct value because it contains $N_{1}$ which is unknown in practice. A simple and practically useful approximation is the following

$$
\begin{equation*}
e\left(N_{0}, n_{1}, c_{1}\right) \sim\left(n_{0}+\frac{1}{\varphi_{0}}\right) /\left(n_{1}+\frac{1}{\varphi_{0}} e^{\varphi_{0}\left(n_{0}-n_{1}\right)}\right) . \tag{112}
\end{equation*}
$$

This formula will, however, give too large efficiencies for small values of $n$ because the decision loss has been overestimated.

In connection with the various approximations developed in the preceding sections is has repeatedly been stated that the value of $c$ found by using the approximations will normally not deviate more than 1 from the correct value (for $(N<200,000)$. It is therefore of importance to know the efficiency of a plan for which $\left|c_{1}-c_{0}\right|=1$.

If $\left|c_{1}-c_{0}\right|=a$ (constant) then $\left|n_{1}-n_{0}\right|=a \beta$ and $e \rightarrow 1$ for $N_{0} \rightarrow \infty$. Expanding the denominator of (112) we find for small values of $\varphi_{0} \alpha \beta$

$$
\begin{equation*}
e\left(N_{0}, n_{1}, c_{1}\right) \sim\left(n_{0}+\frac{1}{\varphi_{0}}\right) /\left(n_{0}+\frac{1}{\varphi_{0}}+\frac{1}{2} \varphi_{0} a^{2} \beta^{2}\right) \tag{113}
\end{equation*}
$$

which converges rather fast to 1 for $n_{0} \rightarrow \infty$ and $a=1$. By means of the results of section 6 it will be seen that this asymptotic efficiency (as a function of $c_{0}$ ) is independent of the "quality level".

An example has been given in Table 8. The costs for each optimum plan have been compared with the costs of using a neighbouring plan, i.e. $c_{1}=c_{0} \pm 1$. The efficiency has been compared with the asymptotic efficiency found from (112). It will be seen that the efficiency is larger than 0.90 for $c \geqq 2$ and that the asymptotic formula gives too high an efficiency for small $N$. ( $N$ has been chosen as the geometric mean of the smallest and largest $N$ for each $c$ ). This conclusion is typical for the cases investigated.

The conversion formulas and tables show how sensitive the solution is to changes of the parameters. A change of $w_{2}$ from 0.05 to 0.10 , say, means, that $N$ has to be multiplied by a factor of about 1.3 and the corresponding $n$ should be increased by about 30. (In most systems of sampling plans in practical use to-day the same plan is used for a rather large $N$-interval, the ratio between endpoints usually being 1.5 or larger). As an example consider the case with $p_{r}=p_{s}=0.010, p_{1}=0.006$ and $p_{2}=0.040$ as shown in the following table.

| Optimum sampling plans. |  |  |  |  |
| ---: | ---: | ---: | ---: | ---: |
| $w_{2}=0.05$ |  |  |  | $w_{2}=0.10$ |
| $N$ | $n$ | $c$ | $n$ | $c$ |
| 500 | 60 | 1 | 85 | 1 |
| 1000 | 115 | 2 | 140 | 2 |
| 5000 | 275 | 5 | 305 | 5 |
| 10000 | 335 | 6 | 405 | 7 |
| 50000 | 505 | 9 | 575 | 10 |
| 100000 | 610 | 11 | 635 | 11 |

For most lot sizes we find the same value of $c$ and a difference in $n$ of about 25 , in other cases the difference in $c$ is 1 and the difference in $n$ correspondingly larger. It is immediately clear that using the plans corresponding to $w_{2}=0.05$ if the true value of $w_{2}$ is 0.10 does not lead to an essential loss in efficiency.

The conclusion is that even if the value of $w_{2}$ used deviates from the true value by a factor of 2 the method will nevertheless lead to a sampling plan of very high efficiency.

Similar conclusions may be drawn for the other parameters by studying the conversion formulas.

The main reason why changes of $p_{r}$ and $w_{2}$ does not affect the optimum solution seriously is that $p_{0}$ and $\varphi_{0}$ are independent of $p_{r}$ and $w_{2}$.


Since the most important relation in the system is

$$
c+\frac{1}{2}=p_{0}(n-\alpha)
$$

it is of importance to know how $p_{0}$ depends on $p_{1}$ and $p_{2}$.
From

$$
\frac{\partial \ln p_{0}}{\partial \ln p_{1}}=\frac{p_{0}-p_{1}}{q_{1} \ln \frac{q_{1}}{q_{2}}}>0 \quad \text { and } \frac{\partial \ln p_{0}}{\partial \ln p_{2}}=\frac{p_{2}-p_{1}}{q_{2} \ln \frac{q_{1}}{q_{2}}>0}
$$

it follows that $p_{0}$ is an increasing function of as well $p_{1}$ as $p_{2}$. Furthermore we have approximately

$$
\frac{\partial \ln p_{0}}{\partial \ln p_{1}}+\frac{\partial \ln p_{0}}{\partial \ln p_{2}} \sim 1 .
$$

Within the domain of variation tabulated the first term is on the average 0.35 and the second 0.65 .

The coefficient $p_{0} \alpha$ varies rather slowly with $\left(p_{1}, p_{2}\right)$.
It follows that $p_{0}$ is known with a relative error of about the same size as the relative errors of $p_{1}$ and $p_{2}$.

If the choice of $p_{1}$ and $p_{2}$ is doubtful then $p_{1}$ should be chosen too large and $p_{2}$ too small (by about half of the percentage error in $p_{1}$ ) because the two errors will tend to counterbalance one another and thus give the correct $p_{0}$. The reason for bringing the two parameters closer together in case of doubt lies also in the fact that $\varphi_{0}$ is a decreasing function of $p_{1}$ and an increasing function of $p_{2}$. Since $n \sim(\log N) / \varphi_{0}$ the proposed rule will lead to a larger sample size than the optimum one which normally gives a better efficiency than too small a sample.

Table 9 shows the efficiency of using a plan obtained by entering the master table by a wrong value of $p_{1}, p_{2}$ or both. It is assumed that the true values of ( $p_{1}, p_{2}$ ) are $(0.006,0.040)$ and optimum plans have been substituted by plans obtained by using neighbouring values of $\left(p_{1}, p_{2}\right)$ in the tables, i. e. the relative error of $p_{1}$ is $17 \%$ and the relative error of $p_{2}$ is $12.5 \%$ downwards and $25 \%$ upwards. The table shows that the efficiency in all cases is larger than $90 \%$ for $N<10,000$. For $N=200,000$, however, the efficiency falls to $58 \%$ in the worst case, i.e. the case where $p_{2}$ is chosen $25 \%$ too large.

The results in the table support the statement above that in case of doubt it is important to use a large value of $p_{1}$ and a small value of $p_{2}$.

A remark on the definition of efficiency. For a lot containing $X$ defectives acceptance without inspection is cheaper than rejection without inspection for $X \leqq\left[N p_{r}\right]$. Classifying all lots in this way the average costs become

$$
K_{N m}=\sum_{X=0}^{\left[N p_{r}\right]}\left(N A_{1}+X A_{2}\right) f_{N}(X)+\sum_{X=\left[N p_{r}\right]+1}^{N}\left(N R_{1}+X R_{2}\right) f_{N}(X) .
$$

It is easily seen that $K_{N m} / N \rightarrow k_{m}$ for $N \rightarrow \infty$, see (15), and that $K_{N m}<N k_{m}=K_{m}$. It would be more correct to define efficiency as the ratio of costs in excess of $K_{N m}$ instead of $K_{m}$ as in (107). This modification will increase the efficiencies for small $N$ slightly whereas the above results regarding asymptotic efficiency will be unchanged. In Table 8 the first 5 efficiencies would be $83,87,89,91$, and 93 , whereas the remaining are unchanged, and in Table 9 the only change would be to increase 5 of the values of $100 e$ for $N=200$ by 1 .

## 13. An example

Consider now an example starting from the original cost functions. To show the various aspects of the method the example will be worked out in more detail than is necessary for routine applications.

Let the three cost functions be $k_{s}(p)=23+35 p, k_{r}(p)=16+35 p, k_{a}(p)=720 p$, the coefficients denoting costs per item in cents, say, i.e. the costs of sampling and testing is 23 cents per item in the sample and the costs of accepting a defective item is 720 cents etc., see section 2.

Let us further assume that lots are generated with probability $w_{1}=0.93$ from a binomially controlled process with $p_{1}=0.009$ and with probability $w_{2}=0.07$ from a process with $p_{2}=0.080$.

The costs may then be described as in the following table:

| $w$ | $p$ | $k_{s}(p)$ | $k_{r}(p)$ | $k_{a}(p)$ | $k_{m}(p)$ | $\left\|k_{r}(p)-k_{a}(p)\right\|$ |
| :--- | :---: | :---: | :---: | :---: | ---: | :---: |
| 0.93 | 0.009 | 23.315 | 16.315 | 6.480 | 6.480 | 9.835 |
| 0.07 | 0.080 | 25.800 | 18.800 | 57.600 | 18.800 | 38.800 |
| Average | 0.014 | 23.489 | 16.489 | 10.058 | 7.342 | 11.863 |

From (12) we find

$$
p_{r}=(16-0) /(720-35)=0.0234
$$

from (28)
and from (22)

$$
p_{m}=0.93 \times 0.009+0.07 \times 0.0234=0.0100,
$$

$$
p_{s}=(23-0) /(720-35)=0.0336
$$

To find the optimum plan for $N=500$ from the master table with $p_{r}=p_{s}$ $=0.010$ we first have to find the conversion factor $\lambda_{s}$ which corrects for the difference between $p_{s}$ and $p_{r}$, i.e.

$$
\lambda_{s}^{-1}=1+\frac{336-234}{0.93(234-90)}=1.76
$$

To use the method of section 8 we find $\gamma_{2}=0.296, \lambda=1.97, p_{1} / \lambda=0.0046$ $\simeq 0.005, p_{2} / \lambda=0.041 \simeq 0.040$, and $N^{*}=1.97 N / 1.76=1.12 N=560$. From the master table we read $\left(n^{*}, c^{*}\right)=(60,1)$ which gives $n=60 / 1.97=30$ as the optimum sample size.

To illustrate the method of section 11 we have to find the conversion factor $f_{1}$ and the correction $g_{1}$ corresponding to the change from $w_{2}=0.05$ to 0.07 . Since $\varrho_{1}=90 / 234=0.385 \simeq 0.40$ and $\varrho_{2}=800 / 234=3.42 \simeq 3.50$ we have $f_{1}=1.13$ and $g_{1}=15$. We then enter the master table with

$$
N^{*}=N \times 2.34 \times 1.13 / 1.76=1.50 N=750
$$

and find $\left(n^{*}, c^{*}\right)=(60,1)$ which finally gives $(n, c)=(30,1)$ since $(60+15) / 2.34=32$.
To find the corresponding value of $R$ we first compute

$$
\gamma_{1}=0.93(234-90) /(336-100)=0.567
$$

and

$$
\gamma_{2}=0.07(800-234) /(336-100)=0.168
$$

which lead to

$$
R=n+(N-n)\left(0.567 Q\left(p_{1}\right)+0.168 P\left(p_{2}\right)\right) .
$$

From a table of the binomial distribution one finds for $(n, c)=(30,1)$ that $Q\left(p_{1}\right)=0.02982$ and $P\left(p_{2}\right)=0.29579$ and consequently

$$
R=30+470 \times 0.0666=61.3 .
$$

The costs of sampling inspection and the average decision losses per lot are thus of nearly the same size.

Returning to the original monetary unit we find

$$
k-k_{m}=R\left(k_{s}-k_{m}\right) / 500=1.98
$$

and finally

$$
k=7.34+1.98=9.32 .
$$

We thus have the following conclusion:
The quality of submitted lots is such that on the average costs per item will be 7.34 cents if all lots are classified correctly, i. e. all lots from process No. 1 are accepted and all lots from process No. 2 are rejected. To decide whether to accept or reject we
inspect a sample of 30 items at the average costs of 0.97 cents per item of the lot. The decision losses will be 1.01 cents per item of the lot on the average. The first part of the costs, 7.34 , depends on the prior distribution and can only be reduced by producing (or buying) lots of better quality. The second part, 1.98, depends on the sampling plan used. Since we have here used the optimum plan any change in sample size or acceptance number will result in increased costs. The average costs of accepting all lots without inspection are 10.06 cents per item.

The two functions $k_{0}(p)=K_{0}(p) / N$ and $k(p)=K(p) / N$ have been shown in Fig. 1 for the example above.

## 14. General remarks

There exists already a great body of theories and tables for constructing single sampling attribute plans based on two specified quality levels ( $p_{1}, p_{2}$ ) and some further requirements. To see how the present paper fits into this the most important systems have been listed below by stating the "further requirements" for each system:
(a). Specification of the producer's and the consumer's risks, see for instance Peach and Littauer [7] and Grubbs [8].
(b). Specification of the consumer's risk and minimization of the average amount of inspection for lots of process average quality $\left(p_{1}\right)$ in the case of rectifying inspection, see Dodge and Romig [9].
(c). Specification of the consumer's risk and minimization of the average costs for lots of process average quality $\left(p_{1}\right)$, i.e. a generalization of the Dodge-Romig LTPD system requiring specification of one cost parameter, see Hald [10].
(d). Specification of two cost parameters, $p_{r}$ and $p_{s}$, and a weight, $w_{2}$, and minimization of the average costs, as for instance in the present paper.

It follows from the results of the present paper that from an economic point of view it is not advisable to fix the consumer's or the producer's risk. On the contrary the producer's and the consumer's risks should both tend to zero with increasing lot size. This theorem is valid not only for the double binomial prior distribution but for any prior distribution, and it is valid not only for the Bayes solution but also for the minimax solution ( $p_{1}<p_{r}<p_{2}$ ), the only difference being the speed of the convergence. For a discrete prior distribution the risks tend to zero inversely proportional to $N$, see (63) and (64). These considerations lead to the result that if one wants a system with a fixed risk then the risk should be fixed to 50 per cent at a point between $p_{1}$ and $p_{2}$. We may therefore increase the list of systems of sampling plans above by the following item:
(e). Minimization of average costs for lots of process average quality ( $p_{1}$ ) under the restriction that $P\left(p_{0}\right)=1 / 2$. Such a system, named the IQL system (Indifference Quality Level) has been discussed by Hald, see [6] and [10], and will be further discussed in a forthcoming paper. This system requires the specification of $p_{0}$ and
a cost parameter. In view of the asymptotic relation (66) between $c$ and $n$ it is clear that $p_{0}$ should be determined from (52).

The simplest possible system based on the specification of two risks and having the same properties as the Bayes solution may be formulated as follows:
(f). Specification of the consumer's or the producer's risk as inversely proportional to lot size, and $P\left(p_{0}\right)=1 / 2$.

This system requires only the specification of one parameter (besides the two quality levels) and it is extremely simple to handle both mathematically and numerically. This is due to the fact that the equation $P\left(p_{0}\right)=1 / 2$ has the solution $c=n p_{0}$ $+\left(p_{0}-2\right) / 3$ (with sufficient accuracy for all practical applications, perhaps apart from the case $c=0$ where the exact solution may be easily found) and that the other equation, $Q\left(p_{1}\right)=\alpha / N$ say, may be solved with respect to $N$ for related values of $(n, c)$ from the first equation. Setting $c=0,1,2, \ldots$ and solving the first equation for $n$, the second equation gives $N=\alpha /\left(1-B\left(c, n, p_{1}\right)\right)$ which may easily be found by means of a table of the binomial (or the Poisson) distribution. The only difficulty lies in the choice of $\alpha$. If the problem is fully specified one may naturally choose $\alpha$ as the coefficient of $1 / N$ in $(63)$ and the system will then asymptotically give an approximation to the Bayes solution. The reason for using the simple system will, however, usually be that some of the parameters in the problem are unknown and in that case the choice of $\alpha$ will to some extent be arbitrary, just as in the other cases the choice of the producer's or the consumer's risk is arbitrary. This system of sampling plans will be discussed in more detail in the forthcoming paper on the $I Q L$ system.

Turning to applications it is important to notice that a system of sampling plans in practice often is required to serve several purposes. In particular we shall here stress (a) that the system should protect the consumer against deterioration of the prior distribution, (b) that the system should work as an incentive for the producer to produce better quality or at least to keep to the quality agreed upon, see Hill [11], and (c) that (average) costs should be minimized. The first two requirements are concerned with consequences of changes of the prior distribution and the problem should therefore really be formulated as a dynamic one. However, since a dynamic model at present is lacking we shall try to indicate how the Bayesian solution may be modified to take requirements (a) and (b) into account.

One of the arguments advanced against the Bayesian method in general has been that a prior distribution does normally not exist. This may be true in many fields but certainly not for industrial mass production with its effective planning and control of operations. Admittedly the prior distribution may change, but changes are usually rather small and slow within a given production period in which the same machinery, techniques, and raw materials are being used. We are here not concerned about isolated very poor lots which may occasionally occur since any sampling plan will detect such lots.

Published data on prior distributions are scarce. Whether the double binomial
distribution is a reasonable approximation to distributions occurring in practice is not known. According to the experience of the author mixed binomial distributions with beta-distributions as weight functions are rather common. (A paper analogous to the present one will present the corresponding theory and tables for the beta-distribution).

One of the drawbacks of the Bayesian solution from a practical point of view is that the solution may be acceptance (or rejection) without inspection. If one is not completely confident that the prior distribution used is the right one and is stable, then a sampling plan is required to guard against deterioration of the prior distribution. One possibility is to use the first or one of the first Bayesian sampling plans in the appropriate table. If that is not satisfactory one may in such cases use an IQL plan.

The same procedure may be used to satisfy requirement (a) above. It should first of all be noted that if a Bayesian sampling plan exists then some protection against deterioration of the prior distribution is automatically obtained and the protection may in the usual way be expressed by means of the OC curve. It is always easy when the plan has been found to compute the consumer's risk and then to decide whether the risk is sufficiently small. If the consumer's risk is too large one may again find a sufficiently large sample in the same table or turn to an $I Q L$ plan or a $L T P D$ plan.

The price to be paid for obtaining the required protection is naturally that the plan used will not minimize costs if the prior distribution holds. If the change in the value of $c$ is not large the increase in costs will, however, be small.

For large lots the consumer's risk for the Bayesian sampling plan will usually be much smaller than 10 per cent so that the problem does only exist for small lots.

The incentive for the producer to keep to the specified quality is usually obtained by alternating between normal and tightened inscpection in a specific way such that the system reacts upon observed changes in the prior distribution. If it was possible to estimate in what way the distribution had changed the reaction could be made to depend on the change. In practice, however, one want to install tightened inspection as soon as possible on the basis of some over-all criterion, for example when the number of lots rejected exceeds some critical limit. A thorough theory does not exist but some rules have been found to work satisfactory in practice. The Military Standard 105 D uses the same sample size for normal and tightened inspection and a reduced acceptance number, $c_{T}$, for tightened inspection. The difference between the two acceptance numbers, $c_{N}-c_{T}$, equals 1 for $2 \leqq c_{N} \leqq 4$, 2 for $5 \leqq c_{N} \leqq 20$, and 3 for $c_{N} \geqq 21$. For $c_{N}=0$ or $1, c_{T}$ is usually equal to $c_{N}$ but the sample size is increased for tightened inspection. Similar rules may be used for the present tables although it has to be realized that the resulting plans will not be minimum-cost plans. The main point is, that under normal conditions the plans will minimize costs and that the plans may be adjusted to changes in the prior distribution so that costs are minimized under the new conditions. If, however, the incentive aspect of sampling inspection is more important for the user of the system than to minimize costs in case of
change of the distribution then some form of tightened inspection may be introduced with the result that during periods of tightened inspection the plans will not minimize costs.

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## Appendix

## Master Tables of Sampling Plans

Tables of Conversion Factors

## Summary of Conversion Formulas

All sampling plans in the master tables assume $p_{r}=p_{s}$ and $w_{2}=0.05$. In the first set of tables $p_{r}=10 \%$ and $\left(p_{1}, p_{2}\right)$ take on the values

$$
\begin{aligned}
& p_{1}=2.0,2.5,3.0,3.5,4.0,5.0,6.0,7.0 \% \\
& p_{2}=15.0,17.5,20.0,25.0,30.0 \% .
\end{aligned}
$$

In the second set of tables $p_{r}=1 \%$ and $\left(p_{1}, p_{2}\right)$ take on the values

$$
\begin{aligned}
& p_{1}=0.20,0.25,0.30,0.35,0.40,0.50,0.60,0.70 \% \\
& p_{2}=1.50,1.75,2.00,2.50,3.00,3.50,4.00,5.00,6.00,7.00 \% .
\end{aligned}
$$

Single Sampling Tables for $p_{1}=2.0 \%$

| $p_{2}=15.0 \%$ |  |  | $p_{2}=17.5 \%$ |  |  | $p_{2}=20.0 \%$ |  |  |  | $p_{2}=25.0 \%$ |  |  |  | $p_{2}=30.0 \%$ |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| $N$ | $n$ | c | $N$ | $n$ | c | $N$ |  | $n$ | c | $N$ |  | $n$ | $c$ | $N$ |  | $n$ | $c$ |
| 1- 1460 | Accept |  | 1- 665 | Accept |  | $1-$ | 334 | Accept |  | $1-$ | 140 | Accept |  | 1 |  | ccept |  |
| 1460-1530 | 16 | 2 | $666-701$ | 16 | 2 | $335-$ | 414 | 7 | 1 | 141- | 149 | 6 | 1 | $46-$ |  | 1 | 0 |
| 1600- 1840 | 29 | 3 | - 1020 | 18 | 2 | 479- | 586 | 17 | 2 | 191- | 259 | 8 | 1 | $93-$ |  | 6 | 1 |
| 2160-2630 | 313 |  |  |  |  | 727- | 946 | 19 | 2 |  |  |  |  | 139 | 199 | 81 |  |
|  |  |  | 1170-1270 | 29 | 3 |  |  |  |  | 358- | 396 | 16 | 2 |  |  |  |  |
| 2630- 2760 | 43 | 4 | 1510-1850 | 31 | 3 | 1020- | 1240 | 29 | 3 | 509- | 682 | 18 | 2 | 316 | 346 | 15 | 2 |
| 3180-3760 | 45 | 4 | 2270- 2710 | 43 | 4 | 1540- | 1980 | 31 | 3 | $945-$ | 1200 | 27 | 3 | 471- | 679 | 17 | 2 |
| 4560-4770 | 58 | 5 | 3280-4100 | 45 | 4 | 2200 | 2530 | 41 | 4 | 1590- | 2240 | 29 | 3 | $974-$ | 1400 | 25 | 3 |
| 5530-6550 | 60 |  |  |  |  | 3120 | 3990 | 43 | 4 |  |  |  |  | 2030 | 2800 | 3 |  |
|  |  |  | 4410-4810 | 56 | 5 |  |  |  |  | 2380 | 2720 | 37 | 4 |  |  |  |  |
| 7990-9560 | 74 | 6 | 760-7080 | 58 | 5 | 4690- | 5040 | 53 | 5 | 3560 | 4880 | 39 | 4 | 2800 | 3970 | 34 | 4 |
| 11300-14000 | 76 | 6 | 8530-10000 | 70 | 6 | 6200 | 7850 | 55 | 5 | 5910 | 7790 | 48 | 5 | 5750- | 7850 | 36 | 4 |
| 14000-16400 | 89 | 7 | 12200-15200 | 72 | 6 | 9860- | 12100 | 66 | 6 | 10500-14200 |  | 50 | 5 | 7850- | 11000 | 43 | 5 |
| 19500-24300 | 91 | 7 |  |  |  | 15300-20400 |  | 68 | 6 |  |  | 15900-21600 |  | 45 |  |  |  |  |
|  |  |  | 16400-17400 | 83 | 7 |  |  | $14200-$ |  | 16800 | 58 |  |  |  |  | 6 |  |
| 24300-28200 | 104 | 8 | 20900-25700 | 85 | 7 | 20400- | 23600 |  | 78 | 7 | 22200- | 30800 | 60 | 6 | 21600- | 29800 | 52 | 6 |
| 33300-40300 | 106 | 8 | $31300-35800$ | 97 | 8 | 29400- | 37800 | 80 | 7 | $34600-$ | 47100 | 69 | 7 | $43000-$ | 58900 | 54 | 6 |
| 42500-48000 | 119 | 9 | 43400-54100 | 99 | 8 | $42500-$ | 45600 | 90 | 8 | $63800-81900$$81900-99600$ |  | 71 | 7 | 58900- | 80000 | 61 | 7 |
| 56700-68400 | 121 |  |  |  |  | 56400-71700 |  | 92 | 8 |  |  | 115000-159000 |  | 63 |  |  |  |  |
|  |  |  | 59700-73500 | 111 | 9 |  |  | 79 |  |  |  | 8 |  |  |  |  |  |  |  |
| 73800-81600 | 134 | 10 | 90200-113000 | 113 | 9 | 87600-1 | 08000 |  | 103 | 9 | 132000-200000 |  | 81 | 8 | 159000-200000 |  | 70 |  |
| 96200-116000 | 136 | 10 | 113000-125000 | 124 | 10 | 136000-1 | 78000 | 105 | 9 |  |  |  |  |  |  |  |  |  |  |  |  |  |  |
| 128000-138000 | 149 | 11 | 151000-188000 | 126 | 10 | 178000-200000 |  | 11510 |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |
| 163000-200000 | 151 | 11 |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |

For $N$ between two intervals adjacent in the table find $(n, c)$ for the first of these intervals and use $(n+1, c)$ as optimum plan.
Single Sampling Tables for $p_{1}=2.5 \%$

For $N$ between two intervals adjacent in the table find $(n, c)$ for the first of these intervals and use $(n+1, c)$ as optimum plan.
Single Sampling Tables for $p_{1}=3.0 \%$

| $p_{2}=15.0 \%$ |  |  | $p_{2}=17.5 \%$ |  |  | $p_{2}=20.0 \%$ |  |  | $p_{2}=25.0 \%$ |  |  |  | $p_{2}=30.0 \%$ |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| N | $n$ | $c$ | N | $n$ | c | N | $n$ | $c$ | $N$ |  |  |  | $N$ | $n$ |  |
| 1- 1760 | Accept |  | 1- 749 | Accept |  | 1- 396 | Accept |  | 1- | 138 | Accept |  | $1-\quad 57$ | Accept |  |
| 1760- 2100 | 35 | 4 | 750- 823 | 14 | 2 | 397- 481 | 14 | 2 | $139-$ | 178 | 6 | 1 | 58- 78 | 1 | 0 |
| 2420- 2960 | 48 | 5 | 824- 980 | 24 | 3 | 636-730 | 24 | 3 | 248- | 299 | 14 | 2 | $79-109$ | 6 | 1 |
| 3450-4180 | 61 | 6 |  |  |  | 923 | 26 | 3 | 404- | 523 | 16 | 2 | $167-209$ | 81 |  |
|  |  |  | 1250- 1340 | 35 | 4 | - 1110 |  |  |  |  |  |  |  |  |  |
| 4960-5910 | 74 | 7 | 1630-1990 | 37 | 4 | 1110- 1340 | 35 | 4 | 524- | 602 | 23 | 3 | 210- 298 | 14 | 2 |
| 7160-8350 | 87 | 8 | 1990-2200 | 47 | 5 | 1730-1920 | 37 | 4 | $807-$ | 1060 | 25 | 3 | $454-503$ | 16 | 2 |
| 10300-11800 | 100 | 9 | 2710-3150 | 49 | 5 | 1920-2440 | 46 | 5 | $1060-$ | 1150 | 32 | 4 | 504-708 | 223 |  |
| 14200-15000 | 102 | 9 |  |  |  |  |  |  | 1530 | 2110 | 34 |  |  |  |  |
|  |  |  | $3150-3600$ | 59 | 6 | $3250-3470$ | 56 | 6 |  |  |  |  | 1130- 1590 | 30 | 4 |
| 15000-16500 | 113 | 10 | 4460- 4970 | 61 | 6 | $4370-5550$ | 58 | 6 | 2110 | 2830 | 42 | 5 | 2520-3460 | 38 | 5 |
| 20000-21700 | 115 | 10 | 4970- 5840 | 71 | 7 | $5550-6120$ | 67 | 7 | 4080- | 5140 | 51 | 6 | 5500-7410 | 46 | 6 |
| 21700-23200 | 126 | 11 | $7300-7820$ | 73 | 7 | 7760-9380 | 69 | 7 | 7120 | 7910 | 53 | 6 | 11100-12000 | 486 |  |
| 27900-31400 | 12811 |  |  |  |  |  |  |  |  |  |  |  |  |  |  |
|  |  |  | 7820-9440 | 83 | 8 | 9380-10700 | 78 | 8 | 7910- | 9280 | 60 | 7 | 12000-15700 | 54 | 7 |
| $31400-39000$ | 140 | 12 | 12200-15200 | 95 | 9 | 13700-15800 | 80 | 8 | 12700- | 15200 | 62 | 7 | 23500-25700 | 56 | 7 |
| 45200-54400 | 153 | 13 | 19100-24400 | 107 | 10 | 15800-18700 | 89 | 9 | $15200-$ | 16600 | 69 | 8 | 25700-33000 | 62 | 8 |
| 65000-75800 | 166 | 14 |  |  |  | 24200-26400 | 91 | 9 | 22400-29000 |  | 71 | 8 | 49100-54800 | 648 |  |
|  |  |  | 29700-31700 | 118 | 11 |  |  |  |  |  |  |  |  |  |  |  |
| 93300-106000 | 179 | 15 | 39100-46300 | 120 | 11 | 26400-32600 | 100 | 10 | 29000- | 39500 | 79 | 9 | 54800-68800 | 70 | 9 |
| 128000-134000 | 181 | 15 | 46300-50400 | 130 | 12 | 44000-56500 | 111 | 11 | 54600- | 69500 | 88 | 10 | 102000-116000 | 72 | 9 |
| 134000-147000 | 192 | 16 | 62500-71800 | 132 | 12 | 73100-77200 | 121 | 12 | 97000-103000 |  | 90 | 10 | 116000-143000 | $78 \quad 10$ |  |
| 178000-193000 | 194 | 16 |  |  |  | $97800-121000$ | 123 | 12 |  |  |  |  |  |  |  |  |
|  |  |  | 71800-80000 | 142 | 13 |  |  |  | 103000-1 | 22000 | 97 | 11 |  |  |  |
| 193000-200000 | 205 | 17 | 99600-111000 | 144 | 13 | 121000-133000 | 132 | 13 | 168000-1 | 94000 | 99 | 11 |  |  |  |
|  |  |  | 111000-127000 | 154 | 14 | 169000-200000 | 13413 |  | 194000-200000 |  | 10612 |  |  |  |  |
|  |  |  | 159000-172000 | 156 | 14 |  |  |  |  |  |  |  |  |  |  |  |  |  |
|  |  |  | 172000-200000 | 166 | 15 |  |  |  |  |  |  |  |  |  |  |  |  |  |

Single Sampling Tables for $p_{1}=3.5 \%$

For $N$ between two intervals adjacent in the table find $(n, c)$ for the first of these intervals and use $(n+1, c)$ as optimum plan.
Single Sampling Tables for $p_{1}=4.0 \%$

| $p_{2}=15.0{ }_{0}$ |  |  | $p_{2}=17.5 \%$ |  |  | $p_{2}=20.0 \%$ |  |  |  | $p_{2}=25.0 \%$ |  |  |  | $p_{2}=30.0 \%$ |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| $N$ | $n$ |  | $N$ | Accept |  |  |  | $n$ |  | $N$ |  | $n$ |  | $N$ |  | $n$ |  |
| 1- 1950 | Accept |  | 1- 806 |  |  | 1- | 402 |  | cept | 1 - | 134 |  | ept | 1- | 65 | Acc | ept |
| 1950- 2130 | 40 | 5 | 807- 923 | 21 | 3 | 403- | 447 | 12 | 2 | 135- | 151 | 5 | 1 | 66- | 79 | 5 | 1 |
| 2260- 2850 | 52 | 6 | $924-981$ | 30 | 4 | 487- | 576 | 21 | 3 | 198- | 276 | 13 | 2 | 119 | 156 | 7 | 1 |
| 2850-3090 | 63 | 7 | 1260- 1470 | 41 | 5 | $736-$ | 944 | 31 | 4 | 359- | 456 | 21 | 3 | 157 | 244 | 13 | 2 |
| 3670- 4130 | 75 | 8 | 1770-2210 | 52 | 6 | 1110 | 1190 | 40 | 5 | 632- | 719 | 29 | 4 | 326- | 442 | 20 | 3 |
| 4730- 5520 | 87 | 9 | 2470- 2650 | 62 | 7 | $1530-$ | 1670 | 42 | 5 | 996- | 1100 | 31 | 4 | $640-$ | 760 | 27 | 4 |
| 6120-7390 | 99 | 10 | $3420-3930$ | 73 | 8 | $1670-$ | 1880 | 50 | 6 | $1100-$ | 1490 | 38 | 5 | 1140 | 1240 | 29 | 4 |
| 7920-10200 | 111 | 11 | 4780-5830 | 84 | 9 | $2490-$ | 2950 | 60 | 7 | 1860- | 2220 | 46 | 6 | $1240-$ | 1860 | 35 | 5 |
| 10200-10800 | 122 | 12 | 6660-8660 | 95 | 10 | $3730-$ | 4610 | 70 | 8 | 3130 | 4530 | 55 | 7 | $2320-$ | 3010 | 42 | 6 |
| 13200-14400 | 134 | 13 | 9170-10100 | 105 | 11 | $5550-$ | 7190 | 80 | 9 | $5210-$ | 6550 | 63 | 8 | $4300-$ | 4870 | 49 | 7 |
| 17000-19100 | 146 | 14 | 12700-14900 | 116 | 12 | 8210 | 11200 | 90 | 10 | 8620 - | 9530 | 71 | 9 | 7280 | 7980 | 51 | 7 |
| 22000-25500 | 158 | 15 | 17500-21900 | 127 | 13 | $12000-$ | 13200 | 99 | 11 | $13300-$ | 14200 | 73 | 9 | 7980 | 11400 | 57 | 8 |
| 28400-33900 | 170 | 16 | 24100-25500 | 137 | 14 | 17500- | 20300 | 109 | 12 | $14200-$ | 18900 | 80 | 10 | $14600-$ | 18000 | 64 | 9 |
| 36600-45100 | 182 | 17 | 33000-37200 | 148 | 15 | 25700- | 31100 | 119 | 13 | $23400-$ | 27100 | 88 | 11 | $26800-$ | 42500 | 72 | 10 |
| 47200-60400 | 194 | 18 | 45500-54300 | 159 | 16 | $37700-$ | 47600 | 129 | 14 | 38300- | 53700 | 97 | 12 | $48500-$ | 65700 | 79 | 11 |
| 60400-64700 | 205 | 19 | 62600-79400 | 170 | 17 | $55200-$ | 73000 | 139 | 15 | 62500- | 76100 | 105 | 13 | 87600-1 | 103000 | 86 | 12 |
| 77800-85600 | 217 | 20 | 85500-91700 | 180 | 18 | 79900- | 85600 | 148 | 16 | 103000-1 | 51000 | 114 | 14 | 158000-200 | 200000 | 94 | 13 |
| 100000-113000 | 229 | 21 | 117000-133000 | 191 | 19 | 115000-1 | 30000 | 158 | 17 | 165000-2 | 200000 | 122 | 15 |  |  |  |  |
| 129000-150000 | 241 | 22 | 160000-200000 | 202 | 20 | 168000-2 | 00000 | 168 | 18 |  |  |  |  |  |  |  |  |
| 165000-200000 | 253 | 23 |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |

For $N$ between two intervals adjacent in the table find $(n, c)$ for the first of these intervals and use $(n+1, c)$ as optimum plan.
Single Sampling Tables for $p_{1}=5.0 \%$

Single Sampling Tables for $p_{1}=6.0 \%$

For $N$ between two intervals adjacent in the table find $(n, c)$ for the first of these intervals and use $(n+1, c)$ as optimum plan.
(Continued on next page)
Single Sampling Tables for $p_{1}=6.0 \%$ (Continued)

| $p_{2}=15.0 \%$ |  |  | $p_{2}=17.5 \%$ |  |  | $p_{2}=20.0 \%$ |  |  | $p_{2}=25.0 \%$ |  |  | $p_{2}=30.0 \%$ |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| $N$ | $n$ | c | N | $n$ | ${ }^{\text {c }}$ | $N$ | $n$ | $c$ | N | $n$ | c | $N$ | $n$ | c |
| 17200-19600 | 231 | 25 | 14700-17500 | 180 | 21 | 20000-24800 | 161 | 20 | 33900-42500 | 128 | 18 | 46200-67600 | 102 | 16 |
| 19600-22200 | 241 | 26 | 17500-19500 | 189 | 22 | 24800-27500 | 169 | 21 | 46500-51600 | 135 | 19 | 67600-80200 | 108 | 17 |
| 22200-25200 | 251 | 27 | 20900-22100 | 198 | 23 | $31100-38600$ | 178 | 22 |  |  |  | 102000-149000 | 115 | 18 |
| 25200- 28600 | 261 | 28 | 25000-29700 | 208 | 24 | 38600-43600 | 186 | 23 | 63800-86100 | 143 | 20 | 149000-183000 | 121 | 19 |
| 28600-32500 | 271 | 29 |  |  |  | 48400-60000 | 195 | 24 | 86100-109000 | 150 | 21 |  |  |  |
|  |  |  | 29700-35300 | 217 | 25 |  |  |  | 117000-132000 | 157 | 22 |  |  |  |
| 32500-36800 | 281 | 30 | 35300-41900 | 226 | 26 | 60000-69200 | 203 | 25 | 161000-200000 | 165 | 23 |  |  |  |
| 36800-40600 | 291 | 31 | 41900-46700 | 235 | 27 | 75000-92900 | 212 | 26 |  |  |  |  |  |  |
| 41900-45000 | 301 | 32 | 50100-52700 | 244 | 28 | 92900-109000 | 220 | 27 |  |  |  |  |  |  |
| 47600-49900 | 311 | 33 | 59700-70800 | 254 | 29 | 116000-144000 | 229 | 28 |  |  |  |  |  |  |
| 54100-55400 | 321 | 34 |  |  |  | 144000-177000 | 237 | 29 |  |  |  |  |  |  |
|  |  |  | 70800-83900 | 263 | 30 |  |  |  |  |  |  |  |  |  |
| 61500-69700 | 332 | 35 | 83900-99300 | 272 | 31 | 177000-187000 | 245 | 30 |  |  |  |  |  |  |
| 69700-79000 | 342 | 36 | 99300-110000 | 281 | 32 |  |  |  |  |  |  |  |  |  |
| 79000-89400 | 352 | 37 | 118000-124000 | 290 | 33 |  |  |  |  |  |  |  |  |  |
| 89400-101000 | 362 | 38 | 141000-167000 | 300 | 34 |  |  |  |  |  |  |  |  |  |
| 101000-115000 | 372 | 39 | 167000-200000 |  | 35 |  |  |  |  |  |  |  |  |  |
| 115000-130000 | 382 | 40 |  |  |  |  |  |  |  |  |  |  |  |  |
| 130000-142000 | 392 | 41 |  |  |  |  |  |  |  |  |  |  |  |  |
| 147000-157000 | 402 | 42 |  |  |  |  |  |  |  |  |  |  |  |  |
| 167000-174000 | 412 | 43 |  |  |  |  |  |  |  |  |  |  |  |  |
| 190000-193000 | 422 | 44 |  |  |  |  |  |  |  |  |  |  |  |  |

Single Sampling Tables for $p_{1}=7.0 \%$

| $p_{2}=15.0 \%$ |  |  | $p_{2}=17.5 \%$ |  |  |  | $p_{2}=20.0 \%$ |  |  |  | $p_{2}=25.0 \%$ |  |  |  | $p_{2}=30.0 \%$ |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| N | Accept |  | $N$ |  | $n \quad c$ |  | $N$ |  | $n \quad c$ |  | N |  | $n \quad c$ |  | $N$ |  | $n \quad c$ |  |
| 1- 1960 |  |  |  | 668 | Accept |  | $1-$ | 287 | Accept |  |  | 73 | Accept |  | 1. | 19 | Accept |  |
|  | Accept |  |  |  |  |  | $288-$ | 335 | 17 | 3 |  | 101 | 5 | 1 |  |  |  |  |
| 60- 2030 | 73 | 10 | 669- | 706 | 31 | 5 | 336- | 357 | 24 | 4 | 102 | 129 | 11 | 2 | 20 | 36 | 1 | 0 |
| 2030-2220 | 82 | 11 | $707-$ | 35 | 39 | 6 |  |  |  |  | $161-$ | 211 | 18 | 3 | $37-$ | 45 | 5 | 1 |
| 2220- 2260 | 91 | 12 | 823- | 958 | 48 | 7 | 424 | 462 | 32 | 5 | 6 | 331 | 25 | 4 | $76-$ | 88 | 11 | 2 |
| 2440-2680 | 101 | 13 | 959 | 1120 | 57 | 8 | $532-$ | 591 | 40 | 6 |  |  |  |  | $135-$ | 213 | 18 | 3 |
| 2680 | 110 | 14 | 1120 | 1220 | 65 |  | 663 | 746 | 48 | 7 | 332 | 460 | 32 | 5 | 214 | 324 | 24 4 |  |
|  |  |  |  |  |  |  | 818- | 934 | 56 | 8 | 461- | 627 | 39 | 6 |  |  |  |  |
| 2940-3240 | 120 | 15 | 1300- | 1520 | 74 | 10 | 1000- | 1160 | 64 | 9 | $628-$ | 837 | 46 | 7 | 325- | 478 | 30 | 5 |
| 3240-3430 | 129 | 16 | 1520- | 1760 | 83 | 11 |  |  |  |  | $838-$ | 892 | 52 | 8 | 479 - | 662 | 36 | 6 |
| $3560-3920$ | 139 | 17 | 1760- | 1970 | 91 | 12 | 1220- | 1430 | 72 | 10 | 1110- | 1220 | 59 | 9 | $701-$$1010-$ | 9121240 | 42 |  |
| 3920-4230 | 148 | 18 | 2040- | 2360 | 100 | 13 | 1490 | 1800 | 80 | 11 |  |  |  |  |  |  | 48 | 8 |
| 4310-4750 | 1581 | 19 | 2360 | 2450 | 108 | 14 | 1800 | 2170 | 88 | 12 | $\begin{aligned} & 1470- \\ & 1930- \end{aligned}$ | 1660 | 66 | 10 | 1460 | 1690 | 54 |  |
|  |  |  |  |  |  |  | 2170 | 2620 | 96 | 13 |  | 2250 | 73 | 11 |  |  |  |  |
| 4750-5210 | 167 | 20 | 2720- | 30 | 117 | 15 | 2620 | 3150 | 104 | 14 | 2520 | 3040 | 80 | 12 | 2080- | 2280 | 60 | 10 |
| 5210- 5740 | 177 | 21 | 3130- | 3610 | 126 | 16 |  |  |  |  | 3290 | 4100 | 87 | 13 | 2980- | 4180 | 67 | 11 |
| 5740-6310 | 186 | 22 | 3610 | 3820 | 134 | 17 | 3150 | 3780 | 112 | 15 | 4280 - | 5550 | 94 | 14 | $4180-$ | 5860 | 73 | 12 |
| 6310-6920 | 196 | 23 | 4150 | 4760 | 143 | 18 | 3780 | 4540 | 120 | 16 |  |  |  |  | $5860-$ | 8180 | 79 | 13 |
| 6920-7620 | 205 | 24 | 4760- | 5470 | 152 | 19 | 4540- | 5430 | 128 | 17 | 5550 | 220 | 101 | 15 | 8180-11400 |  | 85 | 14 |
|  |  |  |  |  |  |  | 5430 | 6490 | 136 | 18 | 7220 | 9390 | 108 | 16 |  |  |  |  |  |
| 7620-8350 | 215 | 25 | 5470- | 5880 | 160 | 20 | 6490- | 7760 | 144 | 19 | 9390 | 12200 | 115 | 17 | 11400- | 15100 | 91 | 15 |
| 8350-9170 | 224 | 26 | 6260- | 7190 | 69 | 21 |  |  |  |  | 12200- | 15800 | 122 | 18 | 16000- | 20000 | 97 | 16 |
| 9170-9420 | 233 | 27 | 7190- | 8220 | 178 | 22 | 7760 | 9260 | 152 | 20 | 15800 | 20300 | 129 | 19 | 22500 | 26400 | 103 | 17 |
| 10000-11000 | 243 | 28 | 8220- | 8990 | 186 | 23 | 9260 | 11000 | 160 | 21 |  |  |  |  | $31500-$ | 34900 | 109 | 18 |
| 11000-11500 | 252 | 29 | 9380-10800 |  | 195 | 24 | 11000- | 13100 | 168 | 22 | 20300- | 21500 | 135 | 20 | 44400-61500 |  | 116 | 19 |
|  |  |  |  |  | 13100- |  | 15600 | 176 | 23 | 26100 | 28500 | 142 | 21 |  |  |  |  |  |  |
| 12100-13300 | 262 | 30 | 10800- | 12300 |  | 204 | 25 | 15600 | - 18600 | 184 | 24 | $\begin{aligned} & 33500- \\ & 43100- \end{aligned}$ | $\begin{aligned} & 37800 \\ & 50200 \end{aligned}$ | 149 | 22 | 61500-85100 |  | $122 \quad 20$ |  |
| 13300-14000 | 271 | 31 | 12300 | 13700 | 212 | 26 | 23 |  |  |  |  |  |  |  | 85100-1 | 117000 | 128 | 21 |
| 14500-15900 | 281 | 32 | $14000-$ | 16000 | 221 | 27 | 18600- | 22100 | 192 | 25 | 55300 | 66700 | 163 | 24 | 117000-1 | 162000 | 134 | 22 |
| 15900-17100 | 290 | 33 | $16000-$ | 16400 | 229 | 28 | $22100-$ | 26300 | 200 | 26 |  |  |  |  | 162000-2 | 200000 | 140 | 23 |
| 17400-19100 | 300 | 34 | 18200- | 20700 | 238 | 29 | 26300- | 31200 | 208 | 27 | 70900- | 90500 | 170 | 25 |  |  |  |  |


For $N$ between two intervals adjacent in the table find ( $n, c$ ) for the first of these intervals and use $(n+1, c)$ as optimum plan.
Mat.Fys.Skr.Dan.Vid.Selsk. 3, no. 2.
Single Sampling Tables for $p_{1}=0.20 \%$

Single Sampling Tables for $p_{1}=0.20 \%$

Single Sampling Tables for $p_{1}=0.25 \%$


| $p_{2}=3.50 \%$ |  | $p_{2}=4.00 \%$ |  | $p_{2}=5.00 \%$ |  |  | $p_{2}=6.00 \%$ |  |  | $p_{2}=7.00 \%$ |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| $N$ | $n \quad c$ | $N$ | $n \quad c$ | $N$ | $n$ | c | $N$ | $n$ | $c$ | $N$ | $n$ | c |
| 1- 259 | Accept | 1- 154 | Accept | $1-75$ | Acce |  | $1-45$ | Acc |  | 1 - 30 | Acc |  |
| 260- 370 | 50 | 155- 220 | 50 |  |  |  |  |  |  |  |  |  |
| 554- 65 | 150 | 320- 486 | 150 | 76- 114 | 5 | 0 | 46- 74 | 5 | 0 | 31- 55 | 5 | 0 |
|  |  |  |  | 171- 258 | 15 | 0 | 119-185 | 15 | 0 | 93- 151 | 15 | 0 |
| 658-736 | $55 \quad 1$ | $556-618$ | $55 \quad 1$ | 414- 453 | 25 | 0 | 304- 404 | 25 | 0 | 260- 365 | 25 | 0 |
| 846- 984 | 65 | 728 - 866 | $65 \quad 1$ |  |  |  |  |  |  |  |  |  |
| 1170- 1410 | 75 | 1050- 1310 | $75 \quad 1$ | 454- 533 | 55 | 1 | 405- 542 | 55 | 1 | 366- 451 | 50 | 1 |
| 1750- 2120 | $85 \quad 1$ | 1680- 2000 | $85 \quad 1$ | $660-831$ | 65 | 1 | $712-960$ | 65 | 1 | $615-864$ | 60 | 1 |
|  |  |  |  | 1080- 1440 | 75 | 1 | 1360-1830 | 75 | 1 | 1280-1820 | 70 | 1 |
| 2120- 2210 | $135 \quad 2$ | 2000- 2160 | $130 \quad 2$ |  |  |  |  |  |  |  |  |  |
| 2550- 2980 | $145 \quad 2$ | 2560-3070 | $140 \quad 2$ | 1870- 2230 | 120 | 2 | 1830- 2310 | 110 | 2 | 1820- 2230 | 100 | 2 |
| $3530-4260$ | $155 \quad 2$ | 3750- 4700 | 1502 | 2780-3550 | 130 | 2 | 3030- 4100 | 120 | 2 | 3060-4360 | 110 | 2 |
| 5270-5790 | $165 \quad 2$ |  |  | 4660-6290 | 140 | 2 | 5800-6920 | 130 | 2 | 6570-7600 | 120 | 2 |
|  |  | 5880-6570 | 2053 |  |  |  |  |  |  |  |  |  |
| 5790-6690 | $220 \quad 3$ | 7830-9480 | 2153 | 6290- 6640 | 180 | 3 | 6920-8730 | 165 |  | 7600-9960 | 150 | 3 |
| 7820- 9270 | $230 \quad 3$ | 11700-14800 | 2253 | 8230-10400 | 190 | 3 | 11500-15700 | 175 | 3 | 13800-19900 | 160 | 3 |
| 11200-13800 | $240 \quad 3$ |  |  | 13500-18100 | 200 | 3 | 22300-24600 | 185 | 3 |  |  |  |
|  |  | 16300-19100 | $280 \quad 4$ |  |  |  |  |  |  | 30000-42800 | 200 | 4 |
| 14800-16600 | 300 | 22900-27900 | $290 \quad 4$ | 20000-23300 | 245 | 4 | 24600-31700 | 220 | 4 | 60000-87800 | 210 | 4 |
| 19500-23100 | 310 | 34800-43700 | $300 \quad 4$ | 29200-37400 | 255 | 4 | 42000-57300 | 230 | 4 |  |  |  |
| 27800-34300 | $320 \quad 4$ |  |  | 49400-62100 | 265 | 4 |  |  |  | 115000-131000 | 245 | 5 |
|  |  | 43700-45500 | $350 \quad 5$ |  |  |  | 84900-112000 | 275 | 5 | 180000-200000 | 255 | 5 |
| 36700-40300 | $380 \quad 5$ | 54200-65500 | $360 \quad 5$ | 62100-80400 | 310 | 5 | 149000-200000 | 285 | 5 |  |  |  |
| 47200-55900 | $390 \quad 5$ | 80400-101000 | $370 \quad 5$ | 102000-132000 | 320 | 5 |  |  |  |  |  |  |
| 67400-82800 | $400 \quad 5$ |  |  |  |  |  |  |  |  |  |  |  |
|  |  | 116000-127000 | $425 \quad 6$ | 186000-200000 | 370 | 6 |  |  |  |  |  |  |
| 89700-96200 | $460 \quad 6$ | 152000-184000 | 4356 |  |  |  |  |  |  |  |  |  |
| 113000-133000 | $470 \quad 6$ |  |  |  |  |  |  |  |  |  |  |  |
| 160000-200000 | $480 \quad 6$ |  |  |  |  |  |  |  |  |  |  |  |

[^0]



[^1]

[^2]Single Sampling Tables for $p_{1}=0.40 \%$



[^3]Mat.Fys.Skr. Dan. Vid.Selsk. 3, no. 2.
Single Sampling Tables for $p_{1}=0.40 \%$

| $p_{2}=3.50 \%$ |  |  | $p_{2}=4.00 \%$ |  |  | $p_{2}=5.00 \%$ |  |  |  | $p_{2}=6.00 \%$ |  |  |  | $p_{2}=7.00 \%$ |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| N | $n$ | c | $N$ | $n$ | $c$ | $N$ |  | $n$ | c | N |  | $n$ | c | N |  | $n$ | $c$ |
| 1- 249 | Accept |  | 1- 133 | Accept |  | 1- | 59 | Accept |  | 1 - | 33 | Accept |  | 1 - | 21 | Accept |  |
| 250 | 5 | 0 | 134- 199 | 5 | 0 |  | 94 | 5 |  |  | 59 | 50 |  | 22 | 43 | 50 |  |
|  |  |  | 312- 406 | 15 |  |  | 235 | 15 | 0 | 98 | 159 | 15 | 0 | 76- | 128 | 15 | 0 0 |
| 497- 526 | 45 | 1 |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |
| 608- 714 | 55 | 1 | 407- 482 | 50 | 1 | 324 | 387 | 50 | 1 | 281- | 373 | 50 | 1 | $249-$ | 293 | 45 | 1 |
| 864-1080 | 65 | 1 | 579-710 | 60 | 1 | 488- | 630 | 60 | 1 | 497- | 691 | 60 | 1 | 402- | 575 | 55 | 1 |
|  |  |  | $903-1170$ | 70 | 1 | $854-$ | 1050 | 70 | 1 | 1000- | 1340 | 100 | 2 | 889- | 972 | 65 | 1 |
| 1290- 1520 | 120 | 2 |  |  |  |  |  |  |  | 1800- | 2570 | 110 | 2 |  |  |  |  |
| 1800- 2180 | 130 | 2 | 1170-1350 | 115 | 2 | 1050 | 1170 | 105 | 2 |  |  |  |  | 973- | 1150 | 90 | 2 |
| 2770-3270 | 190 | 3 | 1620- 2000 | 125 | 2 | $1460-$ | 1890 | 115 | 2 | 2860 - | 3190 | 145 | 3 | $1600-$ | 2350 | 100 | 2 |
| 3890- 4730 | 200 | 3 | 2560- 2720 | 135 | 2 | $2570-$ | 2730 | 125 | 2 | 4220 - | 5830 | 155 | 3 | 3050- | 4030 | 135 | 3 |
|  |  |  |  |  |  |  |  |  |  | 7680 | 9530 | 195 | 4 | $5730-$ | 8830 | 145 | 3 |
| 5600-6610 | 260 | 4 | 2720-3200 | 180 | 3 | $2730-$ | 3020 | 160 | 3 | 12900- | 18200 | 205 | 4 |  |  |  |  |
| 7880- 9640 | 270 | 4 | 3890- 4860 | 190 | 3 | $3800-$ | 4950 | 170 | 3 |  |  |  |  | $8830-$ | 9820 | 175 | 4 |
| 10900-13000 | 330 | 5 | 5860-7200 | 245 | 4 | $6600-$ | 7400 | 215 | 4 | 20100- | 28000 | 245 | 5 | $13600-$ | 20000 | 185 | 4 |
| 15500-19100 | 340 | 5 | 8840-11200 | 255 | 4 | 9390 | 12300 | 225 | 4 | $38600-$ | 51000 | 255 | 5 | $25500-$ | 32200 | 220 | 5 |
|  |  |  |  |  |  |  |  |  |  | $51000-$ | 60900 | 290 | 6 | $45700-$ | 71500 | 230 | 5 |
| 20700-25000 | 400 | 6 | 12100-13000 | 305 | 5 | $15500-$ | 17700 | 270 | 5 | 81900-1 | 15000 | 300 | 6 |  |  |  |  |
| 30000-37000 | 410 | 6 | 15800-19600 | 315 | 5 | $22600-$ | 29800 | 280 | 5 |  |  |  |  | 71500-1 | 05000 | 265 | 6 |
| 39000-47600 | 470 | 7 | 24700-28000 | 370 | 6 | $35800-$ | 41700 | 325 | 6 | 130000-1 | 74000 | 340 | 7 | 153000-2 | 00000 | 275 | 6 |
| 57400-72300 | 480 | 7 | 34100-42700 | 380 | 6 | $53500-$ | 71000 | 335 | 6 |  |  |  |  |  |  |  |  |
| 72300-75900 | 535 | 8 | 50000-59500 | 435 | 7 | 81500- | 97500 | 380 | 7 |  |  |  |  |  |  |  |  |
| 90000-109000 | 545 | 8 | 73200-92500 | 445 | 7 | 125000-1 | 167000 | 390 | 7 |  |  |  |  |  |  |  |  |
| 134000-142000 | 605 | 9 | 100000-126000 | 500 | 8 | 184000-2 | 200000 | 435 | 8 |  |  |  |  |  |  |  |  |
| 169000-200000 | 615 | 9 | 156000-200000 | 510 | 8 |  |  |  |  |  |  |  |  |  |  |  |  |

[^4]
$\boldsymbol{p}_{1}=\mathbf{0 . 5 0} \%$
Single Sampling Tables for $p_{1}=0.50 \%$


For $N$ between two intervals adjacent in the table find $(n, c)$ for the first of these intervals and use $(n+5, c)$ as optimum plan.
Single Sampling Tables for $p_{1}=0.60^{\circ}{ }_{0}$

For $N$ between two intervals adjacent in the table find $(n, c)$ for the first of these intervals and use $(n+5, c)$ as optimum plan.
Single Sampling Tables for $p_{1}=0.60 \%$


For $N$ between two intervals adjacent in the table find $(n, c)$ for the first of these intervals and use $(n+5, c)$ as optimum plan.
Single Sampling Tables for $p_{1}=0.70 \%$

| $p_{2}=1.50 \%$ |  |  | $p_{2}=1.75 \%$ |  |  | $p_{2}=2.00 \%$ |  |  | $p_{2}=2.50 \%$ |  |  | $p_{2}=3.00 \%$ |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| $N$ | $n$ | c | $N$ | $n$ | c | $N$ | $n$ | $c$ | $N$ | $n$ | $c$ | $N$ | $n$ | $c$ |
| 1-21600 | Accept |  | 1- 7440 | Accept |  | 1- 3220 | Accept |  | 1- 826 | Accept |  | 1- 220 | Accept |  |
| 21600-22400 | 800 | 11 | 7440-8330 | 380 | 6 | 3220-3300 | 155 | 3 | 827- 902 | 40 | 1 | 221-365 | 5 | 0 |
| 22400-23500 | 890 | 12 | 8330- 9140 | 465 | 7 | $3470-3630$ | 230 | 4 | 1050-1120 | 100 | 2 |  |  |  |
| 24300-25500 | 985 | 13 | 9550-10100 | 550 | 8 | 4090-4250 | 240 | 4 | 1290 | 110 | 2 | 366- 425 | 45 | 1 |
| 26400-27700 | 1080 | 14 | 11000-12600 | 640 | 9 | - 200 |  |  |  |  |  | $503-612$ | 55 | 1 |
|  |  |  |  |  |  | 4250- 4470 | 310 | 5 | 1580-1680 | 170 | 3 | 52- 885 | 110 | 2 |
| 28800-30100 | 1175 | 15 | 12600-13600 | 725 | 10 | 5050-5220 | 320 | 5 | 1930- 2260 | 180 | 3 | 1050 | $120 \quad 2$ |  |
| 31300-32800 | 1270 | 16 | 14500-15000 | 810 | 11 | 5220- 5460 | 390 | 6 | 2260- 2360 | 240 | 4 |  |  |  |
| 34200-35700 | 1365 | 17 | 16500-18300 | 900 | 12 | $6170-638$ | 400 | 6 | 2700 | 250 | 4 | 1270-1320 | $170 \quad 3$ |  |
| 37200-38800 | 1460 | 18 | 18900-19900 | 985 | 13 |  |  |  |  |  |  | 1550-1860 | 180 | 3 |
| 40600-42200 | 1555 | 19 | 21600-24600 | 1075 | 14 | $6380-6600$ | 470 | 7 | 3130-3660 | 315 | 5 | 1950-2130 | 235 | 4 |
|  |  |  |  |  |  | $7450-7760$ | 480 | 7 | 4200- 4820 | 385 | 6 | 2520-2880 | 245 | 4 |
| 44300-45900 | 1650 | 20 | 24600-26400 | 1160 | 15 | 7760-8930 | 555 | 8 | 5540-6260 | 455 | 7 |  |  |  |
| 48300-49900 | 1745 | 21 | 28000-31800 | 1250 | 16 | 9360-10600 | 635 | 9 | 7230-8020 | 525 | 8 | 2880-3290 | 300 | 5 |
| 52600-54200 | 1840 | 22 | 31800-34700 | 1335 | 17 |  |  |  | 9340-10200 | 595 | 9 | 3950-4120 | 310 | 5 |
| 57300-58900 | 1935 | 23 | 36100-37500 | 1420 | 18 | 11200-12600 | 715 | 10 |  |  |  | $4120-5000$ | 365 | 6 |
| 62500-63900 | 2030 | 24 | 41000-45400 | 1510 | 19 | 13400-14800 | 795 | 11 | 12000-12800 | 665 | 10 | $5750-6280$ | 425 | 7 |
|  |  |  |  |  |  | 16000-17300 | 875 | 12 | 15300-16100 | 735 | 11 | 7500- | 435 | 7 |
| 68000-69300 | 2125 | 25 | 46400-48900 | 1595 | 20 | 19000-20200 | 955 | 13 | 18800-19500 | 745 | 11 |  |  |  |
| 74100-75200 | 2220 | 26 | 52600-59400 | 1685 | 21 | 22400-23600 | 1035 | 14 |  |  |  | 7970- 9250 | 490 | 8 |
| 80700-87800 | 2320 | 27 | 59400-63600 | 1770 | 22 |  |  |  | 19500-23400 | 810 | 12 | 10900-11400 | 550 | 9 |
| 87800-95500 | 2415 | 28 | 67300-75900 | 1860 | 23 | 26500-27400 | 1115 | 15 | 24700-29200 | 880 | 13 | 13600-14900 | 560 | 9 |
| 95500-104000 | 2510 | 29 | 75900-82500 | 1945 | 24 | 31300-36800 | 1200 | 16 | 31200-36200 | 950 | 14 | $20200-24000$ | 615 | 10 |
|  |  |  |  |  |  | 36800-41900 | 1280 | 17 |  |  |  |  | 680 | 11 |
| 104000-113000 | 2605 | 30 | 85700-88500 | 2030 | 25 | 43300-48400 | 1360 | 18 | 39400-44800 | 1020 | 15 |  |  |  |
| 113000-123000 | 2700 | 31 | 96700-107000 | 2120 | 26 | 50800-55800 | 1440 | 19 | 49500-55400 | 1090 | 16 | 27200-29100 | 740 | 12 |
| 123000-133000 | 2795 | 32 | 109000-114000 | 2205 | 27 |  |  |  | 62200-68300 | 1160 | 17 | 34900-36700 | 750 | 12 |
| 133000-145000 | 2890 | 33 | 123000-138000 | 2295 | 28 | 59500-64400 | 1520 | 20 | 78000-84200 | 1230 | 18 | 36700-41800 | 805 | 13 |
| 145000-158000 | 2985 | 34 | 138000-147000 | 2380 | 29 | 69700-74200 | 1600 | 21 | 97700-104000 | 1300 | 19 |  | 870 | 14 |
|  |  |  |  |  |  | 81600-85400 | 1680 | 22 |  |  |  | $66100-72200$ | 930 | 15 |
| 158000-171000 | 3080 | 35 | 156000-176000 | 2470 | 30 | 95400-98200 | 1760 | 23 | 122000-127000 | 1370 | 20 |  |  |  |
| 171000-186000 | 3175 | 36 | 176000-190000 | 2555 | 31 | 112000-130000 | 1845 | 24 | 153000-183000 | 1445 | 21 | 88300-103000 | 995 | 16 |
| 186000-200000 | 3270 | 37 | 198000-200000 | 2640 | 32 |  |  |  | 192000-200000 | 1515 | 22 |  | 1055 | 17 |
|  |  |  |  |  |  | 130000-148000 | 1925 | 25 |  |  |  | $\begin{aligned} & 149000-158000 \\ & 158000-176000 \end{aligned}$ | 1065 | 17 |
|  |  |  |  |  |  | 152000-169000 | 2005 | 26 |  |  |  |  | 112018 |  |
|  |  |  |  |  |  | 177000-194000 | 2085 | 27 |  |  |  |  |  |  |  |




Relation between $p_{r}$ and $w_{2}$ for fixed ( $p_{10}, p_{20}, \gamma_{2}$ ).
Use the same sampling plan for $w_{2}=0.05$ and $p_{r 0}=0.01(0.10)$ as for $w_{2}$ and $p_{r}=0.01 f(0.10 f)$ where $f$ is given in the table.

| $100 \mathrm{w}_{2}$ | 1.5 | 2.0 | $\begin{aligned} \varrho_{2}= & p_{20} / p_{r 0} \\ & 3.0 \end{aligned}$ | 5.0 | 7.0 | $\begin{gathered} \varrho_{1} \\ =p_{10} / p_{r 0} \end{gathered}$ |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| 1 | 0.51 | 0.44 | 0.40 | 0.38 | 0.37 | 0.2 |
|  | 0.78 | 0.77 | 0.76 | 0.76 | 0.76 | 0.7 |
| 2 | 0.70 |  | 0.58 | 0.55 | 0.53 | 0.2 |
|  | 0.85 | 0.84 | 0.83 | 0.82 | 0.82 | 0.7 |
| 3 | 0.83 | 0.78 | 0.73 | 0.70 | 0.69 | 0.2 |
|  | 0.91 | 0.89 | 0.89 | 0.88 | 0.88 | 0.7 |
| 4 | 0.93 | 0.90 | 0.87 | 0.86 | 0.85 | 0.2 |
|  | 0.96 | 0.95 | 0.94 | 0.94 | 0.94 | 0.7 |
| 5 | 1.00 | 1.00 | 1.00 | 1.00 | 1.00 | 0.2 |
|  | 1.00 | 1.00 | 1.00 | 1.00 | 1.00 | 0.7 |
| 6 | 1.06 | 1.09 | 1.11 | 1.14 | 1.15 | 0.2 |
|  | 1.04 | 1.05 | 1.05 | 1.06 | 1.06 | 0.7 |
| 7 | 1.10 | 1.16 | 1.22 | 1.27 | 1.29 | 0.2 |
|  | 1.07 | 1.09 | 1.11 | 1.12 | 1.12 | 0.7 |
| 8 | 1.14 | 1.22 | 1.31 | 1.39 | 1.43 | 0.2 |
|  | 1.10 | 1.13 | 1.16 | 1.17 | 1.18 | 0.7 |
| 9 | 1.18 | 1.28 | 1.40 | 1.51 | 1.56 | 0.2 |
|  | 1.12 | 1.17 | 1.21 | 1.23 | 1.24 | 0.7 |
| 10 | 1.20 | 1.33 | 1.48 | 1.63 | 1.69 | 0.2 |
|  | 1.15 | 1.20 | 1.25 | 1.29 | 1.30 | 0.7 |
| 12 | 1.25 | 1.41 | 1.63 | 1.84 | 1.95 | 0.2 |
|  | 1.19 | 1.27 | 1.34 | 1.40 | 1.42 | 0.7 |
| 14 | 1.28 | 1.48 | 1.75 | 2.03 | 2.19 | 0.2 |
|  | 1.22 | 1.33 | 1.43 | 1.51 | 1.54 | 0.7 |
| 16 | 1.31 | 1.54 | 1.86 | 2.22 | 2.41 | 0.2 |
|  | 1.25 | 1.38 | 1.51 | 1.62 | 1.67 | 0.7 |
| 18 | 1.33 | 1.58 | 1.95 | 2.38 | 2.63 | 0.2 |
|  | 1.27 | 1.42 | 1.59 | 1.72 | 1.79 | 0.7 |
| 20 | 1.35 | 1.63 | 2.03 | 2.54 | 2.84 | 0.2 |
|  | 1.29 | 1.46 | 1.66 | 1.83 | 1.91 | 0.7 |

Table of $b_{1}, b_{2}$, and $b_{3}$.

| $\begin{aligned} & \varrho_{1} \\ & = \\ & p_{1} / p_{r} \end{aligned}$ | 1.5 | 2.0 | $=p_{2} / p_{r}$ | 5.0 | 7.0 |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| 0.2 | 0.63 | 0.64 | 0.65 | 0.66 | 0.67 | $b_{1}$ |
|  | 0.24 | 0.27 | 0.31 | 0.35 | 0.37 | $b_{2}$ |
|  | 1.42 | 1.29 | 1.15 | 1.03 | 0.96 | $b_{3}$ |
| 0.3 | 0.61 | 0.62 | 0.64 | 0.65 | 0.65 | $b_{1}$ |
|  | 0.19 | 0.23 | 0.27 | 0.31 | 0.34 | $b_{2}$ |
|  | 1.69 | 1.48 | 1.28 | 1.11 | 1.03 | $b_{3}$ |
| 0.4 | 0.60 | 0.61 | 0.63 | 0.64 | 0.64 | $b_{1}$ |
|  | 0.16 | 0.19 | 0.24 | 0.29 | 0.32 | $b_{2}$ |
|  | 1.99 | 1.69 | 1.41 | 1.19 | 1.08 | $b_{3}$ |
| 0.5 | 0.59 | 0.60 | 0.62 | 0.63 | 0.63 | $b_{1}$ |
|  | 0.13 | 0.17 | 0.21 | 0.27 | 0.30 | $b_{2}$ |
|  | 2.33 | 1.91 | 1.54 | 1.27 | 1.14 | $b_{3}$ |
| 0.6 | 0.58 | 0.59 | 0.61 | 0.62 | 0.62 | $b_{1}$ |
|  | 0.11 | 0.15 | 0.19 | 0.25 | 0.28 | $b_{2}$ |
|  | 2.74 | 2.15 | 1.68 | 1.34 | 1.19 | $b_{3}$ |
| 0.7 | 0.58 | 0.59 | 0.60 | 0.62 | 0.62 | $b_{1}$ |
|  | 0.09 | 0.13 | 0.18 | 0.23 | 0.26 | $b_{2}$ |
|  | 3.24 | 2.42 | 1.82 | 1.42 | 1.25 | $b_{3}$ |

Conversion factor $f_{2}$ for $N$ due to a change in $p_{r}=p_{s}$ for fixed $\left(p_{1}, p_{2}, w_{2}\right)$. Use $N^{*}=N f_{2}$ as argument in the master table to find $\left(n^{*}, c^{*}\right)$. $p_{r}=p_{s}=0.01 \lambda$ or $0.10 \lambda,\left(p_{1}, p_{2}, w_{2}\right)$ are given in the master tables, $\varrho_{1}=100 p_{1}$ or $10 p_{1}, \varrho_{2}=100 p_{2}$ or $10 p_{2}$.

| $\varrho_{2} \quad \varrho_{1}$ | $\lambda=0.50$ | 0.60 | 0.70 | 0.80 | 0.90 | 1.00 | 1.25 | 1.50 | 1.75 | 2.00 | 3.00 |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| 0.2 | 1.80 | 1.52 | 1.33 | 1.18 | 1.08 | 1.00 | - | - | - | - |  |
| 0.3 | 2.21 | 1.79 | 1.50 | 1.28 | 1.12 | 1.00 | - | - | - | - |  |
| 1.50 .4 | - | 2.11 | 1.69 | 1.38 | 1.16 | 1.00 | - | - | - | - |  |
| 0.5 | - | - | 1.90 | 1.50 | 1.21 | 1.00 | - | - | - | - |  |
| 0.6 | - |  | - | 1.64 | 1.27 | 1.00 | - | - | - | - |  |
| 0.7 | - | - | - | - | - | 1.00 | - | - | - | - | - |
| ( 0.2 | 1.69 | 1.47 | 1.30 | 1.18 | 1.08 | 1.00 | 0.87 | - | - | - | - |
| 0.3 | 1.92 | 1.63 | 1.41 | 1.24 | 1.10 | 1.00 | 0.82 | - | - | - |  |
| 2.00 .4 | - | 1.79 | 1.52 | 1.30 | 1.13 | 1.00 | 0.78 | - | - | - |  |
| 2.00 .5 | - | - | 1.62 | 1.36 | 1.16 | 1.00 | 0.73 | 0.58 | - | - |  |
| 0.6 | - | - | - | 1.42 | 1.19 | 1.00 | 0.69 | 0.52 | - | - |  |
| 0.7 | - | - | - | - | 1.21 | 1.00 | 0.65 | 0.47 | - | - | - |
| 0.2 | 1.53 | 1.38 | 1.25 | 1.15 | 1.07 | 1.00 | 0.88 | 0.80 | - | - |  |
| 0.3 | 1.64 | 1.46 | 1.31 | 1.19 | 1.08 | 1.00 | 0.85 | 0.74 | 0.67 | - |  |
| 3.00 .4 | - | 1.52 | 1.36 | 1.22 | 1.10 | 1.00 | 0.82 | 0.70 | 0.62 | 0.56 |  |
| $3.0 \mid 0.5$ | - | - | - | 1.24 | 1.11 | 1.00 | 0.80 | 0.66 | 0.57 | 0.50 | - |
| 0.6 | - | - | - | - | 1.12 | 1.00 | 0.78 | 0.63 | 0.53 | 0.46 | - |
| 0.7 | - | - | - | - | - | 1.00 | 0.76 | 0.60 | 0.49 | 0.41 | - |
| 0.2 | 1.39 | 1.28 | 1.19 | 1.12 | 1.05 | 1.00 | 0.89 | 0.82 | 0.76 | 0.72 | - |
| 0.3 | - | 1.31 | 1.22 | 1.13 | 1.06 | 1.00 | 0.88 | 0.79 | 0.72 | 0.67 | - |
| 5.00 .4 | - | - | 1.23 | 1.14 | 1.07 | 1.00 | 0.86 | 0.76 | 0.69 | 0.63 | 0.50 |
| 5.00 .5 | - | - | - | 1.14 | 1.07 | 1.00 | 0.85 | 0.74 | 0.66 | 0.60 | 0.46 |
| 0.6 | - | - | - | - | - | 1.00 | 0.85 | 0.73 | 0.65 | 0.58 | 0.42 |
| 0.7 | - | - | - | - | - | 1.00 | 0.85 | 0.73 | 0.63 | 0.56 | 0.39 |
| (0.2 | 1.31 | 1.23 | 1.16 | 1.10 | 1.05 | 1.00 | 0.91 | 0.84 | 0.79 | 0.74 | 0.65 |
| 0.3 | - | 1.24 | 1.17 | 1.11 | 1.05 | 1.00 | 0.89 | 0.81 | 0.75 | 0.70 | 0.59 |
| 70.4 | - | - | - | 1.11 | 1.05 | 1.00 | 0.89 | 0.80 | 0.74 | 0.68 | 0.55 |
| 7.0.5 | - | - | - | - | 1.05 | 1.00 | 0.89 | 0.79 | 0.72 | 0.66 | 0.52 |
| 0.6 | - | - | - | - | - | 1.00 | 0.89 | 0.79 | 0.72 | 0.65 | 0.50 |
| 0.7 | - | - | - | - | - | 1.00 | 0.90 | 0.80 | 0.72 | 0.65 | 0.48 |

Correction $g_{2}$ to $n^{*}$ due to a change in $p_{r}=p_{s}$.
Reference value $p_{r}=p_{s}=0.010 . n=n^{*}+g_{2}$.

| $\varrho_{2} \quad \varrho_{1}$ | $\lambda=0.50$ | 0.60 | 0.70 | 0.80 | 0.90 | 1.00 | 1.25 | 1.50 | 1.75 | 2.00 | 3.00 |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| $(0.2$ | 130 | 100 | 70 | 50 | 25 | 0 | - | - | - | - | - |
| 0.3 | 160 | 120 | 85 | 55 | 30 | 0 | - | - | - | - | - |
| 5 0.4 | - | 150 | 105 | 65 | 35 | 0 | - | - | - | - | - |
| 1.500 .5 | - | - | 135 | 85 | 40 | 0 | - | - | - | - | - |
| 0.6 | - | - | - | 115 | 50 | 0 | - | - | - | - | - |
| 0.7 | - | - | - | - | - | 0 | - | - | - | - | - |
| ( 0.2 | 75 | 55 | 40 | 25 | 15 | 0 | -30 | - | - | - | - |
| 0.3 | 95 | 70 | 50 | 30 | 15 | 0 | -35 | - | - | - | - |
| 0.4 | - | 90 | 60 | 35 | 15 | 0 | -40 | - | - | - | - |
| 2.00 .5 | - | - | 80 | 45 | 20 | 0 | -45 | - 90 | - | - | - |
| 0.6 | - | - | - | 60 | 25 | 0 | - 55 | - 105 | - | - | - |
| 0.7 | - | - | - | - | 40 | 0 | -70 | - 125 | - | - | - |
| 0.2 | 40 | 30 | 20 | 15 | 5 | 0 | - 15 | - 25 | - | - | - |
| 0.3 | 55 | 40 | 25 | 15 | 5 | 0 | -15 | - 30 | -45 | - | - |
| ${ }_{0} 0.4$ | - | 50 | 30 | 20 | 10 | 0 | -20 | - 35 | - 50 | -65 | - |
| 0.5 | - | - | - | 25 | 10 | 0 | -20 | - 40 | - 55 | - 70 | - |
| 0.6 | - | - | - | - | 15 | 0 | -25 | - 45 | -60 | -80 | - |
| 0.7 | - | - | - | - | - | 0 | -30 | - 55 | - 75 | -90 | - |
| 0.2 | 20 | 15 | 10 | 5 | 5 | 0 | - 5 | - 15 | -20 | - 20 | - |
| 0.3 | - | 20 | 15 | 10 | 5 | 0 | - 10 | - 15 | -20 | -25 | - |
| 5.00 .4 | - | - | 15 | 10 | 5 | 0 | - 10 | - 15 | -20 | - 25 | -45 |
| 5.00 .5 | - | - | - | 10 | 5 | 0 | -10 | - 20 | -25 | -30 | - 50 |
| 0.6 | - | - | - | - | - | 0 | - 10 | - 20 | -30 | - 35 | - 55 |
| 0.7 | - | - | - | - | - | 0 | -15 | - 25 | -35 | -40 | -60 |
| 0.2 | 15 | 10 | 5 | 5 | 0 | 0 |  | - 10 | - 10 | - 15 | - 25 |
| 0.3 | - | 15 | 10 | 5 | 0 | 0 | - 5 | - 10 | -10 | -15 | -25 |
| 7.00 .4 | - | - | - | 5 | 5 | 0 | - 5 | - 10 | -15 | -15 | - 25 |
| ${ }^{7.0} 0.5$ | - | - | - | - | 5 | 0 | - 5 | - 10 | -15 | -20 | -30 |
| 0.6 | - | - | - | - | - | 0 | - 10 | - 15 | -20 | -20 | -35 |
| 0.7 | - | - | - | - | - | 0 | -10 | - 15 | -20 | -25 | -35 |

For $p_{r}=p_{s}=0.10$ the correction is $g_{2} / 10$ (rounded down).

Conversion factor $f_{1}$ for $N$ due to a change in $w_{2}$.
Reference value of $w_{2}=0.05, p_{s}=p_{r}$.
Use $N^{*}=N f_{1}$ as argument in the master table to find ( $n^{*}, c^{*}$ ).

| $100 w_{2}$ | 1.5 | 2.0 | $\begin{array}{r} p_{2} / p_{r} \\ 3.0 \end{array}$ | 5.0 | 7.0 | $p_{1} / P_{r}$ |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| 1 | 0.54 | 0.56 | 0.58 | 0.61 | 0.63 | 0.2 |
|  | 0.46 | 0.48 | 0.51 | 0.54 | 0.57 | 0.7 |
| 2 | 0.70 | 0.72 | 0.74 | 0.76 | 0.77 | 0.2 |
|  | 0.65 | 0.66 | 0.68 | 0.71 | 0.73 | 0.7 |
| 3 | 0.82 | 0.83 | 0.84 | 0.86 | 0.87 | 0.2 |
|  | 0.78 | 0.79 | 0.81 | 0.83 | 0.84 | 0.7 |
| 4 | 0.92 | 0.92 | 0.93 | 0.94 | 0.94 | 0.2 |
|  | 0.90 | 0.90 | 0.91 | 0.92 | 0.93 | 0.7 |
| 5 | 1.00 | 1.00 | 1.00 | 1.00 | 1.00 | 0.2 |
|  | 1.00 | 1.00 | 1.00 | 1.00 | 1.00 | 0.7 |
| 6 | 1.07 | 1.07 | 1.06 | 1.06 | 1.05 | 0.2 |
|  | 1.09 | 1.09 | 1.08 | 1.07 | 1.06 | 0.7 |
| 7 | 1.14 | 1.13 | 1.12 | 1.10 | 1.09 | 0.2 |
|  | 1.17 | 1.16 | 1.15 | 1.13 | 1.12 | 0.7 |
| 8 | 1.19 | 1.18 | 1.16 | 1.15 | 1.13 | 0.2 |
|  | 1.25 | 1.24 | 1.21 | 1.19 | 1.17 | 0.7 |
| 9 | 1.25 | 1.23 | 1.21 | 1.18 | 1.17 | 0.2 |
|  | 1.32 | 1.30 | 1.27 | 1.24 | 1.22 | 0.7 |
| 10 | 1.30 | 1.27 | 1.25 | 1.22 | 1.20 | 0.2 |
|  | 1.39 | 1.37 | 1.33 | 1.29 | 1.26 | 0.7 |
| 12 | 1.38 | 1.36 | 1.32 | 1.28 | 1.25 | 0.2 |
|  | 1.51 | 1.48 | 1.43 | 1.38 | 1.34 | 0.7 |
| 14 | 1.46 | 1.43 | 1.38 | 1.33 | 1.30 | 0.2 |
|  | 1.63 | 1.58 | 1.52 | 1.45 | 1.40 | 0.7 |
| 16 | 1.53 | 1.49 | 1.44 | 1.38 | 1.34 | 0.2 |
|  | 1.73 | 1.68 | 1.61 | 1.52 | 1.46 | 0.7 |
| 18 | 1.60 | 1.55 | 1.49 | 1.42 | 1.38 | 0.2 |
|  | 1.83 | 1.77 | 1.68 | 1.58 | 1.52 | 0.7 |
| 20 | 1.65 | 1.60 | 1.53 | 1.46 | 1.41 | 0.2 |
|  | 1.92 | 1.85 | 1.75 | 1.64 | 1.56 | 0.7 |

Correction $g_{1}$ to $n^{*}$ due to a change in $w_{2}$.
Reference value of $w_{2}=0.05, p_{s}=p_{r}=0.01 . n=n^{*}+g_{1}$.

| $100 w_{2}$ | 1.5 | 2.0 | $\begin{array}{r} p_{2} / p_{r} \\ 3.0 \end{array}$ | 5.0 | 7.0 | $p_{1} / p_{r}$ |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| 1 | $\begin{aligned} & -125 \\ & -205 \end{aligned}$ | $\begin{aligned} & -90 \\ & -125 \end{aligned}$ | $\begin{aligned} & -60 \\ & -70 \end{aligned}$ | $\begin{aligned} & -35 \\ & -35 \end{aligned}$ | $\begin{aligned} & -25 \\ & -25 \end{aligned}$ | $\begin{aligned} & 0.2 \\ & 0.7 \end{aligned}$ |
| 2 | $\begin{aligned} & -\quad 70 \\ & -115 \end{aligned}$ | $\begin{array}{r} -\quad 50 \\ -\quad 70 \end{array}$ | $\begin{aligned} & -35 \\ & -40 \end{aligned}$ | $\begin{aligned} & -20 \\ & -20 \end{aligned}$ | $\begin{aligned} & -15 \\ & -15 \end{aligned}$ | $\begin{aligned} & 0.2 \\ & 0.7 \end{aligned}$ |
| 3 | $\begin{array}{r} -\quad 40 \\ -\quad 65 \end{array}$ | $\begin{aligned} & -\quad 30 \\ & -\quad 40 \end{aligned}$ | $\begin{aligned} & -20 \\ & -25 \end{aligned}$ | $\begin{aligned} & -10 \\ & -10 \end{aligned}$ | $\begin{aligned} & -10 \\ & -10 \end{aligned}$ | $\begin{aligned} & 0.2 \\ & 0.7 \end{aligned}$ |
| 4 | $\begin{aligned} & -\quad 20 \\ & -\quad 30 \end{aligned}$ | $\begin{array}{r} -\quad 15 \\ -\quad 20 \end{array}$ | $\begin{aligned} & -10 \\ & -10 \end{aligned}$ | $\begin{aligned} & -5 \\ & -5 \end{aligned}$ | $\begin{array}{r} -5 \\ -5 \end{array}$ | $\begin{aligned} & 0.2 \\ & 0.7 \end{aligned}$ |
| 5 | $\begin{aligned} & 0 \\ & 0 \end{aligned}$ | $\begin{aligned} & 0 \\ & 0 \end{aligned}$ | $\begin{aligned} & 0 \\ & 0 \end{aligned}$ | 0 0 | 0 0 | $\begin{aligned} & 0.2 \\ & 0.7 \end{aligned}$ |
| 6 | $\begin{aligned} & 15 \\ & 25 \end{aligned}$ | $\begin{aligned} & 10 \\ & 15 \end{aligned}$ | $\begin{array}{r} 5 \\ 10 \end{array}$ | 5 5 | 5 5 | $\begin{aligned} & 0.2 \\ & 0.7 \end{aligned}$ |
| 7 | $\begin{aligned} & 25 \\ & 45 \end{aligned}$ | $\begin{aligned} & 20 \\ & 25 \end{aligned}$ | $\begin{aligned} & 15 \\ & 15 \end{aligned}$ | 5 10 | 5 5 | $\begin{aligned} & 0.2 \\ & 0.7 \end{aligned}$ |
| 8 | $\begin{aligned} & 40 \\ & 60 \end{aligned}$ | $\begin{aligned} & 30 \\ & 40 \end{aligned}$ | $\begin{aligned} & 20 \\ & 20 \end{aligned}$ | $\begin{aligned} & 10 \\ & 10 \end{aligned}$ | 5 10 | $\begin{aligned} & 0.2 \\ & 0.7 \end{aligned}$ |
| 9 | $\begin{aligned} & 50 \\ & 80 \end{aligned}$ | $\begin{aligned} & 35 \\ & 50 \end{aligned}$ | $\begin{aligned} & 20 \\ & 25 \end{aligned}$ | 15 15 | 10 10 | $\begin{aligned} & 0.2 \\ & 0.7 \end{aligned}$ |
| 10 | $\begin{aligned} & 55 \\ & 90 \end{aligned}$ | $\begin{aligned} & 40 \\ & 55 \end{aligned}$ | $\begin{aligned} & 25 \\ & 30 \end{aligned}$ | $\begin{aligned} & 15 \\ & 15 \end{aligned}$ | 10 10 | $\begin{aligned} & 0.2 \\ & 0.7 \end{aligned}$ |
| 12 | $\begin{array}{r} 75 \\ 120 \end{array}$ | $\begin{aligned} & 50 \\ & 70 \end{aligned}$ | $\begin{aligned} & 35 \\ & 40 \end{aligned}$ | $\begin{aligned} & 20 \\ & 20 \end{aligned}$ | 15 15 | $\begin{aligned} & 0.2 \\ & 0.7 \end{aligned}$ |
| 14 | $\begin{array}{r} 85 \\ 140 \end{array}$ | $\begin{aligned} & 60 \\ & 85 \end{aligned}$ | $\begin{aligned} & 40 \\ & 50 \end{aligned}$ | 25 25 | 15 15 | $\begin{aligned} & 0.2 \\ & 0.7 \end{aligned}$ |
| 16 | $\begin{aligned} & 100 \\ & 160 \end{aligned}$ | $\begin{array}{r} 70 \\ 100 \end{array}$ | $\begin{aligned} & 45 \\ & 55 \end{aligned}$ | $\begin{aligned} & 25 \\ & 30 \end{aligned}$ | 20 20 | $\begin{aligned} & 0.2 \\ & 0.7 \end{aligned}$ |
| 18 | $\begin{aligned} & 110 \\ & 175 \end{aligned}$ | $\begin{array}{r} 80 \\ 110 \end{array}$ | 50 60 | 30 30 | 20 20 | $\begin{aligned} & 0.2 \\ & 0.7 \end{aligned}$ |
| 20 | $\begin{aligned} & 120 \\ & 195 \end{aligned}$ | $\begin{array}{r} 85 \\ 120 \end{array}$ | $\begin{aligned} & 55 \\ & 65 \end{aligned}$ | $\begin{aligned} & 30 \\ & 35 \end{aligned}$ | 20 25 | $\begin{aligned} & 0.2 \\ & 0.7 \end{aligned}$ |

For $p_{s}=p_{r}=0.10$ the correction is $g_{1} / 10$ (rounded down).

$$
\begin{gathered}
\begin{array}{c}
\text { Summary of conversion formulas } \\
\text { to find }
\end{array} \\
(n, c) \text { corresponding to }\left(N, p_{r}, p_{s}, p_{1}, p_{2}, w_{2}\right) \\
\text { from }
\end{gathered} \begin{aligned}
& \left(n^{*}, c^{*}\right) \text { in the master table for }\left(N^{*}, p_{r 0}, p_{10}, p_{20}\right) . \\
& \text { For }\left\{\begin{array}{l}
p_{r} \leqq 0.05 \\
p_{r}>0.05
\end{array}\right\} \text { use master table with } p_{r 0}=\left\{\begin{array}{l}
0.01 \\
0.10
\end{array}\right. \\
& \lambda_{s}=\left(1+\frac{p_{s}-p_{r}}{w_{1}\left(p_{r}-p_{1}\right)}\right)^{-1} .
\end{aligned}
$$

Formula 1.

$$
\gamma_{2}=\frac{w_{2}\left(p_{2}-p_{r}\right)}{w_{1}\left(p_{r}-p_{1}\right)} \quad \text { and } \quad \lambda p_{r 0}=\frac{p_{2}+19 \gamma_{2} p_{1}}{1+19 \gamma_{2}} .
$$

Use

$$
N^{*}=N \lambda_{s} \lambda, \quad p_{r 0}, \quad p_{10}=p_{1} / \lambda, \quad p_{20}=p_{2} / \lambda
$$

as arguments to find $\left(n^{*}, c^{*}\right)$ in the master table.

$$
(n, c)=\left(n^{*} / \lambda, c^{*}\right) .
$$

If $\left(p_{10}, p_{20}\right)$ fall outside the tabulated range use formula 2.
Formula 2.

$$
\lambda=p_{r} / p_{r 0}, \quad \varrho_{1}=p_{1} / p_{r}, \quad \varrho_{2}=p_{2} / p_{r} .
$$

Use

$$
N^{*}=N \lambda_{s} \lambda f_{1}\left(w_{2}, \varrho_{1}, \varrho_{2}\right), \quad p_{r 0}, \quad p_{10}=\varrho_{1} p_{r 0}, \quad p_{20}=\varrho_{2} p_{r 0}
$$

as arguments to find $\left(n^{*}, c^{*}\right)$ in the master table.

$$
(n, c)=\left(\left(n^{*}+g_{1}\left(w_{2}, \varrho_{1}, \varrho_{2}\right)\right) / \lambda, c^{*}\right) .
$$

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[^0]:    For $N$ between two intervals adjacent in the table find ( $n, c$ ) for the first of these intervals and use ( $n+5, c$ ) as optimum plan.

[^1]:    For $N$ between two intervals adjacent in the table find $(n, c)$ for the first of these intervals and use $(n+5, c)$ as optimum plan.

[^2]:    

[^3]:    For $N$ between two intervals adjacent in the table find $(n, c)$ for the first of these intervals and use $(n+5, c)$ as optimum plan.

[^4]:    

